NUS-RMI Credit Research Initiative Technical Report

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RMI staff article

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his document describes the implementation of the system which the Credit Research Initiative (CRI) at the Risk Management Institute (RMI) of the National University of Singapore (NUS) uses to produce probabilities of default (PDs). As of this version of the Technical Report, RMI covers around 60,400 listed firms (including delisted ones) in 106 economies around the world (see Table A.1). Of the over 39,000 active firms under the CRI coverage, around 34,000 firms have sufficient data to release daily updated PDs. The PD for all firms is freely available to users who can provide evidence of their professional qualifications to ensure that they will not misuse the data. General users who do not request global access are restricted to a list of 3,000 firms. The individual company PD data, along with aggregate PDs at the economy and sector level, can be accessed at http://rmicri.org.

The primary goal of this initiative is to drive research and development in the critical area of credit rating systems. As such, a transparent methodology is essential to this initiative. Having the details of the methodology available to everybody means that there is a base from which suggestions and improvements can be made. The objective of this Technical Report is to provide a full exposition of the CRI system. Readers of this document who have access to the necessary data and who have a sufficient level of technical expertise will be able to implement a similar system on their own. For a full exposition of the conceptual framework of the CRI, see Duan and Van Laere (2012).

The system used by the CRI will evolve as new innovations and enhancements are applied. The changes to the 2012 technical report and operational implementation of our model are: (1) RMI's global coverage; (2) extension of the forecast horizon to five years by applying a

GLOBAL CREDIT REVIEW VOLUME 3 77

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Nelson-Siegel type parameterization and using a Sequential Monte Carlo (SMC) method; (3) changes to treatment of financial statements in monthly calibration; (4) exclusion rule for merger and acquisition (M&A) events; (5) changes to treatment of companies after a default event; (6) change from monthly sigma to daily sigma; (7) changes in the treatment of missing values; and (8) changes in the level and trend calculations and (9) changes in the Distance-to-Default computation. This version of the technical report provides an update on the operational implementation of the CRI and includes all changes to the system that had been implemented by September 2013. More specifically, in addition to Version: 2013 update 1, the current version of the technical report specifies some revisions to the monthly parameter updates that went into effect as of the August 2013 calibration. The latest version of the Technical Report and addenda to the latest version are available via the web portal and will include any changes to the system that have been implemented since the publication of this version.

The remainder of this Technical Report is organized as follows. The next section describes the quantitative model that is currently used to compute PDs from the CRI. The model was first described in Duan *et al.* (2012). The description includes calibration procedures, which are performed on a monthly basis, and individual firm PD computations, which are performed on a daily basis.

Section 2 describes the input variables of the model as well as the data used to produce the variables for input into the model. This model uses both input variables that are common to all firms in an economy and input variables that are firm-specific. Another critical component when calibrating a probability of default estimation system is the default data, and this is also described in this section.

While Section 1 provides a broader description of the model, Section 3 describes the implementation details that are necessary for application, given real world issues of, for example, bad or missing data. The specific technical details needed to develop an operational system are also given, including details on the monthly calibration, daily computation of individual firm PDs and aggregation of the individual firm PDs. Distance-to default (DTD) in a Merton-type model is one of the

firm-specific variables. The calculation for DTD is not the standard one, and has been modified to allow a meaningful computation of the DTD for financial firms. While most academic studies on default prediction exclude financial firms from consideration, it is important to include them given that the financial sector is a critical component in every economy. The calculation for DTD is detailed in this section.

Section 4 shows an empirical analysis for those economies that are currently covered. While the analysis shows excellent results in several economies, there is room for improvement in a few others. This is because, at the CRI's current stage of development, the economies all use the variables used in the academic study of US firms in Duan *et al.* (2012). Future development within the CRI will deal with variable selection specific to different economies, and the performance is then expected to improve. Other planned developments are discussed in Section 5.

I. MODEL DESCRIPTION

The quantitative model that is currently being used by the CRI is a forward intensity model that was introduced in Duan et al. (2012). Certain aspects of the model are taken from Duan and Fulop (2013). This model allows PD forecasts to be made at a range of horizons. In the current CRI implementation of this model, PDs are forecasted from a horizon of one month up to a horizon of five years. At the RMI CRI website, for every firm, the probabilities of that firm defaulting within one month, three months, six months, one year, two years, three years and five years are given. The ability to assess credit quality for different horizons is a useful tool for risk management, credit portfolio management, policy setting and regulatory purposes, since short- and long-term credit risk profiles can differ greatly depending on a firm's liquidity, debt structures and other factors.

The forward intensity model is a reduced form model in which the PD is computed as a function of different input variables. These can be firm-specific or common to all firms within an economy. The other category of the default prediction model is the structural model, whereby the corporate structure of a firm is modeled in order to assess the firm's PD.

78 $\,$ NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT $\,$

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A similar reduced form model by Duffie *et al.* (2007) relies on modeling the time series dynamics of the input variables in order to make PD forecasts for different horizons. However, there is little consensus on assumptions for the dynamics of variables such as accounting ratios, and the model output will be highly dependent on these assumptions. In addition, the time series dynamics will be of very high dimension. For example, with the two common variables and two firm-specific variables that Duffie *et al.* (2007) use, a sample of 10,000 firms gives a dimension of the state variables of 20,002.

Given the complexity in modeling the dynamics of variables such as accounting ratios, this model will be difficult to implement if different forecast horizons are required. The key innovation of the forward intensity model is that PD for different horizons can be consistently and efficiently computed based only on the value of the input variables at the time the prediction is made. Thus, the model specification becomes far more tractable.

Fully specifying a reduced form model includes the specification of the function that computes a PD from the input variables. This function is parameterized, and finding appropriate parameter values is called calibrating the model. The forward intensity model can be calibrated by maximizing a pseudo-likelihood function. The calibration is carried out by groups of economies and all firms within a group of economies will use the same parameter values along with each firm's variables in order to compute the firm's PD.

Subsection 1.1 will describe the modeling framework, including the way PDs are computed based on a set of parameter values for the economy and a set of input variables for a firm. Subsection 1.2 explains how the model can be calibrated. Subsection 1.3 details the way parameters are estimated based on the SMC technique.

1.1. Modeling Framework

While the model can be formulated in a continuous time framework, as done in Duan *et al.* (2012), an operational implementation requires discretization in time. Since the model is more easily understood in discrete time, the following exposition of the model will begin in a discrete time framework.

Variables for default prediction can have vastly different update frequencies. Financial statement data is updated only once a quarter or even once a year, while market data like stock prices are available at frequencies of seconds. A way of compromising between these two extremes is to have a fundamental time period Δt of one month in the modeling framework. As will be seen later, this does not preclude updating the PD forecasts on a daily basis. This is important since, for example, large daily changes in a firm's stock price can signal changes in credit quality even when there is no change in financial statement data.

Thus, for the purpose of calibration and subsequently for computing time series of PD, the input variables at the end of each month will be kept for each firm. The input variables associated with the i^{th} firm at the end of the n^{th} month (at time $t = n\Delta t$) is denoted by $X_i(n)$. This is a vector consisting of two parts: $X_i(n) = (W(n), U_i(n))$. Here, W(n) is a vector of variables at the end of month *n* that is common to all firms in the economy and $U_i(n)$ is a vector of variables specific to firm *i*.

In the forward intensity model, a firm's default is signaled by a jump in a Poisson process. The probability of a jump in the Poisson process is determined by the intensity of the Poisson process. The forward intensity model draws an explicit dependence of intensities at time periods in the future (that is, forward intensities) to the values of input variables at the time of prediction. With forward intensities, PDs for any forecast horizon can be computed knowing only the values of the input variables at the time of prediction, without needing to simulate future values of the input variables.

There is a direct analogy in interest rate modeling. In spot rate models where dynamics on a short-term spot rate are specified, bond pricing requires expectations on realizations of the short rate. Alternatively, bond prices can be computed directly if the forward rate curve is known.

One issue in default prediction is that firms can exit public exchanges for reasons other than default. For example, in mergers and acquisitions involving two public companies, there will be one company that delists from its stock exchange. This is important in

GLOBAL CREDIT REVIEW VOLUME 3 79

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predicting defaults because a default cannot happen if a firm has been previously delisted. An exception is if the exit is a distressed exit and is followed soon after by a credit event. See Subsection 2.4 for details on how this case is handled in the CRI system.

In order to take these other exits into account, defaults and other exits are modeled as two independent Poisson processes, each with their own intensity. While defaults and exits classified as non-defaults are mutually exclusive by definition, the assumption of independent Poisson processes does not pose a problem since the probability of a simultaneous jump in the two Poisson processes is negligible. In the discrete time framework, the probability of simultaneous jumps in the same time interval is non-zero. As a modeling assumption, a simultaneous jump in the same time interval by both the default Poisson process and the non-default type exit Poisson process is considered as a default. In this way, there are three mutually exclusive possibilities during each time interval: survival, default and non-default exit. As with defaults, the forward intensity of the Poisson process for other exits is a function of the input variables. The parameters of this function can also be calibrated.

To further illustrate the discrete framework, the three possibilities for a firm at each time point are diagrammed. Either the firm survives for the next time period Δt , or it defaults within Δt , or it has a non-default exit within Δt . This setup is pictured in Figure 1. Information about firm *i* is known up until time $t = m\Delta t$ and the figure illustrates possibilities in the future between $t = (n - 1)\Delta t$ and $(n + 1)\Delta t$. Here, *m* and *n* are integers with m < n. The probabilities of each branch are, for example: $p_i(m, n)$ the conditional probability viewed from $t = m\Delta t$ that firm *i* will default before $(n + 1)\Delta t$, conditioned on firm *i* surviving up until $n\Delta t$. Likewise, $\bar{p}_i(m, n)$ is the conditional probability viewed from $t = m\Delta t$ that firm *i* will have a non-default exit before $(n + 1)\Delta t$, conditioned on firm *i* surviving up until $n\Delta t$. It is the modeler's objective to determine $p_i(m, n)$ and $\bar{p}_i(m, n)$, but for now it is assumed that these quantities are known. With the conditional default and other exit probabilities known, the corresponding conditional survival probability of firm *i* is $1 - p_i(m, n) - \bar{p}_i(m, n)$.

With this diagram in mind, the probability that a particular path will be followed is the product of the conditional probabilities along the path. For example, the probability at time $t = m\Delta t$ of firm *i* surviving until $(n - 1)\Delta t$ and then defaulting between $(n - 1)\Delta t$ and $n\Delta t$ is:

$$Prob_{t=m\Delta t}[\tau_i = n, \tau_i < \overline{\tau_i}]$$

= $p_i(m, n-1) \prod_{j=m}^{n-2} [1 - p_i(m, j) - \overline{p}_i(m, j)].$ (1)

Here, τ_i is the default time for firm *i* measured in units of months, $\bar{\tau}_i$ is the other exit time measured in units of months, and the product is equal to 1 if there is no term in the product. The condition $\tau_i < \bar{\tau}_i$ is the requirement that the firm defaults before it has a nondefault type of exit. Note that by measuring exits in units of months, if, for example, a default occurs at any time in the interval $[(n-1)\Delta t, n\Delta t]$, then $\tau_i = n$.

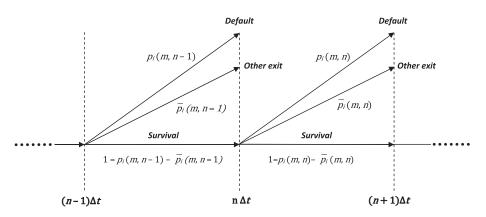


Figure 1. Default-other exit-survival tree for firm *i*, viewed from time $t = m\Delta t$.

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⁸⁰ NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

Using Equation (1), cumulative default probabilities can be computed. At $m\Delta t$ the probability of firm *i* defaulting at or before $n\Delta t$ and not having an other exit before $t = n\Delta t$ is obtained by taking the sum of all of the paths that lead to default at or before $n\Delta t$:

$$Prob_{t=m\Delta t}[m < \tau_i \le n, \tau_i < \overline{\tau}_i] = \sum_{k=m}^{n-1} \left\{ p_i(m,k) \prod_{j=m}^{k-1} [1 - p_i(m,j) - \overline{p}_i(m,j)] \right\}.$$
 (2)

While it is convenient to derive the probabilities given in Equations (1) and (2) in terms of the conditional probabilities, expressions for these in terms of the forward intensities need to be found, since the forward intensities will be functions of the input variable $X_i(m)$. The forward intensity for the default of firm *i* that is observed at time $t = m\Delta t$ for the forward time interval from $t = n\Delta t$ to $(n + 1)\Delta t$, is denoted by $h_i(m,n)$, where $m \leq n$. The corresponding forward intensity for a non-default exit is denoted by $\overline{h}_i(m,n)$. Because default is signaled by a jump in a Poisson process, its conditional probability is a simple function of its forward intensity:

$$p_i(m, n) = 1 - \exp[-\Delta t h_i(m, n)].$$
 (3)

Since joint jumps in the same time interval are assigned as defaults, the conditional other exit probability needs to take this into account:

$$\bar{p}_{i}(m, n) = \exp[-\Delta t \ h_{i}(m, n)] \{1 - \exp[-\Delta t \ \bar{h}_{i}(m, n)]\}.$$
(4)

The conditional survival probabilities in Equations (1) and (2) are computed as the conditional probability that the firm does not default in the period and the firm does not have a non-default exit either:

$$Prob_{t=m\Delta t} [\tau_i, \bar{\tau}_i > n+1 | \tau_i, \bar{\tau}_i > n] = \exp\{-\Delta t [h_i(m, n) + \overline{h}_i(m, n)]\}.$$
 (5)

It remains to be specified the dependence of the forward intensities on the input variable $X_i(m)$. The forward intensities need to be positive so that the conditional probabilities are non-negative. A standard way to impose this constraint is to specify the forward intensities as exponentials of a linear combination of the input variables:

$$\begin{aligned} h_i(m,n) &= \exp[\beta(n-m) \cdot Y_i(m)], \\ \bar{h}_i(m,n) &= \exp[\bar{\beta}(n-m) \cdot Y_i(m)]. \end{aligned}$$
 (6)

Here, β and $\overline{\beta}$ are coefficient vectors that are functions of the number of months between the observation date and the beginning of the forward period (n - m), and $Y_i(m)$ is simply the vector $X_i(m)$ augmented by a preceding unit element: $Y_i(m) = (1, X_i(m))$. The unit element allows the linear combination in the argument of the exponentials in Equation (6) to have a non-zero intercept.

In the current implementation of the forward intensity model in the CRI, the maximum forecast horizon is 60 months (5 years) and there are 12 input variables plus the intercept, so there are 60 sets of β and $\overline{\beta}$. While this is a large set of parameters, as will be seen in Subsection 1.2 and 1.3, the calibration is tractable because the default parameters can be calibrated separately from the other exit parameters, and the total number of parameters are greatly reduced after constraining the term-structure of the parameter estimates to be Nelson-Siegel functions.

Before expressing the probabilities in Equation (1) and (2) in terms of the forward intensities, a notation H is introduced for the forward intensities so that it becomes clear which parameters the forward intensity depends on:

$$H(\beta(n-m), X_i(m)) = \exp[\beta(n-m) \cdot Y_i(m)].$$
(7)

This is the forward default intensity. The corresponding notation for the forward other exit intensity is then just $H(\bar{\beta}(n-m), X_i(m))$. So, the probability in Equation (1) is expressed in terms of the forward intensities, using Equation (3) as the conditional

GLOBAL CREDIT REVIEW VOLUME 3 81

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default probability and Equation (5) as the conditional survival probability:

$$Prob_{t=m\Delta t}[\tau_{i} = n, \tau_{i} < \overline{\tau_{i}}]$$

$$= \{1 - \exp[-\Delta t H(\beta(n-1-m), X_{i}(m))]\}$$

$$\times \prod_{j=m}^{n-2} \exp\{-\Delta t [H(\beta(j-m), X_{i}(m))]$$

$$+ H(\overline{\beta}(j-m), X_{i}(m))]\}$$

$$= \{1 - \exp[-\Delta t H(\beta(n-m-1), X_{i}(m))]\}$$

$$\times \exp\left\{-\Delta t \sum_{j=m}^{n-2} [H(\beta(j-m), X_{i}(m))]$$

$$+ H(\overline{\beta}(j-m), X_{i}(m))]\right\}.$$
(8)

This probability will be relevant in the next part during the calibration. The cumulative default probability given in Equation (2) in terms of the forward intensities is then:

$$Prob_{t=m\Delta t}[m < \tau_{i} \le n, \tau_{i} < \overline{\tau_{i}}]$$

$$= \sum_{k=m}^{n-1} \left\{ \{1 - \exp[-\Delta t H(\beta(k-m), X_{i}(m))]\} \times \exp\left\{-\Delta t \sum_{j=m}^{k-1} [H(\beta(j-m), X_{i}(m)) + H(\overline{\beta}(j-m), X_{i}(m))]\right\} \right\}$$
(9)

This formula is used to compute the main output of the CRI: an individual firm's PD within various time horizons. The β and $\overline{\beta}$ parameters are obtained when the firm's economy is calibrated, and using those together with the firm's input variables yields the firm's PD.

1.2. Pseudo-Likelihood Function

The empirical data set used for calibration can be described as follows. For the economy as a whole, there are N end of month observations, indexed as

82 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

n = 1, ..., N. Of course, not all firms will have observations for each of the *N* months as they may start later than the start of the economy's data set or they may exit before the end of the economy's data set. There are a total of *I* firms in the economy, and they are indexed as i = 1, ..., I. As before, the input variables for the *i*th firm in the *n*th month is $X_i(n)$. The set of all observations for all firms is denoted by *X*.

In addition, the default times τ_i and non-default exit times $\bar{\tau}_i$ for the *i*th firm are known if the default or other exit occurs after time $t = \Delta t$ and at or before $t = N\Delta t$. The possible values for τ_i and $\bar{\tau}_i$ are integers between 2 and N, inclusive. If a firm exits before the month end, then the exit time is recorded as the first month end after the exit. If the firm does not exit before $t = N\Delta t$, then the convention can be used that both of these values are infinite. If the firm has a default type of exit within the data set, then $\bar{\tau}_i$ can be considered as infinite. If instead the firm has a non-default type of exit within the data set, then τ_i can be considered as infinite. The set of all default times and non-default exit times for all firms is denoted by τ and $\overline{\tau}$, respectively. The first month in which firm *i* has an observation is denoted by t_{0i} . Except for cases of missing data, these observations continue until the end of the data set if the firm never exits. If the firm does exit, the last needed input variable $X_i(n)$ is for $n = \min(\tau_i, \overline{\tau}_i) - 1$.

The calibration of the β and $\overline{\beta}$ parameters is done by maximizing a pseudo-likelihood function. The function to be maximized violates the standard assumptions of likelihood functions, but Appendix A in Duan *et al.* (2012) derives the large sample properties of the pseudo-likelihood function.

In formulating the pseudo-likelihood function, the assumption is made that the firms are conditionally independent from each other. In other words, correlations arise naturally from shared common factors W(n) and any correlations between different firms' firm-specific variables. With this assumption, the pseudo-likelihood function for the horizon of ℓ months, a set of parameters β and $\overline{\beta}$ and the data set $(\tau, \overline{\tau}, X)$ is:

$$\mathcal{L}_{\ell}(\beta,\overline{\beta};\tau,\overline{\tau},X) = \prod_{m=1}^{N-1} \prod_{i=1}^{I} P_{\min(N-m,\ell)}(\beta,\overline{\beta};\tau_i,\overline{\tau}_i,X_i(m)).$$
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Here, $P_{\min(N-m,\ell)}$ (β , $\bar{\beta}$; τ_i , $\bar{\tau}_i$, $X_i(m)$) is a probability for the *i*th firm, with the nature of the probability depending on what happens to the firm during the period from month *m* to month *m* + min (N - m, ℓ). This is defined as:

$$\begin{aligned} & P_{\ell}(\beta, \overline{\beta}; \tau_{i}, \overline{\tau_{i}}, X_{i}(m)) \\ &= \mathbf{1}_{\{t_{0i} \leq m, \min(\tau, \overline{\tau}) > m + \ell\}} \\ & \times \exp\left\{-\Delta t \sum_{j=0}^{\ell-1} [H(\beta(j), X_{i}(m)) + H(\overline{\beta}(j), X_{i}(m))]\right\} \\ & + \mathbf{1}_{\{t_{0i} \leq m, \tau_{i} \leq \overline{\tau_{i}}, \tau_{i} \leq m + \ell\}} \{\mathbf{1} - \exp[-\Delta t \ H(\beta(\tau_{i} - m - 1), X_{i}(m))]\} \\ & \times \exp\left\{-\Delta t \sum_{j=0}^{\tau_{i} - m - 2} [H(\beta(j), X_{i}(m)) + H(\overline{\beta}(j), X_{i}(m))]\right\} \\ & + \mathbf{1}_{\{t_{0i} \leq m, \overline{\tau_{i}} \leq \tau_{i}, \overline{\tau_{i}} \leq m + \ell\}} \{\mathbf{1} - \exp[-\Delta t \ H(\overline{\beta}(\overline{\tau_{i}} - m - 1), X_{i}(m))]\} \\ & \times \exp\left\{-\Delta t \ H(\beta(\tau_{i} - m - 1), X_{i}(m))] \\ & \times \exp\left\{-\Delta t \ \sum_{j=0}^{\overline{\tau_{i}} - m - 2} [H(\beta(j), X_{i}(m)) + H(\overline{\beta}(j), X_{i}(m))]\right\} \\ & + \mathbf{1}_{\{t_{0i} > m\}} + \mathbf{1}_{\{\min(\tau_{i}, \overline{\tau_{i}}) \leq m\}}. \end{aligned}$$

In words, if the *i*th firm survives from the observation time at month *m* for the full horizon ℓ until at least $m + \ell$, then the probability is the model-based survival probability for this period. This is the first term in Equation (11). The second term handles the cases where the firm has a default within the horizon, in which case the probability is the model-based probability of the firm defaulting at the month that it ends up defaulting, as given in Equation (8). The third term handles the cases where the firm has a nondefault exit within the horizon, in which case the probability is the model-based probability of the firm having a non-default type exit at the month that the exit actually does occur. The expression for this probability uses the conditional non-default type exit probability given in Equation (4). The final two terms handle the cases where the firm is not in the data set at month m — either the first observation for the firm is after m or the firm has already exited. A constant value is assigned in this case so that this firm will not affect the maximization at this time point.

The pseudo-likelihood function given in Equation (10) can be numerically maximized to give estimates

for the coefficients β and $\overline{\beta}$. Notice though that the sample observations for the pseudo-likelihood function are overlapping if the horizon is longer than one month. For example, when $\ell = 2$, default over the next two periods from month *m* is correlated to default over the next two periods from month *m* + 1 due to the common month in the two sample observations. However, in Appendix A of Duan *et al.* (2012), the maximum pseudo-likelihood estimator is shown to be consistent, in the sense that the estimators converge to the "true" parameter value in the large sample limit.

Notice though that each of the terms in Equation (11) can be written as a product of terms containing only β and terms containing only $\overline{\beta}$. This will allow separate maximizations with respect to β and with respect to $\overline{\beta}$, that is, the defaults and other exits.

The β and $\overline{\beta}$ specific versions of Equation (11) are:

$$\begin{aligned} P_{\ell}^{\beta}(\beta;\tau_{i},\overline{\tau_{i}},X_{i}(m)) \\ &= \mathbf{1}_{\{t_{0i} \leq m,\min(\tau,\overline{\tau_{i}}) > m+\ell\}} \exp\left\{-\Delta t \sum_{j=0}^{\ell-1} H(\beta(j),X_{i}(m))\right\} \\ &+ \mathbf{1}_{\{t_{0i} \leq m,\tau_{i} \leq \overline{\tau_{i}},\tau_{i} \leq m+\ell\}} \exp\left\{-\Delta t \sum_{j=0}^{\tau_{i}-m-2} H(\beta(j),X_{i}(m))\right\} \\ &\times \{\mathbf{1} - \exp[-\Delta t H(\beta(\tau_{i}-m-1),X_{i}(m))]\} \\ &+ \mathbf{1}_{\{t_{0i} \leq m,\overline{\tau_{i}} \leq \tau_{i},\overline{\tau_{i}} \leq m+\ell\}} \exp\left\{-\Delta t \sum_{j=0}^{\overline{\tau_{i}}-m-2} H(\beta(j),X_{i}(m))\right\} \\ &\times \exp[-\Delta t H(\beta(\tau_{i}-m-1),X_{i}(m))] \\ &+ \mathbf{1}_{\{t_{0i} > m\}} + \mathbf{1}_{\{\min(\tau_{i},\overline{\tau_{i}}) \leq m\}}, \\ P_{\ell}^{\overline{\beta}}(\overline{\beta};\tau_{i},\overline{\tau_{i}},X_{i}(m)) \\ &= \mathbf{1}_{\{t_{0i} \leq m,\min(\tau,\overline{\tau}) > m+\ell\}} \exp\left\{-\Delta t \sum_{j=0}^{\ell-1} H(\overline{\beta}(j),X_{i}(m))\right\} \\ &+ \mathbf{1}_{\{t_{0i} \leq m,\overline{\tau_{i}} \leq \overline{\tau_{i}},\overline{\tau_{i}} \leq m+\ell\}} \exp\left\{-\Delta t \sum_{j=0}^{\tau_{i}-m-2} H(\overline{\beta}(j),X_{i}(m))\right\} \\ &+ \mathbf{1}_{\{t_{0i} \leq m,\overline{\tau_{i}} \leq \tau_{i},\overline{\tau_{i}} \leq m+\ell\}} \exp\left\{-\Delta t \sum_{j=0}^{\tau_{i}-m-2} H(\overline{\beta}(j),X_{i}(m))\right\} \\ &+ \mathbf{1}_{\{t_{0i} \leq m,\overline{\tau_{i}} \leq \tau_{i},\overline{\tau_{i}} \leq m+\ell\}} \exp\left\{-\Delta t \sum_{j=0}^{\tau_{i}-m-2} H(\overline{\beta}(j),X_{i}(m))\right\} \\ &\times \{\mathbf{1} - \exp[-\Delta t H(\overline{\beta}(\overline{\tau_{i}}-m-1),X_{i}(m))]\} \\ &+ \mathbf{1}_{\{t_{0i} > m\}} + \mathbf{1}_{\{\min(\tau_{i},\overline{\tau_{i}}) \leq m\}}. \end{aligned} \right$$

GLOBAL CREDIT REVIEW VOLUME 3 83

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Then, the β and $\overline{\beta}$ specific versions of the pseudolikelihood function are given by:

$$\mathcal{L}_{\ell}^{\beta}(\beta;\tau,\overline{\tau},X) = \prod_{m=1}^{N-\ell} \prod_{i=1}^{I} P_{\ell}^{\beta}(\beta;\tau_{i},\overline{\tau}_{i},X_{i}(m))$$

$$\mathcal{L}_{\ell}^{\overline{\beta}}(\overline{\beta};\tau,\overline{\tau},X) = \prod_{m=1}^{N-\ell} \prod_{i=1}^{I} P_{\ell}^{\overline{\beta}}(\overline{\beta};\tau_{i},\overline{\tau}_{i},X_{i}(m)).$$
(13)

With the definitions given in Equation (12) and (13), it can be seen that:

$$\mathcal{L}_{\ell}(\beta,\overline{\beta};\tau,\overline{\tau},X) = \mathcal{L}_{\ell}^{\beta}(\beta;\tau,\overline{\tau},X)\mathcal{L}_{\ell}^{\overline{\beta}}(\overline{\beta};\tau,\overline{\tau},X).$$
(14)

Thus, $\mathcal{L}_{\ell}^{\beta}$ and $\mathcal{L}_{\ell}^{\overline{\beta}}$ can be separately maximized to find their respective parameters. Subsection 1.3 will further explain how the optimal parameters can be estimated.

1.3. Parameter Estimation

Previously, the CRI system produced default predictions up to horizons of two years (RMI, 2012). An extension of the forecast horizon has been implemented as of the PD released on 1 April 2013. With this update, horizons of up to five years are now being computed. Technically speaking, horizons of arbitrary length can be calculated.

This extension to a five-year horizon is done by constraining the term-structure of the parameter estimates to be Nelson-Siegel (Nelson and Siegel, 1987; hereafter NS) functions of the forward-starting time. Horizon-specific parameters β , $\bar{\beta}$ can be obtained from the continuous NS function by using the forward prediction horizon as an input. The term-structures are further constrained so that the effect of risk factors on the forward intensity goes to zero as the horizon increases. This allows tractable and parsimonious extrapolations for horizons beyond five years.

The parameter estimation for the NS functions is based on a new numerical method (a pseudo-Bayesian SMC technique) developed in a working paper by Duan and Fulop (2013). The remainder of this section details the new parameter estimation. 1.3.1 describes the parameterization of the parameters by NS functions, 1.3.2 gives an overview of the SMC method that is used to estimate the NS functions, 1.3.3 details the

84 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

calculation of the confidence intervals for the parameter estimation, and 1.3.4 describes how the parameters can be re-estimated given new data or updates of old data.

1.3.1. Smoothed parameters

Duan *et al.* (2012) formulate the forward intensity model in which the forward default intensity for a firm is a function of a number of covariates. The forward default intensities for different forward starting periods are computed using different sets of parameters.

In Duan *et al.* (2012), the sets of parameters are estimated separately for each forward starting time. Parameters at different forward starting times that are associated with each covariate can be approximated by a function of the forward starting time using NS type term structure functions. Duan *et al.* (2012) show that this approximation by NS functions does not negatively affect prediction performance. The RMI implementation follows Duan and Fulop (2013) to impose the functional restriction during the estimation as opposed to the method used in Duan *et al.* (2012) of fitting the curve after parameter estimates have been obtained. This is done for two reasons.

First, it will significantly reduce the number of parameters. For example, using 12 covariates for forward default intensities up to 60 months would require a joint estimation of $13 \times 60 = 780$ parameters. Here, 13 comes from adding an intercept to the intensity function with 12 covariates. If the coefficients corresponding to each covariate are represented by the NS function of 4 parameters, there will be at most $13 \times 4 = 52$ parameters. In fact, there will be fewer parameters as some of the NS parameters will be constrained to zero.

Second, the NS function will allow extrapolation. For example, the 13 NS functions estimated with predictions up to 60 months can be used for prediction, say, over 72 months.

The NS function with four free parameters is:

$$r(t;\varrho_0,\varrho_1,\varrho_2,d) = \varrho_0 + \varrho_1 \frac{1 - \exp(-t/d)}{t/d} + \varrho_2 \left[\frac{1 - \exp(-t/d)}{t/d} - \exp(-t/d) \right],$$
(15)

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where *t* is the forecast horizon (measured in years). In the RMI implementation, the horizon is 60 months (5 years) so that *t* ranges from 0 to 59/12. Once the four NS parameters are estimated, individual horizon-specific parameters β , $\bar{\beta}$ are obtained from the NS function *r* using the forecast horizon as input to the NS function. In our current implementation with forecast horizons extending to 60 months (5 years), 120 sets of month-specific β and $\bar{\beta}$ are obtained. For all covariates, the restriction d > 0 is imposed so that the functions converge to a value for large *t*. This formulation will be used for forward intensities for both defaults and other types of exit.

For the coefficients of all stochastic covariates, the long-run level ρ_0 is restricted to zero, because the current value of a stochastic covariate should be uninformative of default or other exits when the forward starting time goes to infinity. In other words, the coefficient of such a stochastic covariate should approach zero when t goes to infinity.

The intercept of the forward intensity function is of course non-stochastic. Thus, ρ_0 can have non-zero values for the intercept. With these restrictions on the NS parameters, take the example of 12 covariates, there will be a total of $12 \times 3 + 1 \times 4 = 40$ parameters.

In the RMI implementation, the NS function is further constrained to be non-positive for certain covariates: DTD level and trend, liquidity level and trend, and profitability level and trend. Refer to Section 2 for descriptions of these covariates.

1.3.2. Parameter estimation by SMC

Reliably estimating a system involving 40 parameters presents a numerical challenge. Moreover, the number of parameters can be greater than 40 if there are more than 12 covariates. The RMI implementation follows Duan and Fulop (2013) who use the SMC pseudo-Bayesian method for estimation and self-normalized statistics for inference.

Due to decomposability, the analysis can be performed separately on the forward default and other exit intensities. The data in the RMI implementation are refreshed with monthly frequency, and the sample likelihood used in estimation relies on default predictions running from 1 month to 60 months with a one month increment. Naturally, default prediction is subjected to data availability. Towards the end of the period with available data, the prediction horizon naturally decreases and stops at one-month predictions.

The following exposition closely follows the appendix in Duan and Fulop (2013). It is important to note that the RMI implementation uses the model described in Duan and Fulop (2013), which does not contain any latent frailty or partial conditioning variable, and hence is technically much simpler in parameter estimation. For example, there is no nonlinear filtering problem.

According to the current modeling framework, where for a particular economy there are *N* end of month observations, the input variables of the *i*th firm in the *m*th month is given by $X_i(m)$. Let θ denote a set of NS parameters and ℓ denote the forecast horizon ($\ell = 60$). Then the pseudo-likelihood function at step *m*, denoted by $\mathcal{L}_{m,\min(N-m,\ell)}(\theta)$, takes the form:

$$\mathcal{L}_{m,\min(N-m,\ell)}(\theta) = \prod_{i=1}^{I} P_{\min(N-m,\ell)}$$
(16)
$$(\beta(\theta), \overline{\beta}(\theta); \tau_i, \overline{\tau}_i, X_i(m)),$$

where *I* is the number of firms, $\beta(\theta)$ and $\overline{\beta}$ (θ) are the coefficient vectors from Equation (6) generated from the NS functions with parameter θ . One may notice that $\mathcal{L}_{m,\min(N-m,\ell)}(\theta)$ is one of the terms in the outer-most product in Equation (10).

Let $\pi(\theta)$ denote the prior. Following the notation from Section 1.1, consider the following pseudoposterior distribution at time *n* after one makes the ℓ -period prediction:

$$\gamma_n(\theta) \propto \prod_{m=1}^{n-1} \mathcal{L}_{m,\min(N-m,\ell)}(\theta) \pi(\theta), \text{ for } n = 2, \dots, N.$$
 (17)

One can apply the sequential batch-resampling routine of Chopin (2012) together with tempering steps as in Del Moral *et al.* (2006) to advance the system. For each *n*, this procedure yields a weighted sample of *K* particles, $(\theta^{(k,n)}, w^{(k,n)})$ with k = 1, ..., K, whose empirical distribution function will converge to $\gamma_n(\theta)$ as *K* increases. In the following paragraphs, the superscript *k* denotes the particle index. Note that in the RMI implementation, K = 1000. ()

Initialization: Draw an initial random sample from the prior: $(\theta^{(k,0)} \sim \pi(\theta), w^{(k,0)} = 1/K)$. Here, the only role of the prior $\pi(\theta)$, is to provide the initial particle cloud from which the algorithm can start. Of course, the support of $\pi(\theta)$ must contain the true parameter value θ_0 . In the RMI implementation, normal/truncated normal priors are used. Truncation applies in order to impose the restriction d > 0. To obtain the means of the priors for the SMC method, a least square fit of the MLE parameter estimates of the NS function is conducted. The standard deviations of the priors are set to 5.

Recursions and defining the tempering sequence: Assume there is a particle cloud $(\theta^{(k,n)}, w^{(k,n)})$ whose empirical distribution represents $\gamma_{\mu}(\theta)$. Then, a cloud representing $\gamma_{n+1}(\theta)$ will be reached by combining importance sampling and the Markov Chain Monte Carlo (MCMC) steps. Sometimes moving directly from $\gamma_n(\theta)$ to $\gamma_{n+1}(\theta)$ is too ambitious as the two distributions are too far from each other. This will be reflected in highly variable importance weights if one resorts to direct importance sampling. Hence, following Duan and Fulop (2013) which in turn followed Del Moral et al. (2006), a tempered bridge is built between the two densities and the particles are evolved through the resulting sequence of densities. In particular, assume that at n + 1, there are P_{n+1} intermediate densities:

$$\overline{\gamma}_{n+1,p}(\theta) \propto \gamma_n(\theta) \mathcal{L}_{n,\min(N-n,\ell)}^{\xi p}(\theta),$$
for $p = 1, \dots, P_{n+1}$.
(18)

This construction defines an appropriate bridge: $\xi_0 = 0$ so that $\bar{\gamma}_{n+1,0}(\theta) = \gamma_n(\theta)$, and $\xi_{P_{n+1}} = 1$ so that $\bar{\gamma}_{n+1,P_{n+1}}(\theta) = \gamma_{n+1}(\theta)$. For *p* between 0 and P_{n+1} , ξ_p is chosen from a grid of points to evenly distribute the weights, as decribed below. A particle cloud representing $\bar{\gamma}_{n+1,0}(\theta)$ can be initialized as $(\bar{\theta}^{(k,n+1,0)}, \bar{w}^{(k,n+1,0)}) = (\theta^{(k,n)}, w^{(k,n)})$. Then, for $p = 1, \dots, P_{n+1}$ the sequence proceeds as follows:

• *Reweighting Step*: In order to arrive at a representation of $\bar{\gamma}_{n+1,p}(\theta)$, the particles representing $\bar{\gamma}_{n+1,p-1}(\theta)$ and the importance sampling principle can be used. This leads to:

$$\overline{\theta}^{(k,n+1,p)} = \overline{\theta}^{(k,n+1,p-1)},\tag{19}$$

$$\overline{w}^{(k,n+1,p)} = \overline{w}^{(k,n+1,p-1)} \times \frac{\overline{\gamma}_{n+1,p}(\theta^{(k,n+1,p)})}{\overline{\gamma}_{n+1,p-1}(\overline{\theta}^{(k,n+1,p)})}
= \overline{w}^{(k,n+1,p-1)} \times \mathcal{L}_{n,\min(N-n,\ell)}^{\xi_p - \xi_{p-1}}
(\overline{\theta}^{(k,n+1,p)}).$$
(20)

To avoid particle impoverishment in sequential importance sampling where most of the weight is concentrated in a small number of particles, a resample-move step is run, which is triggered whenever a measure of particle diversity — the efficient sample size (ESS) defined as:

$$ESS = \frac{\left(\sum_{k=1}^{N} \overline{w}^{(k,n+1,p)}\right)^{2}}{\sum_{k=1}^{N} (\overline{w}^{(k,n+1,p)})^{2}},$$
(21)

falls below some preset value *B*. Here, resampling directs the particle cloud towards more likely areas of the sampling space, while the move step enriches particle diversity. In the RMI implementation, *B* is set to 50%. Thus, if ESS < 50%, the following resampling and move steps are performed.

Resampling Step: The particles are resampled proportional to their weights. If *I*^(k,n+1,p) ∈ (1,...,*K*) are particle indices sampled proportional to w
^(k,n+1,p), the equally weighted particles are obtained as:

$$\overline{\theta}^{(k,n+1,p)} = \overline{\theta}^{(I^{(k,n+1,p)},n+1,p)},\tag{22}$$

$$\overline{w}^{(k,n+1,p)} = \frac{1}{K}.$$
(23)

- Move Step: Each particle is passed through a Markov kernel $K_{n+1,p}$ ($\theta^{(k,n+1,p)}$,·) that leaves $\gamma_{n+1,p}(\theta)$ invariant, typically a Metropolis-Hastings kernel:
 - 1. Propose $\theta^{*(k)} \sim Q_{n+1,p} \left(\cdot \left| \overline{\theta}^{(k,n+1,p)} \right) \right)$.
 - 2. Compute the acceptance weight α , where:

$$\alpha = \min\left(1, \frac{\overline{\gamma}_{n+1,p}(\theta^{*(k)})\mathcal{Q}_{n+1,p}\left(\overline{\theta}^{(k,n+1,p)} \middle| \theta^{*(k)}\right)}{\overline{\gamma}_{n+1,p}(\overline{\theta}^{(k,n+1,p)})\mathcal{Q}_{n+1,p}\left(\theta^{*(k)} \middle| \overline{\theta}^{(k,n+1,p)}\right)}\right).$$
(24)

⁸⁶ NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

3. With probability α , set $\bar{\theta}^{(k,n+1,p)} = \theta^{*(k)}$, otherwise keep the old particle.

This step will enrich the support of the particle cloud while conserving its distribution. If the particle set is a poor representation of the target distribution, the move step can also help adjust the location of the support. Crucially, given the importance of the sampling setup, the proposal distribution $Q_{n+1,p}(\cdot | \overline{\theta}^{(k,n+1,p)})$ can be adapted using the existing particle cloud.

In the RMI implementation, block independent normal distribution proposals using the means and the standard deviations implied by the particle set are fitted to the particle cloud before the move. Three (or four) NS parameters corresponding to each covariate form one block. To ensure that the NS parameter d remains positive, any block with a non-positive value for d is discarded. To ensure the smoothness of the term structure of the forward intensity parameters, any block that does not produce an increasing or decreasing structure of the NS function for the first five months is also discarded. Once some block is discarded, the particle is regenerated until it meets the requirements. Note that the likelihood ratio in the Metropolis-Hastings algorithm is not affected by this because the truncated normal creates a common adjustment term in both the numerator and denominator.

As mentioned previously, the coefficients for some covariates are required to be non-positive over all forward starting times. This is achieved by checking whether the NS curve at a particular set of three (or four) parameters meets the condition. If not, the parameter set will be discarded.

To improve the support of the particle cloud, one can execute multiple such Metropolis-Hastings steps each time. In the RMI implementation, such Metropolis-Hastings steps are consecutively performed in each resamplingmove step until the number of unique particles exceeds K/2.

When $p = P_{n+1}$ is reached, a representation of $\gamma_{n+1}(\theta)$ is:

$$(\theta^{(k,n+1)}, w^{(k,n+1)}) = (\overline{\theta}^{(k,n+1,P_{n+1})}, \overline{w}^{(k,n+1,P_{n+1})}).$$
(25)

Following Duan and Fulop (2013), the tempering sequence ξ_p is automatically set to ensure that the efficient sample size stays close to 50%. This is done by a grid search, where the ESS is evaluated at a grid of candidate ξ_p and the one that produces the closest ESS to 50% is chosen.

After the recursion procedure (i.e., ξ_p reaches 1), additional moves using the means implied by the particle set but with all standard deviations increased by a factor of 30% are further performed to enrich the support and adjust the location of the particle set. The number of such moves is set to 20 for the first time point and exponentially declines to 3 mid-way to the sample period and stays at 3 for the remainder. After that, if the number of unique particles is still below K/2, more moves using the means and the standard deviations implied by the particle set (without expansions) are consecutively performed until the particle set meets the requirement. (This case could only happen when ESS $\geq B$ for $\xi_p = 1$.)

1.3.3. Statistical inference

The full sample size has *N* time series data points but one can only make default prediction at N-1 time points; for example, at time point 2, the data is only available for making one-period default prediction at time point 1. Denote the pseudo-posterior mean of the parameter of the whole sample by $\overline{\beta}$ and for n = 2, ..., N,

$$\hat{\theta}_n = \frac{1}{\sum_{k=1}^{K} w^{(k,n)}} \sum_{k=1}^{K} w^{(k,n)} \theta^{(k,n)}.$$
(26)

Note that $(\overline{\theta}^{(k,n+1,0)}, \overline{w}^{(k,n+1,0)}) = (\theta^{(k,n)}, w^{(k,n)})$. is not a true posterior because the likelihood function in equation (17) is not a true likelihood function. Thus, it cannot directly provide valid Bayesian inference. But following Duan and Fulop (2013) — which is in turn based on Shao's self-normalized statistic (Shao, 2010) — inference can be performed using the *t*-like statistic. To test, for example, the hypothesis of the *k*th element of $\overline{\theta}^{(k,n+1,p)} = \overline{\theta}^{(I^{(k,n+1,p)},n+1,p)}$, denoted by $\overline{w}^{(k,n+1,p)} = \frac{1}{K}$, equal to *a*, one has:

$$t^* = \frac{\sqrt{N-1}\left(\hat{\theta}_N^{(k)} - a\right)}{\sqrt{\hat{\delta}_{k,N}}} \xrightarrow{d} \frac{W(1)}{\left[\int_0^1 (W(r) - rW(1))^2 dr\right]^{1/2}}, \quad (27)$$

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GLOBAL CREDIT REVIEW VOLUME 3 87

where W(r) is a Wiener process, is the *k*th diagonal element of $\theta_0^{(k)}$, and:

$$\hat{C}_{N} = \frac{1}{(N-1)^{2}} \sum_{n=2}^{N} n^{2} (\hat{\theta}_{n} - \hat{\theta}_{N}) (\hat{\theta}_{n} - \hat{\theta}_{N})'.$$
(28)

The right-hand-side random variable for t^* does not have a known distribution, but can be easily simulated. Kiefer *et al.* (2000) reported that the 95% quantile is 5.374 and the 97.5% quantile is 6.811. These values can also be used to set up confidence intervals.

1.3.4. Periodic updating

In reality, portfolio credit risk models need to be updated periodically as new data arrive and/or old data are revised. With one new month of data, this means that the final date index N is increased to N + 1. A particular strength of Duan and Fulop's (2013) methodology is that the estimation routine does not need to be re-initialized from the prior as the pseudo-posterior using data up to $N\Delta t$ will provide a much better proposal distribution.

Let the pseudo-posterior at time N (based on the old data set available at time N) be denoted by:

$$\gamma_N^{(N)}(\theta) \propto \prod_{m=1}^{N-1} \mathcal{L}_{m,\min(N-m,\ell)}^{(N)}(\theta) \pi(\theta), \qquad (29)$$

and the pseudo-posterior at time N + 1 (based on the new data set available at time N + 1) by:

$$\gamma_{N+1}^{(N+1)}(\boldsymbol{\theta}) \propto \prod_{m=1}^{N} \mathcal{L}_{m,\min((N+1)-m,\ell)}^{(N+1)}(\boldsymbol{\theta}) \pi(\boldsymbol{\theta}).$$
(30)

The superscript is introduced to differentiate the data set available at time *N* and *N* + 1, respectively. It is important to note that $\mathcal{L}_{m,k}^{(N+1)}(\theta) \neq \mathcal{L}_{m,k}^{(N)}(\theta)$ can be caused by revisions to the old data set. More importantly, there is a generic difference between the pseudo-posterior distribution up to time *N* under the new data set and the corresponding quantity under the old data set specifically due to multiperiod prediction; that is, $\gamma_{N+1}^{(N)}(\theta) \neq \gamma_N^{(N)}(\theta)$ even without any data revisions to the period covered by the old data

88 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

set. To put it concretely, using the new data set and at, say, one period before the last (i.e., time N - 1), one can make default predictions up to two periods, whereas at the same time point, it was only possible to make one-period predictions under the old data set because there were no data beyond time N. Adjustments to the weights are thus necessary to reflect the change in data set before making any sequential updates.

There are several possible ways of advancing the system. The RMI implementation decomposes the move into two steps. First, we take care of data revision up to time N and then act as if we were making predictions with data only up to time N. Doing it this way is meant to maintain the same default prediction setting; that is, for example, only making one-period default prediction at time N - 1 even though the new data set permits predictions up to two periods. Thus, we introduce:

$$\gamma_N^{(N+1,N)}(\theta) \propto \prod_{m=1}^{N-1} \mathcal{L}_{m,\min(N-m,\ell)}^{(N+1)}(\theta) \pi(\theta)$$
(31)

to denote this pseudo-posterior when the superscript (N + 1, N) stands for the updated data set available at time N + 1 but making default predictions as if the data were only available up to time N.

From the previous run up to time *N*, one already has a weighted set of particles $(\theta^{(k,N)}, w^{(k,N)})$ representing the pseudo-posterior distribution $\gamma_N^{(N)}(\theta)$. Next, perform a reweighting by:

$$\boldsymbol{\theta}^{*(k,N)} = \boldsymbol{\theta}^{(k,N)},\tag{32}$$

$$w^{*(k,N)} = w^{(k,N)} \times \frac{\gamma_N^{(N+1,N)}(\theta^{(k,N)})}{\gamma_N^{(N)}(\theta^{(k,N)})}.$$
 (33)

Since the denominator is available from the previous run, one only needs to compute the numerator using the new data set up to time *N*. Then, the weighted set $(\theta^{*(k,N)}, w^{*(k,N)})$ represents the revised pseudo-posterior distribution at time *N*, i.e., $\gamma_N^{(N+1,N)}(\theta)$, specifically to account for data revisions. From this point onward, one can apply the same recursive procedure described

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in Section 1.3.2, starting from equation (18), to complete the updating task.

Reweighting may substantially alter the ESS of the particle set due to a large volume of data changes. If the reweighting leads to a satisfactory ESS, i.e., ESS $\geq B$, advancing to N + 1 continues as usual. Otherwise, the weighted sample will be discarded to prevent the support from degeneration. One can return to the particle set before reweighting and perform resampling to create an equally-weighted particle set. Then, make the Metropolis-Hastings moves by targeting $\gamma_N^{(N+1,N)}(\theta)$ using the Gaussian-type sampler described earlier and starting with the mean and variance implied by the resampled particle set. One should make these Metropolis-Hastings moves until the particle set reaches a desirable level of distinctiveness, and perhaps with a preset minimum number of moves to ensure that the resulting particle set is close enough to the target distribution. In the RMI implementation, the number of moves is set to be 20.

Furthermore, one can update all self-normalized statistics in the way as described earlier to reflect the one additional pseudo-posterior means to the sequence.

The initial parameter estimation is carried out for all calibration groups using the data up to the end of January 2013. Relevant quantities (parameter estimates, the 1000 parameter particles and corresponding weights and sample likelihoods) are saved for periodic updating for all future months. Additional implementation details on the calibration are given in Section 3.

II. INPUT VARIABLES AND DATA

Subsection 2.1 describes the input variables used in the quantitative model. Currently, the same set of input variables is common to all of the economies under the CRI's coverage. Future enhancements to the CRI system will allow different input variables for different economies. The effect of each of the variables on the PD output will be discussed in the empirical analysis of Section 4.

Subsection 2.2 gives the data sources and relevant details of the data sources. There are two categories

of data sources: current and historical. Data sources used for current data need to be updated in a timely manner so that daily updates of PD forecasts are meaningful. They also need to be comprehensive in their current coverage of firms. Data sources that are comprehensive for current data may not necessarily have comprehensive historical coverage for different economies. Thus, other data sources are merged in order to obtain comprehensive coverage of historical and current data.

Subsection 2.3 indicates the fields from the data sources that are used to construct the input variables. For some of the fields, proxies need to be used for a firm if the preferred field is not available for that firm.

Subsection 2.4 discusses the definition and sources of defaults and of other exits used in the CRI.

2.1. Input Variables

Following the notation that was introduced in Section 1, firm *i*'s input variables at time $t = n\Delta t$ are represented by the vector $X_i(n) = (W(n), U_i(n))$ consisting of a vector W(n) that is common to all firms in the same economy, and a firm-specific vector $U_i(n)$ which is observable from the date the firm's first financial statement is released, until the month end before the month in which the firm exits, if it does exit.

In Duan *et al.* (2012), different variables that are commonly used in the literature were tested as candidates for the elements of W(n) and $U_i(n)$. Two common variables and ten firm-specific variables, as described below, were selected as having the greatest predictive power for corporate defaults in the United States. In the current stage of development, this same set of twelve input variables is used for all economies. Future development will include variable selection for firms in different economies.

Common variables

The vector W(n) contains two elements, which are:

- 1. Stock index return: the trailing one-year simple return on a major stock index of the economy;
- 2. Interest rate: a representative 3-month short term interest rate.

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• Firm-specific variables

The ten firm-specific input variables are transformations of measures of six different firm characteristics. The six firm characteristics are:

- 1. volatility-adjusted leverage;
- 2. liquidity;
- 3. profitability;
- 4. relative size;
- 5. market misvaluation/future growth opportunities; and
- 6. idiosyncratic volatility.

Volatility-adjusted leverage is measured as the DTD in a Merton-type model. The calculation of DTD used by the CRI allows a meaningful DTD for financial firms, a critical sector that must be excluded from most DTD computations. This calculation is detailed in Section 3.

Liquidity is measured as a ratio of cash and short term investments to total assets. Profitability is measured as a ratio of net income to total assets. Relative size is measured as the logarithm of the ratio of market capitalization to the economy's median market capitalization.

Duan *et al.* (2012) transformed these first four characteristics into level and trend versions of the measures. For each of these characteristics, the level is computed as the one-year average of the measure, and the trend is computed as the current value of the measure minus the one-year average of the measure. The level and trend of a measure has seldom been used in the academic or industry literature for default prediction, and Duan *et al.* (2012) found that using the level and trend significantly improves the predictive power of the model for short-term horizons.

To understand the intuition behind using the level and trend of a measure as opposed to using just the current value, consider the case of two firms with the same current value for all measures. If the level and trend transformations were not performed, only the current values would be used and the two firms would have identical PD. Suppose that for the first firm the DTD had reached its current level from a high level, and for the second firm the DTD had reached its current level from a lower level (see Figure 2). The first

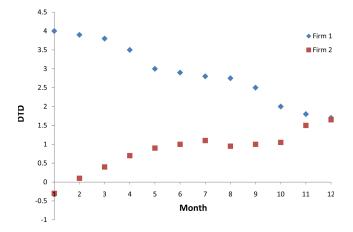


Figure 2. Two firms with all current values equal to each other, but DTD trending in the opposite direction.

firm's leverage is increasing (worsening) and the second firm's leverage is decreasing (improving). If there is a momentum effect in DTD, then firm 1 should have a higher PD than firm 2.

Duan *et al.* (2012) found evidence of the momentum effect in DTD, liquidity, profitability and size. For the other two firm characteristics, applying the level and trend transformation did not improve the predictive power of the model.

One of the remaining two firm characteristics is the market mis-valuation/future growth opportunities characteristic, which is taken as the market-to-book asset ratio and measured as a ratio of market capitalization and total liabilities to total assets. One can see whether the market mis-valuation effect or the future growth opportunities effect dominates this measure by looking at whether the parameter for this variable is positive or negative. This will be further discussed in the empirical analysis of Section 4.

The last firm characteristic is the idiosyncratic volatility which is taken as SIGMA, following Shumway (2001). SIGMA is computed by regressing the daily returns of the firm's market capitalization against the daily returns of the economy's stock index, for the previous 250 days. SIGMA is defined to be the standard deviation of the residuals of this regression. Using daily returns is to ensure that SIGMA provides an accurate and timely measure of idiosyncratic risk of individual companies. Shumway (2001) reasons that SIGMA should be logically related to bankruptcy since firms with more variable cash flows and therefore more variable stock returns relative to a market index are likely to have a higher probability of bankruptcy.

Finally, the vector $U_i(n)$ contains ten elements, consisting of:

- 1. Level of DTD.
- 2. Trend of DTD.
- 3. Level of (Cash + Short term investments)/Total assets, abbreviated as CASH/TA.
- 4. Trend of CASH/TA.
- 5. Level of Net income/Total assets, abbreviated as NI/TA.
- 6. Trend of NI/TA.
- 7. Level of log (Firm market capitalization/ Economy's median market capitalization), abbreviated as SIZE.
- 8. Trend of SIZE.
- 9. Current value of (Market capitalization + Total liabilities)/Total asset, abbreviated as M/B.
- 10. Current value of SIGMA.

The data fields that are needed to compute DTD and short term investments are described in Subsection 2.3. The remaining data fields required are straightforward and standard. The computation for DTD is explained in Section 3.

2.2. Data Sources

There are two data sources that are used for the daily PD forecast updates: Thomson Reuters Datastream and the Bloomberg Data License Back Office Product. Many of the common factors such as short term interest rates and macroeconomic data are retrieved from Datastream.

Firm-specific data comes from Bloomberg's Back Office Product which delivers daily update files by region via FTP after the respective market closes. All relevant data is extracted from the FTP files and uploaded into the CRI database for storage. From this, the necessary fields are extracted and joined with previous months of data.

The Back Office Product includes daily market capitalization data based on closing share prices and also includes new financial statements as companies release them. Firms will often have multiple versions of financial statements within the same period, with different accounting standards, filing statuses (most recent, preliminary, original, reclassified or restated), currencies or consolidated/unconsolidated indicators. A major challenge lies in prioritizing these financial statements to decide which data should be used. The priority rules are described in Section 3.

The firm coverage of the Back Office Product is of sufficient quality that around 34,000 firms can be updated on a daily basis in the 106 economies under the CRI's coverage. While the current coverage is quite comprehensive, historical data from the Back Office Product can be sparse for certain economies. For this reason, various other databases are merged in order to fill out the historical data. The other databases used for historical data are: a database from the Taiwan Economics Journal (TEJ) for Taiwanese firms; a database provided by Korea University for South Korean firms; and data from Prowess for Indian firms.

With all of the databases merged together and for the 106 economies under CRI's coverage, over 60,000 exchange listed firms are in the CRI database. This includes over 30,000 firms that have been delisted at some point in time. The historical coverage of the firm data goes back to the early 1990s. In order to be included in our coverage, a company needs to have common equity traded on a stock exchange. Of these 106 economies, 71 economies have their own stock exchange (see Table A.2). For the other 35 economies under the RMI coverage, we cover companies domiciled in the economy that are quoted on a foreign exchange, either because those economies do not have a stock exchange or because data issues are preventing us from including the companies listed on the local exchange.

2.3. Constructing Input Variables

The chosen stock indices and short term interest rates for the 71 economies with their own stock exchange under the CRI's current coverage are listed in Table A.5 and Table A.6, respectively. All economies are listed by their three letter ISO code given in Table A.4.

Most of the firm-specific variables can be readily constructed from standard fields from firms' financial

GLOBAL CREDIT REVIEW VOLUME 3 91

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statements in addition to daily market capitalization values. The only two exceptions are the DTD and the liquidity measure.

The calculation for DTD is explained in Section 3. In the calculation, several variables are required. One variable is a proxy for a one-year risk-free interest rate, and the choices for each of the 71 economies are listed in Table A.7. Total assets, long-term borrowing and total liabilities are also required, but can be obtained from standard financial statement fields easily.

Total current liabilities are also required, and due to the relatively large number of firms that are missing this value, proxies have to be found. The preferred Bloomberg field for this is BS_CUR_LIAB. If this is missing, then the sum of BS_ST_BORROW, BS_ OTHER_ST_LIAB and BS_CUST_ACCPT_LIAB_ CUSTDY_SEC (customers' acceptance and liabilities/ custody securities) is used. If one or two of these are missing, zero is inserted for those fields, but at least one field is required.

The liquidity measure requires different fields for financial and non-financial firms. For non-financial firms, the numerator of the ratio (Cash + Short term investments) is taken as the sum of BS_CASH_NEAR_CASH_ITEM and BS_MKT_SEC_OTHER_ST_INVEST (marketable securities and other short term investments). If BS_MKT_SEC_OTHER_ST_INVEST is missing, substitute zero (but BS_CASH_NEAR_CASH_ITEM is required).

It was found that this sum frequently overstated the liquidity for financial firms. In place of BS_ MKT_SEC_OTHER_ST_INVEST, financial firms use the sum of ARD_SEC_PURC_UNDER_AGR_TO_ RESELL (securities purchased under agreement to resell), ARD_ST_INVEST and BS_INTERBANK_ ASSET. If one or two of these are missing, zero is inserted for those fields, but at least one field is required. The "ARD" prefix indicates that these are "as reported" numbers directly from the financial statements. As such, for some firms these fields may need to be adjusted to the same units before adding them to other fields.

To summarize, the firm-specific variables include: DTD, Cash/TA, NI/TA, SIZE, M/B, and SIGMA, and the statistics grouped by economy are listed in Table A.8.

2.4. Data for Defaults

The CRI database contains credit events of over 4,000 firms from 1990 to the present. The default events come from numerous sources, including Bloomberg, Compustat, CRSP, Moody's reports, TEJ, exchange websites and news sources.

The default events that are recognized by the CRI can be classified under one of the following events:

- Bankruptcy filing, receivership, administration, liquidation or any other legal impasse to the timely settlement of interest and/or principal payments;
- A missed or delayed payment of interest and/or principal, excluding delayed payments made within a grace period;
- 3. Debt restructuring/distressed exchange, in which debt holders are offered a new security or package of securities that result in a diminished financial obligation (e.g., a conversion of debt to equity, debt with lower coupon or par amount, debt with lower seniority, debt with longer maturity).

The more precise sub-categories of default corporate actions are listed in Table A.9.

Delisting due to other reasons such as failure to meet listing requirements, inactive stock prices or M&A are counted as "other exits" and are not considered as default. However, firms that are delisted from an exchange and then experience a default event within 365 calendar days of the delisting will have an exit event reclassified as credit default. Technical defaults such as covenant violations are not included in our definition of default. The exit events that are not considered as defaults in the CRI system are listed in Table A.10.

In addition to the aforementioned events, there are still cases that require special attention and will be assessed on a case-by-case basis, e.g., subsidiary default. As a general rule, the CRI does not consider related party-default (e.g., subsidiary bankruptcy) as a default event. However, when a non-operating holding parent company relies heavily on its subsidiary, bankruptcy by the subsidiary will cause a considerable economic impact on the parent company. Such cases will be reviewed and final classifications made. (\bullet)

⁹² NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

Complete statistics of the total number of firms, number of defaults and number of other exits in each of the 71 economies from 1992 to 2012 are listed in Table A.11.

III. IMPLEMENTATION DETAILS

Section 1 described the modeling framework underlying the current implementation of the CRI system. It focused on theory rather than the details encountered in an operational implementation. The present section describes how the CRI system handles more specific issues.

Subsection 3.1 describes implementation details related to data, mainly dealing with data cleaning and missing data. Subsection 3.2 describes the specific computation of DTD used by the CRI system that leads to meaningful DTD for financial firms. Subsection 3.3 explains how the calibration previously described in Subsection 1.2 can be implemented. Subsection 3.4 gives the implementation details relevant to the daily output. This includes an explanation of the various modifications needed to compute daily PD so that the daily PD is consistent with the usual month end PD, and a description of the CRI.

3.1. Data Treatment

Fitting data to monthly frequency: Historical end of month data for every firm in an economy is required to calibrate the model. For daily data such as market capitalization, interest rates and stock index values, the last day of the month for which there is valid data is used.

Up to the October 2012 calibration, financial statement (FS) variables data were used, starting from the period end of the statement lagged by 3 months. This is to ensure that predictions are made based on information that was available at the time the prediction was made. However, this treatment can be over-conservative, and many companies actually release their financial statements quicker than 3 months. Therefore, we implement a new logic and we start using the values in an FS as soon as its latest revision was put into the RMI database, unless the FS' release was delayed for more than 3 months. If there was no revision to a FS, the originally released FS is used. Whenever the latest revision is available more than 3 months after the period end, we revert to the previous logic. We start including the FS before the latest revision is actually available as a compromise, to avoid situations like later minor revisions of the FS holding back more up-to-date information. It should be noted that the new approach was only applied for FS input into the RMI database after February 2011, as the revision dates were not accurately recorded before this date. The CRI considers financial statement variables to be valid for one year without restriction, after they were first used.

Priority of financial statements: As described in Subsection 2.2, data provided in Bloomberg's Back Office Product can include numerous versions of financial statements within the same period. If there are multiple financial statements with the same period end, priority rules must be followed in order to determine which to use. The formulation and implementation of these rules are major challenges and areas of continuing development.

The first rule is to prioritize by consolidated/unconsolidated status. This status is relevant only to firms in India, Japan, South Korea and Taiwan, so this rule is only relevant in those economies. Most firms in these economies issue unconsolidated financial statements more frequently than consolidated ones, so these are given higher priority. This simple prioritization can, however, lead to cases where the financial statements used switch from consolidated statements to unconsolidated statements and back again. A more complex prioritization rule is currently under development, with the intention of avoiding this situation.

If, after the first prioritization rule has been applied, there are still multiple financial statements, the second rule is applied. This is prioritization by fiscal period. In most economies, annual statements are required to be audited, whereas other fiscal periods are not necessarily audited. The order of priority from highest to lowest is, therefore: annual, semi-annual, quarterly, cumulative, and finally other fiscal periods.

The third prioritization rule is based on filing status. The "Most Recent" statement is used before the "Original" statement, which is used before the "Preliminary" statement.

GLOBAL CREDIT REVIEW VOLUME 3 93

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The final prioritization rule is based on the accounting standard. Here, financial statements that are reported using Generally Accepted Accounting Principles (GAAP) are given higher priority than financial statements that are reported using International Financial Reporting Standards (IFRS). If an accounting standard is not indicated at all, the financial statement is not used.

Financial statement entries with all other descriptors the same but with different filing statuses will be grouped together. For each variable separately, the variable value is taken from the highest priority financial statement within the group where the value is non-null.

For example, we may consider two financial statement entries having the same period end, and they both are from annual, consolidated statements, and both use the same accounting standard, but the first entry is classified as the "Most Recent" entry and the second is "Original" entry. Suppose the total assets and total liabilities are reported in the "Original" entry, and in the "Most Recent" entry only the total liabilities have been updated but the total assets have been replaced with a null value, then the total liabilities will be taken from the "Original" entry.

The rule mentioned above allows us to group the "Most Recent" and the "Original" entries together, as Bloomberg occasionally only updates values that change without updating other values. If the entries are not grouped, most of the variables would have null values.

One variable that requires special attention is the net income. Net income is a flow variable and needs to be adjusted based on the period of the financial statement. More specifically, we transform the net income into a monthly net income by dividing the net income by the number of months that the financial statement covers. Due to the different coverage periods, several types of net income can still be used. For example, the monthly net income can be computed from the annual net income divided by 12, the semi-annual net income divided by six and the quarterly net income divided by three. When the monthly net income can be obtained from different sources simultaneously, the quarterly net income will have higher priority than any others because it covers a more recent period of time. **Treatment of stale market capitalization prices:** The market capitalization of a firm is required in a few input variables: DTD, SIZE, M/B and SIGMA. For most firms, the market capitalization is available from Bloomberg on a daily basis.

A check on the trading volume of shares is used to remove stale prices. Specifically, if there are more than two consecutive days of identical market capitalization prices, subsequent identical prices are removed only if the trading volume is equal to zero. This is to avoid, for example, cases where the shares of a company are under a trading suspension but the market capitalization data is incorrectly carried forward.

An exception is for Indian companies, where it is common for some companies to have market capitalizations reported only once a month with several consecutive months having identical prices and positive trading volume. These prices are very likely not to be accurate reflections of the firms' value. So, the trading volume is not checked for Indian firms and market capitalizations are excluded after more than two repeated prices.

For some firms, there are gaps in the market capitalization data provided by Bloomberg. Previously, the first recourse was to use the share price multiplied by the shares outstanding listed in the balance sheet and multiplied by an adjustment factor that Bloomberg provides to account for splits, dividends, etc. However, this data is frequently in error and using the shares outstanding as the previous available market capitalization divided by the price on that day was found to be more reliable.

If the gap in market capitalization data is more than a year, then the previous computation using the shares outstanding from the balance sheet is again used. If there are still remaining gaps in the data, then shares outstanding from Compustat data is used.

Currency conversion: Currency conversions are required if the market capitalization or any of the financial statement variables are reported in a currency different than the currency of the economy. If a currency conversion is required, the foreign exchange rate used is that reported at the relevant market close. For firms traded in Asia and Asia-Pacific, the Tokyo closing rate is used; for firms traded in Western Europe, the London closing rate is used; and for firms traded in North America, the New York closing rate is used. For market capitalizations, the FX rate used is

94 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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for the date that the market capitalization is reported. For financial statement variables, the FX rate used is for the date of the period end of the statement.

Provisions for missing values and outliers: Missing values and outliers are dealt with by a threestep procedure. In the first step, the ten firm-specific input variables are computed for all firms and all months. In the second step, outliers are eliminated by winsorization. In the final step, missing values are replaced under certain conditions.

The first step is to compute the input variables and to determine which are missing. As mentioned previously, financial statement variables are carried forward for one year after the date that they are first used. This is generally three months after the period end of the statement. If no financial statement is available for the company within this year, then the financial statement variable will be missing. For market capitalization, if there is no valid market capitalization value within the calendar month, then the value is set to missing.

For illiquid stocks, if there has been no valid market capitalization value for a firm within the last 90 calendar days, then the market capitalization is deemed to not properly reflect the value of the firm. The firm is considered to have exited with a non-default event. Once the firm starts trading again and a new financial statement is released, the firm can enter back into the calibration. With regard to historical PD, the PD can be reported again once there are enough valid variables.

With regard to the level variables, the current month and the last eleven months are averaged to compute the level. A minimum of six observations are required to calculate the level variables. However, this condition is not enforced during the first six months of a firm. In the absence of six valid observations after the initial six months of a company, the level variable will be considered as missing.

To compute the trend variables, the level is subtracted from the current month value. If the current month value is missing, the trend variable is set to be the last valid value during the previous one year.

The value of M/B is set to be missing if any of the following values are missing: market capitalization, total liabilities or total assets of the firm. For the computation of SIGMA, at least 50 valid returns over the last 250 days of possible returns are required for the

regression. If there are less than 50 valid returns, SIGMA is set to be missing.

In this way, the eight trend and level variables as well as M/B and SIGMA are computed and identified as missing or present. Winsorization can then be performed as a second step to eliminate outliers. The volume of outliers is too large to be able to determine whether each one is valid or not, so winsorization applies a floor and a cap on each of the variables. The historical 0.1 percentile and 99.9 percentile for all firms in the economy are recorded for each of the ten variables. Any values that exceed these levels are set to equal these boundary values.

With a winsorization level and 0.1 percentile and 99.9 percentile, the boundary values still may not be reasonable. For example, NI/TA levels of nearly –25, meaning an annual net income –25 times larger than the total assets of a firm, have been observed at this stage. In these cases, a more aggressive winsorization level is applied, until the boundary values are reasonable. Thus, the winsorization level is economy- and variable-specific, and will depend on the data quality for that economy and variable. Winsorization levels different from the default of 0.1 percentile and 99.9 percentile are indicated in RMI (2013).

A third and final step can be taken to deal with missing values. If during a particular month, no variable is missing for a particular firm, the PD can then be computed. If six or more of these ten variables are missing, there is deemed to be too many missing observations and no replacement shall be made.

If between one and five variables are missing out of the ten, the first step is to trace back for at most twelve months to use previous values of these variables instead. If this does not succeed in replacing all of the variables, a replacement by sector medians is done. The median is for the financial or non-financial firms (as indicated by their Bloomberg industry sector code) within the economy during that month. Replacement by the sector median should have a neutral effect on the PD of the firm; the firm is assessed by the other variables that it does have values for. This sector median is always performed in calibration. However, when reporting historical PD, the sector replacement is not done if it results in a relative change in PD of 10% or more where the initial PD was at or above

GLOBAL CREDIT REVIEW VOLUME 3 95

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100 bps, or an absolute change in PD of 10 bps or more where the initial PD was below 100 bps.

However, this treatment of missing values is not always meaningful and occasionally results in counterintuitive patterns in a company's historical PD. Accordingly, the RMI CRI team is reconsidering the treatment of missing values in two stages of development, with the first stage focusing on the replacement of missing values in the initial phase of a company, and the second stage of development focusing on later periods in the company's time series. The first stage has been implemented and is explained below.

In the initial phase of a company — up until six months after IPO — it can be expected that the company's data availability and quality is relatively low due to, for example, a delay in the issuance of financial statements or illiquid trading. So, many companies require missing value replacements during that period. However, as observed in our data, replacing the missing values during these first six months with a sector median affects a company's PD in an unmeaningful way, sometimes resulting in extreme spikes and falls in the company's PD. Since this occurs at the beginning of a company's history, there are no previous PD values to compare to as can be done at later periods in a company's history.

Hence, in order to avoid this, as of the 2013 February calibration, we set a criterion to start the missing value treatment only six months after the beginning of a company's data. Doing so ensures that PDs in the beginning of a company's history are more reflective of the true creditworthiness of that individual company.

The RMI CRI team is currently developing a method to deal with missing values later in the history of a company in a more meaningful way. This second stage of development for treating missing values will be completed in the coming months.

Inclusion/exclusion of companies for calibration: Firms are included within an economy for calibration when the primary listing of the firm is on an exchange in the economy. This ensures that all firms within the economy are subject to the same disclosure and accounting rules.

There are a relatively small number of firms that are dual listed, in which two corporations listed in different exchanges operate as a single entity but retain separate legal status. In the CRI system, a combined company will be assigned to the single economy it is most associated with. An example is the Rio Tinto Group. This consists of Rio Tinto plc, listed in the UK; and Rio Tinto Limited, listed in Australia. Most of Rio Tinto's operations are in Australia rather than the UK, so Rio Tinto is assigned to Australia.

In the US, firms traded on the OTC markets or the Pink Sheets are not considered as exchange listed so are not included in calibration or in the reporting of PD forecasts. Many of these firms are small or start-up firms. Including this large group of companies would skew the calibration and the aggregate results. The TSX Venture Exchange in Canada also contains only small and start-up firms, so firms listed here are also excluded.

Other examples include Taiwan's GreTai Securities Market and Singapore's Catalist. The challenge for markets outside of the US or Canada is that the data on whether firms are listed on the smaller markets rather than the main board is difficult to obtain. For all economies besides the US and Canada, there is continuing work being done in the CRI system to exclude firms that are not listed on major exchanges within a country.

Firms that record an exit (other than due to no trading for 90 calendar days) will not enter back into the calibration even if the firm continues to trade and issues financial statements, as that can happen after the firms declare bankruptcy. There are two exceptions to this exclusion. The first, determined on a case by case basis, is if the firm should be deemed to have re-emerged from bankruptcy. The second exception is for all firms in China, where two situations are prevalent. The first situation is that the firm experiences few repercussions from the default and continues operating normally. The other situation is for one firm to take over a defaulted firm's listing. This happens due to the limited supply of exchange listings. Both of these situations can be considered as emerging from default, so the CRI system enters all of these companies back into the calibration as new companies.

3.2. Distance-to-Default Computation

The DTD computation used in the CRI system is not a standard one. Standard computations exclude

96 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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financial firms, which is of course a critical part of any economy. Thus, the standard DTD computation must be extended to give meaningful estimates for financial firms as well. Duan and Wang (2012) have provided a review of different DTD calculations with several examples for financial and non-financial firms.

The description of the specialized DTD computation starts with a brief description of the Merton (1974) model. Merton's model makes the simplifying assumption that firms are financed by equity and a single zero-coupon bond with maturity date T and principal L. The asset value of the firm V_t follows a geometric Brownian motion:

$$dV_t = \mu V_t dt + \sigma V_t dB_t.$$
(34)

Here, B_i is the standard Brownian motion, μ is the drift of the asset value in the physical measure, and σ is the volatility of the asset value. Equity holders receive the excess value of the firm above the principal of the zerocoupon bond and have limited liability, so the equity value at maturity is: $E_i = \max(V_i - L, 0)$. This is just a call option payoff on the asset value with a strike value of L. Thus, the Black-Scholes option pricing formula can be used to calculate the equity value at times t before T,

$$E_{t} = V_{t}N(d_{+}) - e^{-r(T-t)}LN(d_{-}), \qquad (35)$$

where *r* is the risk-free rate, $N(\cdot)$ is the standard normal cumulative distribution function, and

$$d_{\pm} = \frac{\log\left(\frac{V_T}{L}\right) + \left(r \pm \frac{\sigma^2}{2}\right)(T-t)}{\sigma\sqrt{T-t}}.$$
(36)

Following the Merton (1974) model, the probability of the company's default at time T evaluated at time tis $N(-DTD_t)$, where DTD at time t is defined as:

$$\text{DTD}_{t} = \frac{\log\left(\frac{V_{T}}{L}\right) + \left(\mu - \frac{\sigma^{2}}{2}\right)(T-t)}{\sigma\sqrt{T-t}}.$$
 (37)

The standard KMV assumptions given in Crosbie and Bohn (2003) are to set the time to maturity T - t at a value of one year, and the principal of the zero-coupon

bond L to a value equal to the firm's current liabilities plus one half of its long-term debt. Here, the current liabilities and long-term debt are taken from the firm's financial statements. If the firm is missing the current liabilities field, then various substitutes for this field can be used, as described in Subsection 2.3.

This is a poor assumption of the debt level for financial firms, since they typically have large liabilities, such as deposit accounts, that are neither classified as current liabilities nor long-term debt. Thus, using these standard assumptions means ignoring a large part of the debt of financial firms.

To properly account for the debt of financial firms, Duan (2010) included a fraction δ of a firm's other liabilities. The other liabilities are defined as the firm's total liabilities minus both the short and longterm debt. The debt level *L* then becomes the current liabilities plus half of the long-term debt plus the fraction δ multiplied by the other liabilities, so that the debt level is a function of δ . The standard KMV assumptions are then a special case where $\delta = 0$.

The fraction δ can be optimized along with μ and σ in the maximum likelihood estimation method developed in Duan (1994, 2000). Following Duan *et al.* (2012), the firm's market value of assets is standardized by its book value A_i , so that the scaling effect from a major investment or financing by the firm will not distort the time series from which the parameter values are estimated. Thus, the log-likelihood function is:

$$\mathcal{L}(\mu,\sigma,\delta) = -\frac{n-1}{2}\log(2\pi) - \frac{1}{2}\sum_{t=2}^{n}\log(\sigma^{2}h_{t}) - \sum_{t=2}^{n}\log\left(\frac{\hat{V}_{t}(\sigma,\delta)}{A_{t}}\right) - \sum_{t=2}^{n}\log[N(\hat{d}_{+}(\hat{V}_{t}(\sigma,\delta),\sigma,\delta))] - \frac{1}{2\sigma^{2}}\sum_{t=2}^{n}\frac{1}{h_{t}}\left[\log\left(\frac{\hat{V}_{t}(\sigma,\delta)}{A_{t}} \times \frac{A_{t-1}}{\hat{V}_{t-1}(\sigma,\delta)}\right) - \left(\mu - \frac{\sigma^{2}}{2}\right)h_{t}\right]^{2}, \quad (38)$$

where *n* is the number of days with observations of the equity value in the sample, \hat{V}_t is the implied asset value found by solving Equation (35), \hat{d}_+ is computed with Equation (36) using the implied asset value, and h_t is the number of trading days as a fraction of the year between observations t - 1 and t. Notice that the implied asset value and \hat{d}_+ are dependent on δ by virtue of the dependence of L on δ .

GLOBAL CREDIT REVIEW VOLUME 3 97

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Implementation of DTD computation: The DTD at the end of each month is needed for every firm in order to calibrate the forward intensity model. A moving window, consisting of the last one year of data before each month end is used to compute the month end DTD. Daily market capitalization data based on closing prices is used for the equity value in the implied asset value computation of Equation (34). If there are fewer than 50 days of valid observations for the market capitalization, then the DTD value is set to missing. An observation is valid if there is positive trading volume that day. If the trading volume is not available, the observation is assumed to be valid if the value for the market capitalization changes often enough. The precise criterion is as follows: if the market capitalization does not change for three days or more in a row, the first day is taken as a valid observation and the remaining days with the same value are set to missing.

A straightforward idea for the DTD computation is to first estimate the three variables μ , σ and δ via maximizing the log-likelihood function (38) over $\sigma \ge 0$ and $0 \le \delta \le 1$ and then to calculate the DTD from Equation (37). Let $(\hat{\mu}, \hat{\sigma}, \hat{\delta})$ be an optimal solution to the maximization problem. By direct calculation, it is not hard to see that:

$$\hat{\mu} = \frac{\hat{\sigma}^2}{2} + \frac{1}{\sum_{t=2}^n h_t} \log\left(\frac{\hat{V}_n(\hat{\sigma},\hat{\delta})}{A_n} \times \frac{A_1}{\hat{V}_1(\hat{\sigma},\hat{\delta})}\right). \quad (39)$$

In view of this, maximizing the three dimensional function $\mathcal{L}(\mu, \sigma, \delta)$ can be equivalently reduced to maximizing the two dimensional function $\widetilde{\mathcal{L}}(\sigma, \delta)$ taking the form:

$$\begin{split} \tilde{\mathcal{L}}(\sigma,\delta) &= -\frac{n-1}{2}\log(2\pi) - \frac{1}{2}\sum_{t=2}^{n}\log(\sigma^{2}h_{t}) - \sum_{t=2}^{n}\log\left[\frac{\hat{V}_{t}(\sigma,\delta)}{A_{t}}\right] \\ &- \sum_{t=2}^{n}\log[N(\hat{d}_{+}(\hat{V}_{t}(\sigma,\delta),\sigma,\delta))] \\ &- \frac{1}{2\sigma^{2}} \left\{\sum_{t=2}^{n}\frac{1}{h_{t}} \left[\log\left(\frac{\hat{V}_{t}(\sigma,\delta)}{A_{t}} \times \frac{A_{t-1}}{\hat{V}_{t-1}(\sigma,\delta)}\right)\right]^{2} \\ &- \frac{1}{\sum_{t=2}^{n}h_{t}} \left[\log\left(\frac{\hat{V}_{n}(\sigma,\delta)}{A_{n}} \times \frac{A_{1}}{\hat{V}_{1}(\sigma,\delta)}\right)\right]^{2} \right\}. \end{split}$$

$$(40)$$

98 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

However, with quarterly financial statements there will never be more than three changes in the corporate structure (defined in this model by *L* and *A_i*) throughout the year, leading to possibly unstable estimates of δ . This problem is mitigated by performing a two-stage optimization for σ and δ .

In the first stage, the maximization of $\mathcal{L}(\sigma, \delta)$ for each firm is performed over both σ and δ . For each firm, at the first month in which DTD can be computed the maximization is constrained in $\sigma \ge 0$ and $0 \le \delta \le 1$. Thereafter, at month *n*, the maximization is still constrained in $\sigma \ge 0$ while δ is constrained in the interval, [max(0, $\hat{\delta}_{n-1} - 0.05$), min(1, $\hat{\delta}_{n-1} + 0.05$)], where $\hat{\delta}_{n-1}$ is the estimate of δ made in the previous month. In other words, a ten percent band around the previous estimate of δ (where that band is floored with 0 and capped with 1) is applied so that the estimates do not fluctuate too much from month to month.

However, for many firms, the estimate of δ would frequently lie on the boundary of the constraining interval, meaning that the estimates of δ were not stable. Therefore, a second stage is implemented to impose greater stability. All financial sector firms in the same economy are assumed to share the same estimate of δ , chosen to be the average of all its individual estimates. The same is done for non-financial firms. Accordingly, with δ being fixed to be the sector average, the original maximization of $\mathcal{L}(\sigma, \delta)$ is reduced to a one-dimensional maximization in σ . Thus, this maximization is used to perform the estimates of σ for each firm.

Since the first stage is done to obtain a stable sector-average estimate of δ , the criteria used to include a firm-month is more strict. In the first stage, a two-year window is used instead of one year, and a minimum of 250 days of valid observations of the market capitalization are required instead of 50. If a firm has less than 250 days of valid observations within the last two years of a particular month end, δ will not be estimated for that firm and that month end.

It was found that after applying the two-stage procedure described above, the estimate of μ was frequently unstable and could lower the explanatory power of DTD. For example, suppose a firm has a large drop in its implied asset value in January 2011,

so that the estimated μ is negative for the DTD calculation at the end of December 2011. If there is little change in the company in January 2012, then the drop in implied asset value in January 2011 is no longer within the observation window for the DTD calculation at the end of January 2012. There will be a large increase in the estimated μ , resulting in a substantial improvement of the DTD just because of the moving observation window. To avoid this problem, we now set μ to be equal to $\sigma^2/2$. So in calculating DTD, the second term in the numerator of Equation (37) is eliminated.

In summary, the DTD for each firm is computed using the economy and sector (financial or non-financial) average for δ in that month, and the estimate of σ based on the last year of data for the firm.

Carrying out this two-stage procedure would take 70 hours of computation time on a single PC, given the millions of firm months that are required. However, each of the stages is parallelizable. In the first stage the DTD can be computed independently between firms. In the second stage, once the sector averages of the δ have been computed for each month, the DTD can again be computed independently between firms. In the current CRI system, by using a computational grid administered by the NUS computer center, the DTD computational time for all firms over the full history of twenty years takes only about 3.5 hours.

3.3. Calibration

Implementation: As shown in Section (1), the calibration of the forward intensity model involves multiple maximum pseudo-likelihood estimations, where the pseudo-likelihood functions are given in Equation (13). The maximizations are on the logarithm of these expressions, and the default parameters' maximization is performed independently from the non-default exit parameters. Parameter estimates for the entire horizon up to five years for the default and non-default exits can be obtained directly from the NS function.

A few input variables have an unambiguous effect on a firm's probability of default. Increments of both the level and trend of DTD, CASH/TA, and NI/TA should indicate that a firm is becoming more creditworthy and should lead to a decreasing PD. For large and relatively clean data sets such as the US, an unconstrained optimization leads to parameter values which mostly have the expected sign. For each of the DTD level and trend, CASH/TA level and trend, and NI/TA level, the default parameters at all horizons are negative. A negative default parameter at a horizon means that if the variable increases, the forward intensity will decrease (based on Equation (6)), so that the conditional default probability at that horizon will decrease.

Grouping for economies: There are not enough defaults in some small economies and calibrations of these individual economies are not statistically meaningful. In order to ensure that there are enough defaults for calibration, the 71 economies are categorized into groups according to similarities in their stage of development and their geographic locations. Within these groups, the economies are combined and calibrated together.

Starting from the Aug 2012 calibration, Canada and the US remain in the same calibration group, and the developed economies of Asia-Pacific (Australia, Hong Kong, Japan, Singapore, South Korea, Taiwan and New Zealand) form another calibration group. China and India, the two major emerging economies of Asia-Pacific are each calibrated as individual groups. All the European countries covered by the CRI are in a single calibration group, which now includes Israel, Russia and Turkey. The other emerging economies of Asia-Pacific (Kazakhstan, Indonesia, Malaysia, Pakistan, Philippines, Sri Lanka, Thailand and Vietnam) are grouped together with the Latin American economies (Argentina, Brazil, Colombia, Chile, Mexico, Peru, and Venezuela), Middle-East economies (Bahrain, Jordan, Kuwait, Saudi Arabia and United Arab Emirates) and African economies (Egypt, Morocco, Nigeria and South Africa), to form the "emerging markets" calibration group.

All economies in these new calibration groups share the same coefficients for all variables except for the benchmark risk-free interest rate variable. The benchmark interest rates coefficients will be allowed to vary, because different economies based in different ()

currencies naturally have different dependencies on their interest rates and the interest rate levels can differ significantly across economies. After adopting the euro, all eurozone countries use Germany's three-month Bubill rate as this is more reflective of monetary rather than sovereign credit conditions in each economy, which is the intent of this variable. For the period before joining the eurozone, their own interest rates are used.

In addition, the benchmark interest rate is entered as the current value minus the historical month-end mean. This allows the variable to reflect its value relative to the historical average. When an economy does not have enough default events to identify a separate interest rate coefficient, the interest rate variable will be disabled for that economy by inputting a zero value for the whole time series. In fact, that is also why we de-mean all interest rate series so that setting the interest rate series of a particular economy to zero, when necessary, does not induce a bias by the base economy in the same group.

Since all eurozone countries except Germany do not have enough default events prior to joining the eurozone, their benchmark interest rate is entered as zero for that period. Among the non-eurozone members of the European group, Denmark, Norway, Sweden and the UK each have separate coefficients for the benchmark interest rate. Switzerland and Iceland do not use this variable for their whole history.

In the Developed Asia-Pacific group, all economies have their own coefficient for the benchmark interest rate, except for Japan and New Zealand who share the same coefficient. For the North American group, both Canada and the US have their own coefficient for the benchmark interest rate.

In the Emerging Markets group, there are insufficient defaults in the Latin American economies to calibrate individual economy benchmark interest rate coefficients in a statistically significant way, so all Latin American economies share the same benchmark interest rate coefficient. Among the Asian economies in the Emerging Markets group, Indonesia, Malaysia, and Philippines have their own coefficient for the benchmark interest rate, while Vietnam does not use this variable. All the other economies in the Emerging Markets group share the same benchmark interest rate coefficient. Indonesia also has its own intercept and uses its own coefficients for the stock index return, CASH/TA level and Relative Size level. These coefficients are required because these characteristics for defaulting firms in Indonesia are substantially different than in other economies. Separate coefficients are required to improve the accuracy of the PD forecasts.

Relative Size: For the calibration data set, the median market cap of firms in an economy for each month end includes the market cap from the last trading day of each firm in the month. If a firm does not trade in a particular month, the firm's market cap is not included in the median. For certain economies, many firms are illiquid and the median market cap experiences large variations due to the change in composition of firms rather than the market value of the firms. Another problem is data quality at the beginning of the historical sample: if a data provider starts including the market cap for a large number of firms in one month compared to the previous, there can be a large jump in the median market cap.

To avoid this problem, we use a combination of the economy's stock index and the economy's median market cap as the divisor in the Relative Size variable:

- 1. We choose a recent month where there is a more complete set of firms in the economy that have trading activity, and calculate the ratio of the economy's median market cap to stock index value at the end of the month.
- 2. For each month, the divisor for the Relative Size variable of firms in the economy is taken as the month end stock index multiplied by that ratio.

3.4. Daily Output

Individual firms' PD: In computing the pseudo-log-likelihood functions in Equation (13), only the end of month data is needed. The data needs to be extended to daily values in order to produce daily PDs.

For the level variables, the last 12 end-of-month observations (before averaging) are combined with the current value. The current value is scaled by a fraction equal to the current day of the month divided by the number of calendar days in the month. The earliest monthly value is scaled by one minus this fraction. The sum is then divided by the number of valid monthly observations, with the current value and the earliest monthly value jointly having the weight of one observation if either or both are not

100 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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missing. Not performing this scaling can lead to an artificial jump in PD at the beginning of the month. When performing the scaling, the change in level is more gradual throughout the month.

SIGMA is computed by regressing the daily returns of the firm's market capitalization against the daily returns of the economy's stock index for the previous 250 days.

Exclusion rule for Mergers & Acquisitions: Mergers & Acquisitions (M&A) events are common occurrences. When an important M&A deal is closed, the Market Capitalization (MC) of the acquirer changes immediately as the market cap of the acquirer will now reflect the joint value of the acquirer and the target. However, the financial statement will not immediately reflect the new situation. In this case, the Distance-to-Default (DTD) and market-to-book ratio, which are important inputs for the PD computation, will be distorted due to a mismatch in the MC and the financial statement variables. In order to ensure the accuracy and reliability of our PD estimates, we apply a rule to disable PD calculations for companies that are involved in important M&A deals.

An important M&A deal is defined as an M&A event on which all the following three criteria apply:

- 1. Upon the deal's completion, the acquiring company owns 20% or more of the target company.
- 2. The size of the deal is material to the acquirer. This is measured in terms of total assets. If α is the percentage of the target that is being acquired, the size is considered material if the product of α and the total assets of the target is greater than or equal to 20% of the total assets of the acquirer.
- The change in MC is material, with the largest absolute daily MC return, within 20 days of the M&A completion day, larger than or equal to 5%.

In the event of an important M&A, the PDs of the acquirer will be not be computed until we are able to include financial statement variables reflecting the new situation (typically between 3 and 6 months after deal completion). The RMI CRI team is currently developing a method to deal with M&A cases more systematically, and will avoid having to disable PDs for companies involved in an important M&A deal.

Aggregating PD: The CRI provides term structures of the probability distributions for the number of

defaults as well as the expected number of defaults for different groups of firms. The companies are grouped by economy (using each firm's country of domicile), by sector (using the firm's Bloomberg industrial sector code) and sectors within economies. With the individual firms' PD, the expected number of defaults is trivial to compute. The algorithm used to compute the probability distribution of the number of defaults was originally reported in Anderson *et al.* (2003). It assumes conditional independence and uses a fast recursive scheme to compute the necessary probability distribution.

Note that while this algorithm is currently used to produce the probability distribution of the number of defaults within an economy or sector, it can easily be generalized to compute loss distributions for a portfolio manager, in which case the portfolio's exposure to each firm should be aggregated.

Inclusion of firms in aggregation: As explained in Subsection 3.1, firms are included in an economy for calibration if the firms' primary listing is on an exchange in that economy. This is to ensure that all firms in an economy are subject to the same disclosure and accounting requirements. In contrast, a firm is included in an economy's aggregate results if the firm is domiciled in that economy. This is because users typically associate firms with their economy of domicile rather than the economy where their primary listing is, if they are different. For example, the Bank of China has its primary listing in Hong Kong, but its economy of domicile is China so the Bank of China is included in the aggregation forecasts for China, and is included under China when searching for the individual PDs.

Treatment of companies after a default event: When a company experiences a default event, the CRI system discontinues the PD calculation for that company. However, if the company resumes operations after some time, it will be treated as a new company, and we continue to generate PD. The new company's PDs are not affected by the financial statement or market cap data prior to the event. So, the PDs calculated are independent of the PDs that were generated before the default event. On our website, the PDs are however displayed on a single graph for the convenience of our users.

In order to implement the treatment, default events are classified into hard defaults and soft defaults (see Table 1). ()

Table 1.	Classification	of	default	events.
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Default-Action Type	Subcategory	
Hard Defaults	Administration, Arrangement, Canadian CCAA, Chapter 7, Chapter 11,	
(Default events that are typically	Chapter 15, Conservatorship, Insolvency, Japanese CRL, Judicial Management	
permanent)	Liquidation, Pre-Negotiated Chapter 11, Protection, Receivership,	
	Rehabilitation, Rehabilitation (Thailand 1997), Reorganization, Restructuring,	
	Section 304, Supreme court declaration, Winding Up, Work Out, Other,	
	Unknown	
Soft Defaults	Coupon & Principal Payment, Coupon Payment Only, Debt Restructuring,	
(Default events that companies	Interest Payment, Loan Payment, Principal Payment, ADR (Japan only),	
can emerge from)	Declared Sick (India only), Unknown	

Hard defaults are default events that are typically permanent. In other words, companies typically cannot emerge from hard defaults. An example of a hard default is a forced liquidation of a company. PDs will not be computed after the default event unless there is an exceptional circumstance that warrants a manual intervention. General Motors is an example of such an event. Although GM filed for Chapter 11 reorganization in June 2009, the company resumed operations in March 2011. As of March 2011, after the company resumed operations, we decided to treat GM as a new company.

Soft defaults are default events that companies can typically emerge from. An example of a soft default is a debt restructuring. More specifically, after a soft default, if there is sufficient data for the company, then the company is assumed to have been able to continue its operations and PDs are computed. The PDs are generated once sufficient history of both the market capitalization data and the new financial statement data (released after the event) becomes available. Take the Australian company Marion Energy Ltd as an example, which had a debt restructuring in April 2010. We stopped calculating PD after 31 March 2010. As debt restructuring is considered as a soft default, we started calculating PD again from 30 Sept 2010 onwards, when data requirements were met.

This treatment does not apply to Chinese companies, based on two reasons: (1) a firm typically experiences few repercussions from the default and continues operating normally; and (2) it is common for another firm to take over a defaulted firm's listing, due to the limited supply of exchange listings. Both of these situations can be considered as emerging from default,

102 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

so the CRI system enters all of these companies back into the calibration as new companies.

IV. EMPIRICAL ANALYSIS

This section presents an empirical analysis of the CRI outputs for the 71 economies with their own exchange that are currently being covered. In Subsection 4.1, an overview is given of the default parameter estimates. Subsection 4.2 explains and provides the accuracy ratios for the different countries under the CRI coverage.

4.1. Parameter Estimates

With 60 months of forecast horizons, 13 variables and 6 different groups of economies, tables of the parameter estimates occupy over 20 pages and are not included in this Technical Report. In Figures B.1 and B.2, the parameter estimates are from calibrations performed in September 2013 using data up until the end of August 2013. As an example, plots of the default parameters for the US are given in figures included in Figures B.1 and B.2 in Appendix B. In this part, a brief overview is given of the general traits and patterns seen in the default parameter estimations of the cRI.

Recall that if a default parameter for a variable at a particular horizon is estimated to be positive (negative) from the maximum pseudo-likelihood estimate, then an increasing value in the associated variable will lead to an increasing (decreasing) value of the forward

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intensity at that horizon, which in turn means an increasing (decreasing) value for the conditional default probability at that horizon.

For the stock index one-year trailing return variable, most groups have default parameters that are slightly negative in the shorter horizons and then become positive in the longer horizons. When the equity market performs well, this is only a short-term positive for firms and in the longer term, firms are actually more likely to default. This seemingly counterintuitive result could be due to correlation between the market index and other firm-specific variables. For example, Duffie *et al.* (2009) suggested that a firm's DTD can overstate its creditworthiness after a strong bull market. If this is the case, then the stock index return serves as a correction to the DTD levels at these points in time.

As expected we observe a different relationship between the short-term interest rate and default across economies. This observation possibly indicates different lead-lag relationships between credit conditions and the raising and cutting of short-term interest rates.

DTD is a measure of the volatility-adjusted leverage of a firm. Low or negative DTD indicates high leverage and high DTD indicates low leverage. Therefore, PD would be expected to increase with decreasing DTD. Indeed, almost all of the calibrations for the different groups lead to negative default parameters for the DTD level.

The ratio of the sum of cash and short-term investments to total assets (CASH/TA) measures liquidity of a firm. This indicates the availability of a firm's funds and its ability to make interest and principal payments. As expected, for almost all economies (Indonesia being the only exception) the default parameters for CASH/TA level in shorter horizons are significantly negative. The magnitude of the default parameters typically decreases for longer horizons, indicating that CASH/TA level is a better indicator of a firm's ability to make payments in the short term than the long term.

The ratio of net income to total assets (NI/TA) measures profitability of a firm. The relationship between PD and NI/TA is as expected: the default parameters for NI/TA level is negative for most economies and most horizons.

The logarithm of the market capitalization of a firm over the median market capitalization of firms within the economy (SIZE) does not have a consistent effect on PD across different economies. For example, in the US the default parameters for SIZE level are positive for all horizons, suggesting that the complexity of larger firms outweighs the potential benefits, such as diversified business lines and funding sources. On the other hand, in Europe the default parameters for SIZE level are negative across all horizons. These differences may reflect differences in the business environments in the respective economies.

The default parameters associated with DTD Trend, CASH/TA Trend, SIZE Trend and NI/TA Trend are negative across almost all economies and horizons. The trend variables reflect momentum. The momentum effect is a short-term effect, and evidence of this is seen in the lower magnitude of the default parameters at longer horizons than at shorter horizons. The exception is the NI/TA Trend, which for some calibration groups has a higher magnitude at longer horizons.

The ratio of the sum of market capitalization and total liabilities to total assets (M/B) can either indicate the market mis-valuation effect or the future growth effect. This default parameter is negative in most economies, indicating that higher M/B implies higher PD, and the future growth effect dominates. On the other hand, in China and in the Developed Asia-Pacific calibration group, the default parameter for M/B is positive, indicating that for these economies, the market mis-valuation effect dominates.

Shumway (2001) argued that a high level of the idiosyncratic volatility (SIGMA) indicates highly variable stock returns relative to the market index, which is equivalent to highly variable cash flows. Currently, for the different economies, this variable is no longer significant.

4.2. Prediction Accuracy

In-sample testing: Various tests are carried out to test the prediction accuracy of the RMI PD forecasts. These tests are conducted in-sample.

A single calibration is conducted for the in-sample tests, using data to the end of the data sample. As an example, one-year PD forecasts are made for December 31, 2000 by using the data at or before

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December 31, 2000 and the parameters from the calibration. These PD forecasts can be compared to actual defaults that occurred at any time in 2001.

Accuracy Ratio: The accuracy ratio (AR) is one of the most popular and meaningful tests of the discriminatory power of a rating system (BCBS, 2005). The AR and the equivalent Area Under the Receiver Operating Characteristic (AUROC) are described in Duan and Shrestha (2011). In short, if defaulting firms had been assigned among the highest PD of all firms before they defaulted, then the model has discriminated well between safe and distressed firms. This leads to higher values of AR and AUROC. The range of possible AR values is in [0,1], where 0 is a completely random rating system and 1 is a perfect rating system. The range of possible AUROC values is in [0.5, 1]. AUROC and AR values are related by: AR = $2 \times AUROC$ -1.

The AR and AUROC values for different horizons are available in Table B.1 of this technical report. Only economies with more than 20 defaults entering into the AR and AUROC computation are listed. The PD are taken to be non-overlapping. For example, the one-year AR is based on PDs computed on 31/12/2000, 31/12/2001,..., 31/12/2009 and firms defaulting within one year of those dates, while the two-year AR is based on PDs computed on 31/12/2000, 31/12/2002,..., 31/12/2008 and firms defaulting within two years of those dates.

The AUROC values have been provided only for the purpose of comparison, if other rating systems report their results in terms of AUROC. The discussion will focus only on AR. The model is able to achieve strong AR results mostly greater than 0.80 at the one and six-month horizons for developed economies. There is a drop in AR at one, two and five-year horizons, but the AR are still mostly acceptable. Australia, the UK and Singapore have sharp drops in AR at longer horizons. Hong Kong has comparatively worse AR over all horizons as compared to other developed economies.

The AR in emerging market economies such as Brazil, China, India, Indonesia, Malaysia, Philippines, Russia, and Vietnam are noticeably weaker than the results in the developed economies. This can be due to a number of issues. The quality of data is worse in emerging markets, in terms of availability and data errors. This may be due to lower reporting and auditing standards. Also, variable selection is likely to play a more important role in emerging markets. The variables were selected based on the predictive power in a developed economy, the US. Performing variable selections specific to the calibration group are expected to improve predictive accuracy, especially in emerging market economies. Finally, there could be structural differences in how defaults and bankruptcies occur in emerging market economies. If the judicial system is weak and there are no repercussions for default, firms may be less reluctant to default.

The Basel Committee states in BCBS (2004) that banks must also use other quantitative validation tools and comparisons with relevant external data sources. Thus, rating institutions are required to think about further validation of their models using external data sources. We compare the accuracy ratio of the RMI PD with other rating agencies such as Moody's, S&P's and Fitch separately. For a given rating agency, comparisons are made for firms that have both a corporate rating from the rating agency and an RMI PD over the sample period. We expanded our data to monthly data for each firm from the first month a rating is given until the most recent month.

For S&P, this data set has 4580 listed firms from around world. It consists of 393 default events and 439,652 end of month observations from 1990 to January 2013. The 1-year AR for the RMI PD is 0.87 and the AR for the S&P ratings is 0.81. For Moody's, this sample includes 1805 listed firms from different countries. It covers 198 default events and 117,262 end of month observations. The AR for the Moody's ratings is 0.81 and the 1-year RMI PD is 0.88. In the comparison against ratings from Fitch, there are 1490 firms with 49 default events. The end of month data points are 150,917. The Fitch AR and RMI PD are 0.62 and 0.83, respectively. In summary, the RMI PDs have a stronger performance in terms of AR than any of the three major credit rating agencies.

The AR is a test of discriminatory power, or how well the rating system ranks firms in terms of creditworthiness. In a separate article included in Volume 2 of the GCR (2012), we provided a more qualitative check on the RMI PD in which we compare the

 $104 \quad \text{nus-rmi credit research initiative technical report}$

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behaviour of the RMI PD to the rating actions of external credit rating agencies such as Moody's and S&P for some well known default cases.

Aggregate defaults: The time series of aggregate predicted number of defaults and actual number of defaults in each calibration group are also available in Figures B.3 to B.8. For China and India in particular, these figures show that there is room for improvement in the predictive power of the model.

V. ONGOING DEVELOPMENTS

The CRI can be developed along a number of directions. We now comment on obvious ones that in our view are likely to bring meaningful and measurable benefits. Besides modifications to the current modeling framework of the forward intensity, a change in modeling platform will be undertaken if another model proves more promising in terms of accuracy and robustness of results. For this type of development, we also rely on the collective efforts by the worldwide credit research community to challenge and improve the existing modeling platform.

As an example, the CRI will soon start a parallel implementation using the partially-conditioned forward intensity approach proposed in Duan and Fulop (2013) to study its practicality and performance. In fact, the parameter constraints on the forward-intensity function essential to the implementation of that approach has already been incorporated into the current CRI system in making longer-term default predictions.

Within the current modeling framework, future developments involve, for example, variable selection where more experiments are needed to identify common risk factors and company-specific attributes that are more indicative of defaults in emerging markets.

Finally, a series of new applications and tools using the RMI PD as an input are currently being developed. More specifically, RMI plans to actively work with users to discuss the possibility of taking advantage of the world class research infrastructure at the institute to propagate real world applications in credit rating and testing. Some interesting areas include research in counterparty risk management and stress testing of financial systems by policy makers. RMI also remains committed to making its vast resources available for academic research.

The RMI Credit Research Initiative is premised on the concept of credit ratings as a "public good". Being a non-profit undertaking allows a high level of transparency and collaboration that other commercial credit rating systems can not replicate. The research and support infrastructure is in place and researchers from around the world are invited to contribute to this initiative. Any methodological improvements that researchers develop will be incorporated into the CRI system. In essence, the initiative operates as a "selective wikipedia" where many can contribute but implementation control is retained.

If you have feedback on this technical report or wish to work with us in this endeavor, please contact us at rmicri@globalcreditreview.com.

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GLOBAL CREDIT REVIEW VOLUME 3 105

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106 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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APPENDIX A: DATA

Table A.1. All countries under the RMI coverage.

Region Economy	
Asia-Pacific (21)	Australia, China, Hong Kong, Indonesia, India, Japan, Kazakhstan, Malaysia, Pakistan, Philippines,
	Singapore, South Korea, Sri Lanka, Taiwan, Thailand, Vietnam, New Zealand, Cambodia, Macau, Mongolia
	and Papua New Guinea.
North America (2)	Canada, the United States.
Europe (43)	Austria, Belgium, Bermuda, Bulgaria, Croatia, Cyprus, Czech Republic, Denmark, Estonia, Finland, France
	Germany, Greece, Hungary, Iceland, Ireland, Italy, Latvia, Lithuania, Luxembourg, Macedonia, Malta,
	Netherlands, Norway, Poland, Portugal, Romania, Russian Federation, Slovakia, Slovenia, Spain, Sweden,
	Switzerland, Turkey, the United Kingdom, Ukraine, Faeroe Island, Gibraltar, Guernsey, Isle of Man, Jersey,
	Liechtenstein and Monaco.
Latin America (17)	Argentina, Brazil, Columbia, Chile, Mexico, Peru, Venezuela, Bahamas, Belize, Cayman Islands, Curacao,
	Dominican Republic, Falkland Islands, Panama, Puerto Rico, Virgin Islands and Virgin Islands, British.
Middle-East (10)	Bahrain, Israel, Jordan, Kuwait, Saudi Arabia, United Arab Emirates, Azerbaijan, Iraq, Qatar and Sudan.
Africa (13)	Angola, Egypt, Morocco, Nigeria, South Africa, Gabon, Mauritius, Mozambique, Namibia, Sierra Leone,
	Tanzania, United Republic of Togo and Zambia.

Table A.2. The 71 countries under the RMI coverage for which we cover companies listed on the exchange.

Region	Economy
Asia-Pacific (17)	Australia, China, Hong Kong, Indonesia, India, Japan, Kazakhstan, Malaysia, Pakistan, Philippines, Singapore,
	South Korea, Sri Lanka, Taiwan, Thailand, Vietnam and New Zealand.
North America (2)	Canada and the United States.
Europe (35)	Austria, Belgium, Bulgaria, Croatia, Cyprus, Czech Republic, Denmark, Estonia, Finland, France, Germany,
	Greece, Hungary, Iceland, Ireland, Italy, Latvia, Lithuania, Luxembourg, Macedonia, Malta, Netherlands,
	Norway, Poland, Portugal, Romania, Russian Federation, Slovakia, Slovenia, Spain, Sweden, Switzerland,
	Turkey and the United Kingdom, Ukraine.
Latin America (7)	Argentina, Brazil, Columbia, Chile, Mexico, Peru and Venezuela.
Middle-East (6)	Bahrain, Israel, Jordan, Kuwait, Saudi Arabia and United Arab Emirates.
Africa (4)	Egypt, Morocco, Nigeria and South Africa.

Table A.3. The 35 countries under the RMI coverage for which we cover companies domiciled in the economy but listed on a foreign exchange included in Table A.2. The bolded text indicate that these economies also have their own local stock exchange.

Angola	Gibraltar	Panama
Azerbaijan	Guernsey	Papua New Guinea
Bahamas	Iraq	Puerto Rico
Belize	Isle of Man	Qatar
Bermuda	Jersey	Sierra Leone
Cambodia	Liechtenstein	Sudan
Cayman Islands	Macau	Tanzania, United Republic of
Curacao	Mauritius	Togo
Dominican Republic	Monaco	Virgin Islands
Faeroe Island	Mongolia	Virgin Islands, British
Falkland Islands	Mozambique	Zambia
Gabon	Namibia	

GLOBAL CREDIT REVIEW VOLUME 3 107

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b1728 Global Credit Review Volume 3

ISO Code	Economy	Calibration Group
ARE	United Arab Emirates	Emerging
ARG	Argentina	Emerging
AUS	Australia	Developed Asia-Pacific
AUT	Austria	Europe
BEL	Belgium	Europe
BGR	Bulgaria	Europe
BHR	Bahrain	Emerging
BRA	Brazil	Emerging
CAN	Canada	North America
CHE	Switzerland	Europe
CHL	Chile	Emerging
CHN	China	China
COL	Colombia	Emerging
СҮР	Cyprus	Europe
CZE	Czech Republic	Europe
DEU	Germany	Europe
DNK	Denmark	Europe
EGY	Egypt	Emerging
ESP	Spain	Europe
EST	Estonia	Europe
FIN	Finland	Europe
FRA	France	Europe
GBR	United Kingdom	Europe
GRC	Greece	Europe
HKG	Hong Kong	Developed Asia-Pacific
HRV	Croatia	Europe
HUN	Hungary	Europe
IDN	Indonesia	Emerging
IND	India	India
IRL	Ireland	Europe
ISL	Iceland	Europe
ISR	Israel	Europe
ITA	Italy	Europe
JOR	Jordan	Emerging
JPN	Japan	Developed Asia-Pacific
KAZ	Kazakhstan	Emerging
KOR	South Korea	Developed Asia-Pacific
KWT	Kuwait	Emerging
LKA	Sri Lanka	Emerging
LTU	Lithuania	Europe
LUX	Luxembourg	Europe
LVA	Latvia	Europe
MAR	Morocco	Emerging
MEX	Mexico	Emerging

Table A.4. ISO codes for economies currently covered by the CRI and the group that each economy is calibrated in.

(Continued)

 $108 \quad \text{nus-rmi credit research initiative technical report}$

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ISO Code	Economy	Calibration Group
MKD	Macedonia	Europe
MLT	Malta	Europe
MYS	Malaysia	Emerging
NGA	Nigeria	Emerging
NLD	Netherlands	Europe
NOR	Norway	Europe
NZL	New Zealand	Developed Asia-pacific
PAK	Pakistan	Emerging
PER	Peru	Emerging
PHL	Philippines	Emerging
POL	Poland	Europe
PRT	Portugal	Europe
ROM	Romania	Europe
RUS	Russian Federation	Europe
SAU	Saudi Arabia	Emerging
SGP	Singapore	Developed Asia-Pacific
SVK	Slovakia	Europe
SVN	Slovenia	Europe
SWE	Sweden	Europe
THA	Thailand	Emerging
TUR	Turkey	Europe
TWN	Taiwan	Developed Asia-Pacific
UKR	Ukraine	Emerging
USA	United States	North America
VEN	Venezuela	Emerging
VNM	Vietnam	Emerging
ZAF	South Africa	Emerging

Table A.4.(Continued)

Table A.5. The stock indices used for each economy in computing the first common variable.

Country	Stock Exchange	Period Used
ARE	FTSE NASDAQ DUB UAE 20	6/28/2006-Present
ARG	Buenos Aires Stock Exchange Merval Index	
AUS	All Ordinaries Index	
AUT	Austrian Traded ATX Index	
BEL	Belgian All Shares Return Index	
BGR	Bulgaria Stock Exchange Sofix Index	10/24/2000-Present
BHR	Bahrain Bourse All Share Index	7/8/2004-Present
BRA	Brazil Bovespa Stock Index	
CAN	S&P/TSX Composite Index	
CHE	SPI Swiss Performance Index	
CHL	Santiago Stock Exchange IPSA Index	
CHN	Shanghai Stock Exchange Composite Index	
COL	FTSE All World Series Colombia Local	

(Continued)

GLOBAL CREDIT REVIEW VOLUME 3 109

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b1728 Global Credit Review Volume 3

Country	Stock Exchange	Period Used
СҮР	Cyprus Stock Exchange General Index	9/3/2004-Present
	Cyprus Stock Exchange General	4/2/1996-9/2/2004
CZE	Prague Stock Exch Index	4/5/1994-Present
DEU	CDAX Performance Index	
DNK	OMX Copenhagen 20 Index	
EGY	EGX 100 Index	5/1/2006-Present
ESP	IBEX 35 Index	
EST	OMX Tallinn OMXT	
FIN	OMX Helsinki Index	
FRA	CAC 40 Index	
GBR	FTSE 100 Index	
GRC	Athex Composite Share Price Index	
HKG	Hang Seng Index	
HRV	Croatia Zagreb CROBEX	6/14/2002-Present
HUN	Budapest Stock Exch Index	1/2/1991-Present
DN	Jakarta Composite Index	
ND	BSE Sensex 30 Index	
RL	Irish Overall Index	
ISL	OMX Iceland All-Share Price Index	
SR	Tel Aviv 100 Index	12/31/1991-Present
TA	Italy Stock Market BCI Comit Global	
IOR	Jordan Traded Value Index	11/10/2005-Present
IPN	Nikkei 500	
KAZ	Kazakhstan Stock Exchange Index KASE	7/12/2000-Present
KOR	KOSPI Index	
KWT	Kuwait SE Weighted Index	1/2/2012-Present
	Kuwait Global General Index	1/2/1984-1/2/2012
LKA	Sri Lanka Colombo All-Share Index	1/2/1985-Present
LTU	OMX Vilnius OMXV	1/4/2000-Present
LUX	Luxembourg Stock Exchange LuxX Index	1/4/1999-Present
	Luxembourg Stock Exchange 13 'Dead'	1/2/1998-1/3/1999
LVA	OMX GIRA OMXR	1/2/2000-Present
MAR	CFG 25	12/31/1993-Present
MEX	Mexico Bolsa Index	
MKD	Macedonian Stock Exchange MBI 10	12/30/2004-Present
MLT	Malta Stock Exchange	
MYS	FTSE Bursa Malaysia KLCI	
NGA	NIGERIA STCK EXC ALL SHR	1/30/1998-Present
NLD	AEX Index	
NOR	OBX Price Index	
NZL	NZX All Index	3/30/1992-Present
PAK	Karachi All Share Index	3/11/1998-Present
PER	Bolsa de Valores de Lima General Sector Index	
PHL	PSEI-Philippine Stock Exchange Index	
POL	WSE WIG Index	4/16/1991-Present

Table A.5.(Continued)

(Continued)

110 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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Country	Stock Exchange	Period Used
PRT	PSI General Index	
ROM	BSE COMPOSITE INDEX	4/17/1998-Present
RUS	MICEX INDEX	9/22/1997-Present
SAU	TADAWUL ALL SHARE INDEX	1/31/1994-Present
SGP	Straits Times Index	1/10/2008-Present
	Straits Times Old Index	8/31/1999-1/9/2008
SVK	Slovak Share Index	
SVN	HSBC Slovenia Dollar	
SWE	OMX Stockholm All-Share Index	
THA	Stock Exchange of Thailand Index	
TUR	Istanbul Stock Exchange National 100 Index	1/4/1988-Present
TWN	Taiwan Taiex Index	
UKR	Ukraine PFTS Index	1/12/1998-Present
USA	S&P 500 Index	
VEN	Caracas Stock Exchange Stock Market Index	
VNM	Ho Chi Minh Stock Index	7/28/2000-Present
ZAF	MSCI South Africa Index	12/31/1992-Present

 Table A.5.
 (Continued)

Note: *A blank Period Used column indicates that there is only a single index that is used throughout the whole period.

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GLOBAL CREDIT REVIEW VOLUME 3 111

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b1728 Global Credit Review Volume 3

Country	Short-Term Interest Rate	Period Used
ARE	UAE Ibor 3 Month	5/15/2000-Present
ARG	Argentina Deposit 90 Day	
AUS	Australia Dealer Bill 90 Day	
AUT	Germany 3 Month Bubill	1/1/1999-Present
	_	-12/31/1998
BEL	Germany 3 Month Bubill	1/1/1999-Present
	_	-12/31/1998
BGR	Bulgaria Interbank 3 Month	2/17/2003-Present
BHR	Bahrain Ibor 3 Month	12/14/2006
BRA	Andima Brazil Govt Bond Fixed Rate 3 Months	4/3/2000-Present
	Brazil CDB (up to 30 Days)	10/10/1994-3/31/2000
CAN	Canada Treasury Bill 3 Month	
CHE	_	
CHL	Chile TAB UF Interbank Rate 90 Days	
CHN	China Time Deposit Rate 3 Month	
COL	Colombia CD Rate 90-Day	
СҮР	Germany 3 Month Bubill	1/1/2008-Present
	_	-12/31/2007
CZE	Czech Republic Interbank 3 MTH	4/22/1992-Present
DEU	Germany 3 Month Bubill	5/25/1993-Present
	Germany Interbank 3 Month	1/2/1986-5/24/1993
DNK	Denmark Interbank 3 Month	
EGY	Egypt 91 Day T-Bill	7/6/2004-Present
ESP	Germany 3 Month Bubill	1/1/1999-Present
	_	-12/31/1998
EST	Germany 3 Month Bubill	1/1/2011-Present
	_	-12/31/2010
FIN	Germany 3 Month Bubill	1/1/1999-Present
	_	-12/31/1998
FRA	Germany 3 Month Bubill	1/1/1999-Present
	_	-12/31/1998
GBR	UK Treasury Bill Tender 3 Month	
GRC	Germany 3 Month Bubill	1/1/2001-Present
	—	-12/31/2000
HKG	Hong Kong Exchange Fund Bill 3 Month	
HRV	Croatia Zibor Rate 3 Month	6/2/1997-Present
HUN	Hungary Interbank 3 Month	9/7/1995-Present
IDN	Indonesia SBI 90 Day	7/10/2003-Present
	Indonesia SBI/DISC 90 Day	1/1/1985-7/9/2003
IND	India T-Bill Secondary 91 Day	
IRL	Germany 3 Month Bubill	1/1/1999-Present
	—	-12/31/1998
ISL	—	
ISR	Israel T-Bill Secondary 3 Month	5/30/1995-Present
ITA	Germany 3 Month Bubill	1/1/1999-Present

 Table A.6.
 The interest rates used for each economy as the second common variable.

 $112 \quad {\rm nus-rmi\ credit\ research\ initiative\ technical\ report}$

b1728_Ch-07.indd 112

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Country	Short-Term Interest Rate	Period Used		
	_	-12/31/1998		
JOR	Amman Interbank 3 Month	3/9/2001-Present		
JPN	Japan Treasury Discount Bills 3 Month	7/10/1992-Present		
	Japanese Government Bond Interest Rate-1 Year Maturity	9/24/1974-7/9/1992		
KAZ	Kazakhstan KIBOR/KIBID 90 Days Interbank	9/29/2001-Present		
KOR	Korea Commercial Paper 91 Day			
KWT	Kuwait Interbank 3 Month	8/17/1983-Present		
LKA	Sri Lanka Treasury Bill 3 Month	1/6/1989-Present		
LTU	VILNIUS Interbank Three Month	1/6/1999-Present		
LUX	Germany 3 Month Bubill	1/1/1999-Present		
	_	-12/31/1998		
LVA	Treasury Bill Rate 3 Month	5/11/1994-Present		
MAR	Morocco Deposit Rate 3 Month	6/6/2003-Present		
MEX	Mexico Cetes 2ND MKT. 90 Day	6/26/1996 - Present		
	Mexico Cetes 91 Dat AVG.RET.AT AUC.	3/9/1989-6/25/1996		
MKD	Macedonia Skibor 3 Months	7/2/2007 - Present		
MLT	Germany 3 Month Bubill	1/1/2008-Present		
	_	-12/31/2007		
MYS	Malaysia Deposit 3 Month			
NGA	Nigeria Interbank Offered Rate 3 Month	1/30/2004-Present		
NLD	Germany 3 Month Bubill	1/1/1999-Present		
	_	-12/31/1998		
NOR	Norway Govt Treasury Bills 3 Month	6/27/1995-Present		
	Norway Interbank 3 Month(effective)	1/2/1986-6/26/1995		
NZL	New Zealand Dollar Deposit 3 Month	9/27/1988-Present		
PAK	PKR 3 Month Repo	10/29/1999-Present		
PER	Peru Savings Rate			
PHL	Philippine Treasury Bill 91 Day			
POL	Poland Interbank 3 Month (EOD)	6/4/1993-Present		
PRT	Germany 3 Month Bubill	1/1/1999-Present		
	_	-12/31/1998		
ROM	Romanian Interbank 3 Month	8/1/1995-Present		
RUS	Russian Federation Interbank 31-90 Day	9/1/1994-Present		
SAU	Saudi Interbank 3 Month	1/1/1987-Present		
SGP	Singapore T-Bill 3 Month			
SVK	Germany 3 Month Bubill	1/1/2009-Present		
	_	-12/31/2008		
SVN	Germany 3 Month Bubill	1/1/2007-Present		
	_	-12/31/2006		
SWE	Sweden T-Bill 3 Month	5/25/1993-Present		
	Sweden Treasury Bill 90 Day	4/25/1989-5/24/1993		
THA	Thailand Repo 3 Month(BOT)			
TUR	Turkish Interbank 3 Month	8/1/2002-Present		
TWN	Taiwan Money Market 90 Day			
UKR	Ukraine Interbank 3 Months	3/1/2001-Present		

Table A.6.(Continued)

GLOBAL CREDIT REVIEW VOLUME 3 113

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Table A.6. (Continued)									
Country	Short-Term Interest Rate	Period Used							
USA	US Generic Govt 3-Month Yield								
VEN	Venezuela Overnight								
VNM	Vietnam Interbank 3 Month	12/11/1998-Present							
ZAF	South Africa T-Bill 91 Days (Tender Rates)	12/31/1980-Present							

Note: *A blank Period Used column indicates that there is only a single interest rate that is used throughout the whole period.

Country	Interest Rate Name	Period Used
ARE	UAE Ibor 1 Year	5/15/2000-Present
ARG	Aregentina Deposit 90 Day (PA.)	
AUS	Australia Govt. Bonds Generic Mid Yield 1 Year	
AUT	German Government Bonds 1 Year BKO	1/1/1999-Present
	Austria VIBOR 12 Month	6/10/1991-12/31/1998
BEL	German Government Bonds 1 Year BKO	1/1/1999-Present
	Belgium Treasury Bill 1 Year	4/2/1991-12/31/1998
BGR	Bulgaria Interbank 3 Month	2/17/2003-Present
BHR	Bahrain Ibor 1 Year	12/14/2006
BRA	Andima Brazil Govt Bond Fixed Rate 1 Year	4/3/2000-Present
	BRAZIL CDB (UP TO 30 DAYS)	10/10/1994-3/31/2000
CAN	Canada Treasury Bill 1 Year	
CHE	Swiss Interbank 1 Year (ZRC:SNB)	
CHL	Chile TAB UF Interbank Rates 360 Days	8/1/1996- Present
	Chile TAB UF Interbank Rate 90 Days	11/2/1992-7/30/1996
CHN	China Household Savings Deposits 1-Year Rate	
COL	Colombia Government Generic Bond 1 Year Yield	3/1/2001 - Present
	Colombia CD Rate 360-Dat	7/12/1993-2/8/2001
СҮР	Cyprus Treasury Bill Rate — 13 Week	
CZE	Czech Republic Interbank 3 MTH	4/22/1992-Present
DEU	German Government Bonds 1 Year BKO	1/10/1995-Present
	Germany Interbank 12 Month	11/2/1990-1/9/1995
DNK	Denmark Government Bonds 1 Year Note Generic Bid Yield	6/1/2008-Present
	Denmark Euro-Krone 1 Year(FT/ICAP/TR)	6/14/1985-5/31/2008
EGY	Egypt 364 Day T-Bill	7/6/2004-Present
ESP	German Government Bonds 1 Year BKO	1/1/1999-Present
	Spain 12 Month Treasury Bill Yield	11/30/1992-12/31/1998
	Spain Interbank 12 Month	12/19/1991-11/29/1992
EST	Estonia, Interest Rates, Prices, Production & Labour, Interest Rates, Deposit Rate	
FIN	German Government Bonds 1 Year BKO	1/1/1999-Present
	Finland Interbank Close 12 Month	4/2/1992-12/31/1998
FRA	German Government Bonds 1 Year BKO	1/1/1999-Present
	France Treasury Bill 12 Months	1/3/1989-12/31/1998
GBR	UK Govt. Bonds 1 Year Note Generic	9/12/2001-Present

Table A.7. The interest rates used for each economy in the DTD calculation.

(*Continued*)

114 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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Country	Interest Rate Name	Period Used
	UK Govt. Liability Nominal Spot Curve 12 Month	12/13/1985-9/11/2001
GRC	German Government Bonds 1 Year BKO	1/1/2001-Present
	Greece Treasury Bill 1 Year	1/2/1990-12/31/2000
IKG	HKMA Hong Kong Exchange Fund Bill 12 Month	
IRV	Croatia Zibor Rate 3 Month	6/2/1997-Present
IUN	Hungary Interbank 3 Month	9/7/1995-Present
DN	Indonesia SBI 90 Day	7/10/2003-Present
	Indonesia SBI/DISC 90 Day	1/1/1985-7/9/2003
ND	India T-Bill Secondary 1 Year	
RL	UK Govt. Liability Nominal Spot Curve 12 Month	
SL	Iceland Interbank 12 Month	2/1/2000-Present
	Iceland Interbank 3 Month	8/4/1998-1/31/2000
	Iceland 90-day CB Notes	5/12/1987-8/3/1998
SR	Israel T-Bill Secondary 1 Year	11/15/1994-Present
TA	German Government Bonds 1 Year BKO	1/1/1999-Present
	Italy Bots Treasury Bill 12 Month Gross Yields	9/5/1994-12/31/1998
	Italy T-Bill Auction Gross 12 Month	3/31/1987–9/4/1994
OR	Amman Interbank 1 Year	3/9/2001-Present
PN	Japan Treasury Bills 12 Month	12/14/1999-Present
	Japanese Government Bond Interest Rate-1 Year Maturity	9/24/1979-12/13/1999
AZ	Kazakhstan KIBOR/KIBID 90 Days Interbank	9/29/2001-Present
KOR	Korea Monetary Stabilization Bonds 1 Year	
TWT	Kuwait Interbank 1 Year	8/17/1983-Present
KA	Sri Lanka Fixed Deposit 1 Year	3/31/1988-Present
TU	Vilnius Interbank 12 Month	3/29/2000-Present
.UX	Long-Term Government Bond Yields — Maastricht Definition (Avg.)	
VA	Treasury Bill Rate 1 Year	4/3/1998-Present
/IAR	Morocco Deposit Rate 1 Year	6/6/2003-Present
/KD	Macedonia Skibor 3 Month	7/2/2007-Present
1EX	Mexico Cetes 2ND MKT. 360 Day	6/26/1996 -Present
	Mexico Cete 91 DAY AVG.RET.AT AUC.	3/9/1989- 6/25/1996
/ILT	Long-Term Government Bond Yields — Maastricht Definition (Avg.)	
/IYS	Bank Negara Malaysia 1 Year Govt. Securities Indicative YTM	6/21/2005-Present
	Malaysia Deposit 1 Year	1/1/1985-6/20/2005
IGA	Nigeria Interbank Offered Rate 3 Month	1/30/2004-Present
ILD	German Government Bonds 1 Year BKO	1/1/1999-Present
	Netherlands Interbank 1 Year	1/2/1987-12/31/1998
IOR	Norway Govt Treasury Bills 12 Month	7/1/1997-Present
	Norway Interbank 1 Year	1/2/1986-6/30/1997
ZL	New Zealand Dollar Deposit 1 Year	9/27/1988–Present
AK	PKR 12 Month Repo	10/29/2004-Present
'ER	Peru Savings Rate	
PHL	Philippine Treasury Bill 364 Day	
POL	Poland Interbank 1 Year (EOD)	10/11/1995-Present

 Table A.7. (Continued)

GLOBAL CREDIT REVIEW VOLUME 3 115

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Country	Interest Rate Name	Period Used
	Portugal 1-Year-LISBOR-Act/365 Day convention	8/16/1993-12/31/1998
ROM	Romanian Interbank 12 Month	8/1/1995-Present
RUS	Russian Federation Interbank 31–90 Day	9/1/1994-Present
SAU	Saudi Interbank 1 Year	1/1/1987-Present
SGP	Singapore T-Bill 3 Month	
SVK	Slovak Rep.Interbank 1 Year	
SVN	Slovenia Treasury Bill 3 Month 'Dead'	
SWE	Sweden Interbank 1 Year	5/25/1993–Present
	Sweden Treasury Bill 1 Year Note	4/25/1989-5/24/1993
THA	Thailand Govt. Bond 1 Year Note	8/7/2000-Present
	Thailand Deposit 12 Month(KT)	1/2/1991-8/6/2000
TUR	Turkish Interbank 12 Month	8/1/2002–Present
TWN	Taiwan Deposit 12 Month	
UKR	UAE Ibor 1 Year	5/15/2000-Present
USA	US Treasury Constant Maturities 1 Year	
VEN	Venezuela Overnight	
VNM	Vietnam Interbank 3 Month	12/11/1998-Present
ZAF	South Africa T-Bill 91 Days (Tender Rates)	12/31/1980-Present

Table A.7. (Continued)

Note: *A blank Period Used column indicates that there is only a single interest rate that is used throughout the whole period.

	DTD Level										
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations			
ARE	-0.80	1.77	2.76	3.83	13.33	3.07	1.88	5040			
ARG	-1.75	1.37	2.68	3.89	19.82	2.89	2.25	12432			
AUS	-1.43	1.85	2.98	4.20	18.39	3.29	2.18	279166			
AUT	-2.68	1.95	3.12	5.02	27.36	4.06	4.11	20255			
BEL	-2.68	2.52	4.38	6.73	27.36	5.03	3.77	28803			
BGR	-1.78	1.07	2.01	3.20	27.36	2.43	2.26	9953			
BHR	-0.27	1.69	2.73	4.57	18.32	3.60	2.97	1344			
BRA	-1.85	0.71	1.84	3.35	24.67	2.39	2.69	49196			
CAN	-1.13	1.91	3.26	4.93	24.77	3.69	2.57	212877			
CHE	-2.68	2.69	4.08	5.90	23.69	4.53	2.85	50300			
CHL	-1.85	3.58	5.22	6.77	25.75	5.62	3.44	28137			
CHN	0.05	3.09	4.17	5.74	16.75	4.66	2.29	261715			
COL	-1.35	2.34	3.86	5.70	20.21	4.31	3.03	5590			
СҮР	-1.19	0.89	1.54	2.46	23.81	2.07	2.28	16037			
CZE	-2.68	1.30	2.42	3.70	20.27	2.71	2.21	5949			
DEU	-2.68	1.62	2.89	4.38	27.36	3.31	2.68	174121			
								(Continued)			

Table A.8. Summary statistics of input variables (based on data from the statistics of input variables (based on data from the statistics of the statistics	om January 1991 to August 2013).
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116 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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	Table A.8. (Continued)										
				D	FD Level						
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations			
DNK	-1.92	1.88	3.17	4.70	27.36	3.65	2.96	41358			
EGY	-1.85	1.78	2.80	4.07	25.75	3.11	2.09	13226			
ESP	-2.68	2.07	3.46	4.98	27.36	3.87	3.11	34401			
EST	-0.30	1.40	2.89	3.99	11.63	3.27	2.53	714			
FIN	-2.68	2.27	3.44	4.93	16.52	3.79	2.34	27721			
FRA	-2.68	1.83	3.01	4.59	27.36	3.48	2.71	157441			
GBR	-2.68	2.23	3.57	5.29	27.36	4.03	2.72	373188			
GRC	-2.68	1.38	2.40	3.73	23.59	2.72	2.13	54384			
HKG	-1.43	1.52	2.54	3.97	18.39	3.03	2.29	204365			
HRV	-2.68	1.15	2.29	3.60	18.21	2.64	2.17	10054			
HUN	-1.53	1.62	2.74	4.31	27.36	3.14	2.41	7000			
IDN	-1.85	0.65	1.66	2.80	25.55	1.98	2.08	57698			
IND	-1.73	0.82	1.70	2.84	17.86	2.10	2.09	440591			
IRL	-1.73	1.96	3.26	4.78	14.51	3.49	2.23	9861			
ISL	-1.48	1.76	2.96	4.31	20.01	3.29	2.31	4305			
ISR	-2.19	1.26	2.35	3.61	27.36	2.73	2.40	71614			
ITA	-2.68	1.61	2.83	4.36	27.36	3.20	2.52	57802			
JOR	-1.09	2.37	3.46	4.98	15.49	3.99	2.49	11234			
JPN	-1.43	2.08	3.13	4.49	18.39	3.53	2.19	792157			
KAZ	-1.59	0.54	1.24	2.99	25.75	2.49	4.13	781			
KOR	-1.43	1.23	2.16	3.32	18.39	2.50	2.10	278669			
KWT	-0.44	2.28	3.23	4.47	25.75	3.70	2.30	21036			
LKA	-1.85	1.61	2.41	3.73	16.12	2.80	1.94	16090			
LTU	-1.30	1.43	3.13	5.21	20.95	3.68	3.19	4158			
LUX	-0.17	3.06	4.75	7.40	27.36	6.02	4.48	2800			
LVA	-1.45	1.30	2.27	3.89	27.36	2.90	2.54	2873			
MAR	-0.69	2.62	3.90	5.34	21.53	4.18	2.45	6959			
MEX	-1.85	2.01	3.64	5.50	25.75	4.13	3.19	18307			
MKD	-1.09	1.35	1.93	2.84	16.51	2.59	2.56	1826			
MLT	-0.65	2.30	3.52	4.86	14.99	4.11	3.01	892			
MYS	-1.85	1.56	2.83	4.59	25.75	3.49	2.91	183585			
NGA	-1.78	1.20	2.28	3.22	25.75	2.81	3.50	13202			
NLD	-2.68	2.44	3.96	5.77	27.36	4.38	3.00	35364			
NOR	-2.63	1.24	2.36	3.79	20.49	2.63	2.04	42178			
NZL	-1.30	2.87	4.88	6.94	18.39	5.21	3.20	16703			
PAK	-1.85	0.42	1.59	3.23	14.52	2.03	2.26	21899			
PER	-1.85	1.96	3.11	4.55	22.71	3.55	2.53	11048			
PHL	-1.85	1.11	2.25	3.65	25.75	2.66	2.30	37197			

GLOBAL CREDIT REVIEW VOLUME 3 117

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	Table A.8. (Continued)											
				D	TD Level							
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations				
POL	-1.94	1.37	2.31	3.46	27.36	2.63	2.06	55811				
PRT	-2.68	1.10	2.29	3.82	20.10	2.73	2.37	13389				
ROM	-2.68	0.98	1.80	2.38	27.36	1.92	1.60	15254				
RUS	-2.27	1.33	2.61	4.03	27.36	3.27	4.01	20091				
SAU	-1.52	3.59	5.23	7.49	25.75	6.05	3.67	13908				
SGP	-1.19	1.54	2.65	4.28	18.39	3.17	2.36	113973				
SVK	-0.39	1.62	2.41	3.04	27.36	3.36	5.06	905				
SVN	-2.47	2.10	3.39	5.49	16.88	3.92	2.95	6426				
SWE	-2.68	1.72	3.02	4.54	27.36	3.37	2.50	77340				
THA	-1.71	1.65	2.89	4.40	25.75	3.29	2.53	91012				
TUR	-2.34	1.61	2.87	4.57	22.59	3.47	2.82	37376				
TWN	-1.43	2.62	3.72	5.07	18.39	4.04	2.21	225465				
UKR	-1.69	0.98	1.70	2.76	21.52	1.89	1.53	4847				
USA	-1.13	1.79	3.04	4.70	24.77	3.53	2.62	1454952				
VEN	-1.85	0.35	1.21	2.65	17.01	2.05	3.01	3393				
VNM	-1.85	0.99	1.74	2.83	25.75	2.08	1.74	33109				
ZAF	-1.85	1.33	2.76	4.52	25.75	3.35	3.12	76694				

DTD Trend

	DID Trend									
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations		
ARE	-4.43	-0.37	0.02	0.38	6.00	-0.02	0.84	5040		
ARG	-7.73	-0.51	-0.01	0.41	7.38	-0.04	1.01	12432		
AUS	-5.68	-0.47	-0.02	0.38	5.40	-0.05	0.98	279166		
AUT	-8.45	-0.56	-0.03	0.42	7.80	-0.14	1.58	20255		
BEL	-8.45	-0.61	-0.01	0.59	7.80	-0.02	1.51	28803		
BGR	-8.45	-0.44	0.00	0.30	7.80	-0.08	1.01	9953		
BHR	-7.73	-0.35	0.01	0.26	4.36	-0.10	0.91	1344		
BRA	-7.73	-0.33	0.01	0.37	7.38	0.02	1.00	49196		
CAN	-6.37	-0.53	-0.02	0.45	5.44	-0.05	1.10	212877		
CHE	-8.45	-0.59	0.01	0.62	7.80	0.02	1.28	50300		
CHL	-7.73	-0.64	0.05	0.61	7.38	-0.01	1.49	28137		
CHN	-5.93	-0.56	-0.02	0.47	5.41	-0.06	1.03	261715		
COL	-7.73	-0.42	0.04	0.65	7.38	0.11	1.29	5590		
CYP	-8.45	-0.35	-0.07	0.17	7.80	-0.12	0.77	16037		
CZE	-7.78	-0.34	0.00	0.38	5.78	0.00	0.87	4949		
DEU	-8.45	-0.49	-0.03	0.42	7.80	-0.03	1.10	174121		
DNK	-8.45	-0.50	-0.01	0.43	7.80	-0.03	1.19	41358		
EGY	-7.73	-0.46	-0.01	0.45	7.38	-0.01	0.98	13226		
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118 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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	Table A.8. (Continued) DTD Trend										
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations			
ESP	-8.45	-0.49	0.01	0.47	7.80	-0.00	1.26	34401			
EST	-3.42	0.00	0.30	0.87	3.80	0.43	0.81	714			
FIN	-8.45	-0.45	0.04	0.54	7.80	0.04	1.04	27721			
FRA	-8.45	-0.46	0.00	0.45	7.80	-0.01	1.08	157441			
GBR	-8.45	-0.54	-0.02	0.42	7.80	-0.08	1.25	373188			
GRC	-8.45	-0.53	-0.08	0.32	7.80	-0.10	0.96	54384			
HKG	-5.68	-0.47	0.00	0.45	5.40	-0.02	0.97	204365			
HRV	-4.96	-0.55	-0.03	0.26	7.80	-0.12	0.93	10054			
HUN	-6.80	-0.43	0.00	0.40	7.80	-0.06	0.91	7000			
IDN	-7.73	-0.30	0.02	0.33	7.38	-0.00	0.77	57698			
IND	-6.55	-0.36	-0.03	0.33	5.21	-0.02	0.82	440591			
IRL	-6.45	-0.48	0.00	0.46	7.23	-0.04	1.01	9861			
ISL	-8.45	-0.69	-0.06	0.42	6.70	-0.17	1.35	4305			
ISR	-8.45	-0.43	0.00	0.40	7.80	-0.03	1.05	71614			
ITA	-8.45	-0.56	-0.02	0.47	7.80	-0.05	1.13	57802			
JOR	-7.38	-0.34	0.00	0.36	7.38	0.03	0.92	11234			
JPN	-5.68	-0.46	-0.01	0.43	5.40	-0.01	0.88	792157			
KAZ	-7.73	-0.50	0.00	0.40	7.38	0.01	1.16	781			
KOR	-5.68	-0.43	0.00	0.42	5.40	-0.01	0.91	278669			
KWT	-7.73	-0.44	0.00	0.40	7.38	-0.04	1.06	21036			
LKA	-7.73	-0.34	0.00	0.41	7.38	0.04	0.87	16090			
LTU	-6.02	-0.64	-0.01	0.65	7.80	0.01	1.35	4158			
LUX	-8.45	-0.64	0.01	0.55	7.80	-0.09	1.40	2800			
LVA	-8.45	-0.37	0.03	0.37	7.80	-0.04	1.11	2873			
MAR	-7.73	-0.59	-0.08	0.39	7.38	-0.10	1.09	6959			
MEX	-7.73	-0.42	0.06	0.61	7.38	0.07	1.18	18307			
MKD	-6.14	-0.37	-0.05	0.34	6.55	0.04	0.84	1826			
MLT	-6.66	-0.60	-0.04	0.55	4.26	0.02	1.27	892			
MYS	-7.73	-0.47	-0.01	0.42	7.38	-0.04	1.07	183585			
NGA	-7.73	-0.39	0.00	0.39	7.38	-0.01	1.42	13202			
NLD	-8.45	-0.63	-0.03	0.55	7.80	-0.05	1.22	35364			
NOR	-8.45	-0.41	-0.00	0.38	7.80	-0.03	0.90	42178			
NZL	-5.68	-0.57	0.03	0.63	5.40	0.02	1.42	16703			
PAK	-5.24	-0.26	0.03	0.35	6.17	0.04	0.74	21899			
PER	-7.73	-0.43	0.00	0.50	7.38	0.03	1.21	11048			
PHL	-7.73	-0.34	0.00	0.34	7.38	0.01	0.92	37197			
POL	-8.45	-0.49	-0.03	0.37	7.80	-0.08	0.91	55811			
PRT	-8.45	-0.43	-0.02	0.33	7.34	-0.04	0.94	13389			

GLOBAL CREDIT REVIEW VOLUME 3 119

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	Table A.8. (Continued) DTD Trend											
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations				
ROM	-8.45	-0.27	0.00	0.23	7.80	-0.01	0.73	15254				
RUS	-8.45	-0.46	0.00	0.42	7.80	-0.12	1.54	20091				
SAU	-7.73	-0.74	0.13	1.01	7.38	0.10	1.83	13908				
SGP	-5.68	-0.44	-0.01	0.41	5.40	-0.03	0.96	113973				
SVK	-8.45	-0.15	0.06	0.31	7.80	0.04	1.39	905				
SVN	-5.13	-0.57	-0.10	0.20	7.80	-0.17	1.00	6426				
SWE	-8.45	-0.47	-0.02	0.44	7.80	-0.02	1.04	77340				
THA	-7.73	-0.51	-0.01	0.45	7.38	-0.04	1.05	91012				
TUR	-8.45	-0.50	0.09	0.66	7.80	0.09	1.31	37376				
TWN	-5.68	-0.53	0.00	0.56	5.40	0.01	1.02	225465				
UKR	-5.71	-0.51	-0.00	0.33	6.49	-0.14	0.90	4847				
USA	-6.37	-0.47	0.00	0.47	5.44	-0.01	0.99	1454952				
VEN	-6.72	-0.29	-0.00	0.28	7.38	0.00	0.90	3393				
VNM	-7.73	-0.39	-0.05	0.24	7.38	-0.07	0.66	33109				
ZAF	-7.73	-0.45	-0.01	0.37	7.38	-0.06	1.18	76694				
				CAS	H/TA Leve	l						
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations				
ARE	0.00	0.07	0.15	0.23	0.94	0.17	0.14	6142				
ARG	0.00	0.02	0.05	0.11	0.69	0.08	0.08	12964				
AUS	0.00	0.04	0.13	0.35	0.97	0.23	0.25	289939				
AUT	0.00	0.03	0.07	0.15	0.96	0.11	0.13	21975				
BEL	0.00	0.03	0.07	0.18	0.99	0.14	0.18	31291				
BGR	0.00	0.01	0.03	0.08	0.58	0.06	0.08	10253				
BHR	0.00	0.09	0.18	0.27	0.91	0.20	0.15	2778				
BRA	0.00	0.02	0.08	0.17	0.94	0.12	0.13	52443				
CAN	0.00	0.01	0.06	0.21	0.99	0.16	0.21	218264				
CHE	0.00	0.05	0.10	0.20	0.99	0.15	0.16	54975				
CHL	0.00	0.01	0.03	0.08	0.94	0.06	0.09	29891				
CHN	0.00	0.08	0.14	0.25	0.89	0.19	0.16	267324				
COL	0.00	0.03	0.06	0.09	0.76	0.07	0.08	6101				
СҮР	0.00	0.01	0.05	0.15	0.93	0.11	0.14	16617				
CZE	0.00	0.02	0.05	0.11	0.99	0.09	0.13	6647				
DEU	0.00	0.03	0.08	0.19	0.99	0.14	0.18	181477				
DNK	0.00	0.03	0.09	0.18	0.99	0.14	0.17	45170				
EGY	0.00	0.04	0.12	0.23	0.94	0.16	0.14	14709				
ESP	0.00	0.02	0.05	0.11	0.82	0.08	0.10	38443				
EST	0.00	0.03	0.05	0.12	0.53	0.09	0.09	2467				

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120 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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				ole A.8. (C	,			CASH/TA Level										
	Min	25%	Median	75%	H/TA Level Max	Mean	StdDev	# Observations										
FIN	0.00	0.03	0.08	0.16	0.99	0.12	0.14	29245										
FRA	0.00	0.03	0.08	0.17	0.99	0.13	0.14	163863										
GBR	0.00	0.03	0.09	0.22	0.99	0.17	0.21	378789										
GRC	0.00	0.02	0.06	0.13	0.83	0.10	0.11	56058										
HKG	0.00	0.06	0.14	0.26	0.97	0.19	0.17	210091										
HRV	0.00	0.01	0.02	0.05	0.52	0.05	0.08	11945										
HUN	0.00	0.02	0.06	0.13	0.74	0.09	0.10	7476										
IDN	0.00	0.03	0.08	0.17	0.90	0.12	0.12	61228										
IND	0.00	0.01	0.03	0.07	0.81	0.06	0.09	602790										
IRL	0.00	0.05	0.09	0.21	0.97	0.15	0.17	10353										
ISL	0.00	0.02	0.04	0.08	0.53	0.06	0.06	4725										
ISR	0.00	0.03	0.10	0.22	0.99	0.18	0.21	72757										
ITA	0.00	0.03	0.06	0.14	0.99	0.10	0.11	61244										
JOR	0.00	0.01	0.04	0.14	0.94	0.11	0.15	13308										
JPN	0.00	0.08	0.13	0.22	0.97	0.17	0.14	795679										
KAZ	0.00	0.08	0.14	0.18	0.36	0.13	0.07	922										
KOR	0.00	0.04	0.09	0.18	0.97	0.13	0.13	282260										
KWT	0.00	0.03	0.07	0.20	0.94	0.15	0.18	22089										
LKA	0.00	0.02	0.04	0.10	0.94	0.09	0.13	16415										
LTU	0.00	0.01	0.03	0.07	0.51	0.06	0.09	4325										
LUX	0.00	0.05	0.11	0.18	0.97	0.15	0.14	3191										
LVA	0.00	0.01	0.04	0.12	0.44	0.08	0.09	3146										
MAR	0.00	0.01	0.05	0.13	0.78	0.09	0.11	10481										
MEX	0.00	0.03	0.06	0.11	0.77	0.08	0.08	19948										
MKD	0.00	0.02	0.04	0.16	0.59	0.11	0.13	2324										
MLT	0.00	0.04	0.08	0.21	0.50	0.14	0.14	1330										
MYS	0.00	0.02	0.07	0.16	0.94	0.12	0.13	186337										
NGA	0.00	0.03	0.07	0.20	0.73	0.13	0.15	14526										
NLD	0.00	0.02	0.05	0.12	0.99	0.10	0.13	37561										
NOR	0.00	0.04	0.09	0.19	0.99	0.15	0.18	44342										
NZL	0.00	0.01	0.03	0.10	0.97	0.10	0.17	17600										
PAK	0.00	0.01	0.05	0.14	0.90	0.10	0.12	28455										
PER	0.00	0.01	0.04	0.13	0.71	0.09	0.11	11736										
PHL	0.00	0.02	0.08	0.18	0.94	0.13	0.15	38970										
POL	0.00	0.02	0.06	0.14	0.99	0.11	0.12	57259										
PRT	0.00	0.01	0.03	0.07	0.54	0.06	0.08	14419										
ROM	0.00	0.01	0.03	0.07	0.73	0.00	0.10	15955										

GLOBAL CREDIT REVIEW VOLUME 3 121

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				CAS	H/TA Leve	1		
	Min	25%	Median	75%	Max	Mean	StdDev	# Observation
RUS	0.00	0.02	0.06	0.15	0.99	0.11	0.12	22186
SAU	0.00	0.04	0.09	0.18	0.94	0.15	0.18	14598
SGP	0.00	0.06	0.13	0.24	0.97	0.17	0.15	115827
SVK	0.00	0.02	0.05	0.11	0.59	0.08	0.09	1414
SVN	0.00	0.01	0.04	0.08	0.41	0.06	0.07	7000
SWE	0.00	0.04	0.09	0.20	0.99	0.16	0.19	79823
THA	0.00	0.02	0.06	0.14	0.88	0.10	0.12	93020
TUR	0.00	0.02	0.06	0.15	0.99	0.11	0.14	60029
TWN	0.00	0.06	0.11	0.21	0.95	0.16	0.14	227903
UKR	0.00	0.01	0.02	0.07	0.88	0.06	0.12	5636
USA	0.00	0.03	0.07	0.24	0.99	0.18	0.22	1509399
VEN	0.00	0.04	0.07	0.18	0.94	0.12	0.11	3937
VNM	0.00	0.03	0.08	0.17	0.93	0.13	0.14	34586
ZAF	0.00	0.03	0.08	0.16	0.94	0.12	0.14	80493
				CAS	H/TA Trend	1		
	Min	25%	Median	75%	Max	Mean	StdDev	# Observation
ARE	-0.36	-0.02	-0.00	0.01	0.40	-0.01	0.06	6142
ARG	-0.36	-0.01	0.00	0.01	0.40	-0.00	0.04	12964
AUS	-0.42	-0.03	-0.00	0.01	0.44	-0.01	0.09	289939
AUT	-0.46	-0.01	0.00	0.00	0.46	-0.00	0.04	21975
BEL	-0.46	-0.01	0.00	0.00	0.46	-0.00	0.05	31291
BGR	-0.25	-0.00	0.00	0.00	0.46	-0.00	0.03	10253
BHR	-0.36	-0.02	0.00	0.01	0.40	-0.00	0.07	2778
BRA	-0.36	-0.01	0.00	0.01	0.40	-0.00	0.05	52443
CAN	-0.44	-0.02	0.00	0.01	0.42	-0.00	0.07	218264
CHE	-0.46	-0.01	0.00	0.01	0.46	-0.00	0.04	54975
CHL	-0.36	-0.01	-0.00	0.00	0.40	-0.00	0.04	29891
CHN	-0.30	-0.03	-0.00	0.01	0.30	-0.01	0.05	267324
COL	-0.36	-0.01	0.00	0.01	0.40	0.00	0.04	6101
СҮР	-0.46	-0.01	0.00	0.00	0.46	-0.00	0.04	16617
CZE	-0.33	-0.00	0.00	0.00	0.46	0.00	0.04	6647
DEU	-0.46	-0.01	0.00	0.00	0.46	-0.00	0.06	181477
DNK	-0.46	-0.01	-0.00	0.01	0.46	-0.00	0.06	45170
EGY	-0.36	-0.02	-0.00	0.01	0.40	-0.00	0.05	14709
ESP	-0.46	-0.01	0.00	0.01	0.46	-0.00	0.04	38443
EST	-0.25	-0.01	0.00	0.01	0.17	-0.00	0.03	2467
FIN	-0.46	-0.01	-0.00	0.01	0.46	-0.00	0.05	29245

122 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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					Continued) H/TA Trend	1		
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations
FRA	-0.46	-0.01	0.00	0.01	0.46	-0.00	0.04	163863
GBR	-0.46	-0.02	0.00	0.01	0.46	-0.01	0.07	378789
GRC	-0.46	-0.01	-0.00	0.01	0.46	-0.00	0.05	56058
HKG	-0.42	-0.02	0.00	0.01	0.44	-0.00	0.07	210091
HRV	-0.18	-0.01	-0.00	0.00	0.44	-0.00	0.03	11945
HUN	-0.46	-0.01	-0.00	0.01	0.46	-0.00	0.04	7476
IDN	-0.36	-0.01	0.00	0.01	0.40	-0.00	0.04	61228
IND	-0.35	-0.00	0.00	0.00	0.36	-0.00	0.04	602790
IRL	-0.46	-0.01	0.00	0.01	0.46	-0.00	0.05	10353
ISL	-0.36	-0.01	0.00	0.00	0.40	-0.00	0.03	4725
ISR	-0.46	-0.02	-0.00	0.01	0.46	-0.00	0.08	72757
ITA	-0.46	-0.01	-0.00	0.01	0.46	-0.00	0.04	61244
JOR	-0.36	-0.01	0.00	0.00	0.40	-0.00	0.04	13308
JPN	-0.42	-0.01	0.00	0.01	0.44	-0.00	0.04	795679
KAZ	-0.17	-0.02	0.00	0.01	0.30	-0.00	0.04	922
KOR	-0.42	-0.02	-0.00	0.01	0.44	-0.00	0.06	282260
KWT	-0.36	-0.01	-0.00	0.01	0.40	-0.00	0.06	22089
LKA	-0.36	-0.01	0.00	0.01	0.40	0.00	0.05	16415
LTU	-0.20	-0.01	-0.00	0.00	0.31	-0.00	0.03	4325
LUX	-0.39	-0.01	0.00	0.00	0.26	0.00	0.04	3191
LVA	-0.21	-0.01	0.00	0.01	0.32	-0.00	0.04	3146
MAR	-0.36	-0.01	0.00	0.01	0.40	-0.00	0.04	10481
MEX	-0.30	-0.01	-0.00	0.01	0.40	-0.00	0.03	19948
MKD	-0.18	-0.00	0.00	0.00	0.31	0.00	0.04	2324
MLT	-0.32	-0.01	0.00	0.00	0.18	-0.00	0.03	1330
MYS	-0.36	-0.01	0.00	0.01	0.40	-0.00	0.05	186337
NGA	-0.36	-0.01	0.00	0.00	0.40	-0.00	0.06	14526
NLD	-0.46	-0.01	0.00	0.00	0.46	-0.00	0.04	37561
NOR	-0.46	-0.02	-0.00	0.01	0.46	-0.00	0.06	44342
NZL	-0.42	-0.01	0.00	0.00	0.44	-0.00	0.06	17600
PAK	-0.36	-0.01	0.00	0.00	0.40	-0.00	0.04	28455
PER	-0.32	-0.01	0.00	0.01	0.39	0.00	0.04	11736
PHL	-0.36	-0.01	0.00	0.01	0.40	-0.00	0.06	38970
POL	-0.46	-0.01	-0.00	0.00	0.46	-0.00	0.05	57259
PRT	-0.40	-0.01	0.00	0.00	0.46	-0.00	0.03	14419
ROM	-0.46	-0.00	0.00	0.00	0.46	-0.00	0.03	15955
RUS	-0.46	-0.01	0.00	0.01	0.46	0.00	0.06	22186

 Table A.8.
 (Continued)

(Continued)

GLOBAL CREDIT REVIEW VOLUME 3 123

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			Tal	ble A.8. ((Continued)						
					H/TA Trend						
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations			
SAU	-0.36	-0.01	0.00	0.01	0.40	-0.00	0.06	14598			
SGP	-0.42	-0.02	0.00	0.01	0.44	-0.00	0.06	115827			
SVK	-0.13	-0.01	0.00	0.00	0.15	-0.00	0.02	1414			
SVN	-0.30	-0.00	0.00	0.00	0.28	-0.00	0.02	7000			
SWE	-0.46	-0.01	-0.00	0.01	0.46	-0.00	0.06	79823			
THA	-0.36	-0.01	-0.00	0.01	0.40	-0.00	0.04	93020			
TUR	-0.46	-0.01	0.00	0.01	0.46	-0.00	0.06	60029			
TWN	-0.42	-0.02	0.00	0.02	0.44	0.00	0.05	227903			
UKR	-0.23	-0.00	0.00	0.00	0.32	0.00	0.03	5636			
USA	-0.44	-0.02	-0.00	0.01	0.42	-0.00	0.06	1509399			
VEN	-0.18	-0.01	0.00	0.00	0.31	-0.00	0.03	3937			
VNM	-0.36	-0.02	-0.00	0.01	0.40	-0.00	0.05	34586			
ZAF	-0.36	-0.01	0.00	0.01	0.40	-0.00	0.05	80493			
	NI/TA Level										
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations			
ARE*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	6180			
ARG*	-0.04	-0.00	0.00	0.01	0.03	0.00	0.01	13017			
AUS	-0.47	-0.02	-0.00	0.00	0.10	-0.02	0.05	290464			
AUT	-0.56	0.00	0.00	0.00	0.08	-0.00	0.02	22103			
BEL	-0.35	0.00	0.00	0.01	0.08	0.00	0.01	31396			
BGR	-0.19	-0.00	0.00	0.01	0.08	0.00	0.01	11193			
BHR*	-0.03	0.00	0.00	0.01	0.03	0.01	0.01	2825			
BRA*	-0.04	-0.00	0.00	0.01	0.03	0.00	0.01	52496			
CAN	-0.39	-0.01	0.00	0.00	0.20	-0.01	0.04	218860			
CHE	-0.56	0.00	0.00	0.00	0.08	0.00	0.02	55184			
CHL*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	29989			
CHN	-0.17	0.00	0.00	0.01	0.13	0.00	0.01	267646			
COL*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	6149			
СҮР	-0.56	-0.00	0.00	0.00	0.08	-0.00	0.03	17142			
CZE	-0.29	0.00	0.00	0.00	0.04	0.00	0.01	6688			
DEU	-0.56	-0.00	0.00	0.00	0.08	-0.00	0.02	182385			
DNK	-0.56	0.00	0.00	0.00	0.08	-0.00	0.03	45393			
EGY*	-0.04	0.00	0.00	0.01	0.03	0.01	0.01	14769			
ESP	-0.56	0.00	0.00	0.00	0.08	0.00	0.03	38509			
EST	-0.09	0.00	0.00	0.01	0.05	0.00	0.01	2489			
FIN	-0.22	0.00	0.00	0.01	0.08	0.00	0.01	29292			
FRA	-0.56	0.00	0.00	0.00	0.08	0.00	0.02	164611			

(Continued)

124 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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	Table A.8. (Continued)											
				N	I/TA Level							
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations				
GBR	-0.56	-0.00	0.00	0.01	0.08	-0.01	0.04	379983				
GRC	-0.56	-0.00	0.00	0.01	0.08	0.00	0.01	56167				
HKG	-0.47	-0.00	0.00	0.01	0.10	-0.00	0.03	210117				
HRV	-0.11	-0.00	0.00	0.00	0.08	0.00	0.01	12210				
HUN	-0.13	-0.00	0.00	0.01	0.03	0.00	0.01	7500				
IDN*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	61292				
IND*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	608425				
IRL	-0.56	-0.00	0.00	0.01	0.08	0.00	0.02	10395				
ISL	-0.13	0.00	0.00	0.01	0.02	0.00	0.01	4758				
ISR	-0.56	-0.00	0.00	0.00	0.08	-0.01	0.05	72805				
ITA	-0.23	-0.00	0.00	0.00	0.08	0.00	0.01	61297				
JOR*	-0.04	-0.00	0.00	0.00	0.03	0.00	0.01	13325				
JPN	-0.47	0.00	0.00	0.00	0.10	0.00	0.01	795722				
KAZ*	-0.04	0.00	0.00	0.00	0.03	0.00	0.01	935				
KOR	-0.47	-0.00	0.00	0.01	0.10	-0.00	0.02	285324				
KWT*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	22170				
LKA*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	16489				
LTU	-0.05	-0.00	0.00	0.01	0.05	0.00	0.01	4335				
LUX	-0.36	0.00	0.00	0.01	0.08	-0.00	0.04	3341				
LVA	-0.06	-0.00	0.00	0.01	0.08	0.00	0.01	3266				
MAR*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	10552				
MEX*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	20085				
MKD	-0.50	0.00	0.00	0.00	0.03	-0.00	0.03	2421				
MLT	-0.14	0.00	0.00	0.00	0.04	0.00	0.01	1334				
MYS*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	186385				
NGA*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	14838				
NLD	-0.56	0.00	0.00	0.01	0.08	0.00	0.02	37588				
NOR	-0.56	-0.00	0.00	0.00	0.08	-0.00	0.03	44583				
NZL	-0.47	0.00	0.00	0.01	0.10	-0.01	0.05	17627				
PAK*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	28534				
PER*	-0.04	0.00	0.00	0.01	0.03	0.01	0.01	11808				
PHL*	-0.04	-0.00	0.00	0.01	0.03	0.00	0.01	39045				
POL	-0.36	-0.00	0.00	0.01	0.08	0.00	0.02	57462				
PRT	-0.22	-0.00	0.00	0.00	0.06	0.00	0.01	14528				
ROM	-0.56	-0.00	0.00	0.01	0.08	0.00	0.02	17822				
RUS	-0.56	0.00	0.00	0.01	0.08	0.00	0.02	22582				
SAU*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	14649				

 Table A.8.
 (Continued)

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GLOBAL CREDIT REVIEW VOLUME 3 125

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			Tab	ole A.8. (C	Continued)						
				NI	/TA Level						
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations			
SGP	-0.47	0.00	0.00	0.01	0.10	0.00	0.02	115917			
SVK	-0.02	0.00	0.00	0.00	0.03	0.00	0.01	1503			
SVN	-0.08	0.00	0.00	0.00	0.03	0.00	0.01	7072			
SWE	-0.56	-0.01	0.00	0.01	0.08	-0.01	0.03	80221			
THA*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	93064			
TUR	-0.56	-0.00	0.00	0.01	0.08	0.00	0.03	60041			
TWN	-0.29	0.00	0.00	0.01	0.10	0.00	0.01	227950			
UKR	-0.10	-0.00	0.00	0.01	0.06	0.00	0.01	5789			
USA	-0.39	-0.00	0.00	0.01	0.20	-0.00	0.03	1508765			
VEN*	-0.04	0.00	0.00	0.01	0.03	0.00	0.01	3971			
VNM*	-0.04	0.00	0.00	0.01	0.03	0.01	0.01	34835			
ZAF*	-0.04	0.00	0.01	0.01	0.03	0.00	0.01	80780			
	NI/TA Trend										
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations			
ARE*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	6180			
ARG*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	13017			
AUS	-0.37	-0.00	0.00	0.00	0.28	-0.00	0.04	290464			
AUT	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.01	22103			
BEL	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.01	31396			
BGR	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.02	11193			
BHR*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.00	2825			
BRA*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	52496			
CAN	-0.31	-0.00	0.00	0.00	0.24	0.00	0.03	218860			
CHE	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.01	55184			
CHL*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	29989			
CHN	-0.20	-0.00	-0.00	0.00	0.14	-0.00	0.01	267646			
COL*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.00	6149			
СҮР	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.02	17142			
CZE	-0.27	-0.00	0.00	0.00	0.26	-0.00	0.01	6688			
DEU	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.02	182385			
DNK	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.02	45393			
EGY*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	14769			
ESP	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.02	38509			
EST	-0.32	-0.00	0.00	0.00	0.11	-0.00	0.02	2489			
FIN	-0.20	-0.00	0.00	0.00	0.29	-0.00	0.01	29292			
FRA	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.01	164611			

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126 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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	Table A.8. (Continued) NI/TA Trend										
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations			
GBR	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.03	379983			
GRC	-0.33	-0.00	-0.00	0.00	0.29	-0.00	0.01	56167			
HKG	-0.37	-0.00	0.00	0.00	0.28	-0.00	0.03	210117			
HRV	-0.22	-0.00	0.00	0.00	0.10	-0.00	0.01	12210			
HUN	-0.10	-0.00	0.00	0.00	0.11	-0.00	0.01	7500			
IDN*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	61292			
IND	-0.13	-0.00	0.00	0.00	0.11	-0.00	0.01	608425			
IRL	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.02	10395			
ISL	-0.12	-0.00	0.00	0.00	0.13	-0.00	0.01	4758			
ISR	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.03	72805			
ITA	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.01	61297			
JOR*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	13325			
JPN	-0.37	-0.00	0.00	0.00	0.28	-0.00	0.01	795722			
KAZ*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	935			
KOR	-0.37	-0.00	0.00	0.00	0.28	-0.00	0.03	285324			
KWT*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	22170			
LKA*	-0.03	-0.00	0.00	0.00	0.03	0.00	0.01	16489			
LTU	-0.12	-0.00	0.00	0.00	0.11	-0.00	0.01	4335			
LUX	-0.09	-0.00	0.00	0.00	0.22	0.00	0.01	3341			
LVA	-0.19	-0.00	0.00	0.00	0.08	-0.00	0.01	3266			
MAR*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.00	10552			
MEX*	-0.03	-0.00	0.00	0.00	0.03	0.00	0.01	20085			
MKD	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.02	2421			
MLT	-0.04	-0.00	-0.00	0.00	0.03	-0.00	0.00	1334			
MYS*	-0.03	-0.00	-0.00	0.00	0.03	-0.00	0.01	186385			
NGA*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	14838			
NLD	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.02	37588			
NOR	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.02	44583			
NZL	-0.37	-0.00	0.00	0.00	0.28	-0.00	0.03	17627			
PAK*	-0.03	-0.00	0.00	0.00	0.03	0.00	0.01	28534			
PER*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	11808			
PHL*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	39045			
POL	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.02	57462			
PRT	-0.33	-0.00	0.00	0.00	0.21	-0.00	0.01	14528			
ROM	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.02	17822			
RUS	-0.33	-0.00	0.00	0.00	0.29	0.00	0.02	22582			
SAU*	-0.03	-0.00	0.00	0.00	0.03	0.00	0.01	14649			
SGP	-0.37	-0.00	-0.00	0.00	0.28	-0.00	0.02	115917			

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GLOBAL CREDIT REVIEW VOLUME 3 127

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			Tab		Continued) /TA Trend						
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations			
SVK	-0.05	-0.00	0.00	0.00	0.06	-0.00	0.01	1503			
SVN	-0.17	-0.00	0.00	0.00	0.06	-0.00	0.01	7072			
SWE	-0.33	-0.00	0.00	0.00	0.29	-0.00	0.03	80221			
THA*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	93064			
TUR	-0.33	-0.00	-0.00	0.00	0.29	-0.00	0.02	60041			
TWN	-0.37	-0.00	-0.00	0.00	0.28	-0.00	0.01	227950			
UKR	-0.11	-0.00	0.00	0.00	0.13	-0.00	0.01	5789			
USA	-0.31	-0.00	0.00	0.00	0.24	-0.00	0.02	1508765			
VEN*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.00	3971			
VNM*	-0.03	-0.00	-0.00	0.00	0.03	-0.00	0.01	34835			
ZAF*	-0.03	-0.00	0.00	0.00	0.03	-0.00	0.01	80780			
	SIZE Level										
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations			
ARE	-4.78	-0.86	0.09	1.15	4.25	0.15	1.50	7005			
ARG	-6.58	-1.42	0.26	1.58	7.06	0.14	2.07	14330			
AUS	-6.46	-1.22	-0.11	1.54	6.96	0.31	2.05	312426			
AUT	-6.75	-1.37	-0.12	1.32	4.50	-0.07	2.00	23602			
BEL	-6.75	-1.39	0.10	1.52	6.91	0.12	2.26	37742			
BGR	-6.75	-1.70	-0.33	0.91	8.06	-0.37	1.84	16735			
BHR	-3.60	-1.16	-0.29	0.98	3.26	-0.12	1.50	3451			
BRA	-6.58	-1.74	-0.08	1.31	7.45	-0.15	2.58	59280			
CAN	-5.98	-1.50	-0.22	1.26	6.00	-0.08	2.08	239767			
CHE	-6.75	-1.26	-0.02	1.30	6.31	0.10	1.95	54559			
CHL	-6.58	-1.09	0.10	1.28	4.30	0.01	1.83	33043			
CHN	-2.48	-0.74	-0.25	0.30	3.79	-0.15	0.88	289226			
COL	-5.42	-1.45	-0.09	1.09	4.43	-0.25	1.68	7196			
СҮР	-4.64	-1.04	0.00	1.02	6.87	0.01	1.63	20496			
CZE	-6.75	-1.51	-0.16	0.89	5.36	-0.23	1.91	8939			
DEU	-6.75	-0.37	1.13	2.78	8.06	1.19	2.50	210512			
DNK	-6.75	-0.27	0.91	2.23	7.41	1.04	1.89	46094			
EGY	-6.58	-1.01	0.05	1.52	5.40	0.24	1.75	17615			
ESP	-6.75	-1.66	-0.27	1.15	5.31	-0.31	2.11	40961			
EST	-3.55	-0.45	0.40	1.59	5.24	0.54	1.73	2663			
FIN	-6.36	-1.78	-0.49	1.08	6.40	-0.33	1.97	29719			
FRA	-6.75	-1.32	0.12	1.87	7.66	0.36	2.31	190957			
GBR	-6.75	-1.14	0.24	1.86	8.06	0.46	2.23	412607			
GRC	-6.08	-0.44	0.50	1.59	6.55	0.65	1.60	59060			

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128 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

b1728_Ch-07.indd 128

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	Table A.8. (Continued)										
				SI	ZE Level						
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations			
HKG	-8.79	-1.51	-0.50	0.86	6.96	-0.21	1.85	224713			
HRV	-6.75	-0.70	0.42	1.63	6.10	0.48	1.79	15242			
HUN	-6.75	-0.87	0.88	2.43	6.49	0.83	2.26	8352			
IDN	-6.58	-1.04	0.13	1.40	6.08	0.24	1.83	68953			
IND	-5.23	-1.20	0.23	2.04	8.38	0.55	2.32	531335			
IRL	-6.75	-2.05	-0.83	0.62	4.79	-0.72	1.95	10816			
ISL	-6.75	-2.00	-1.07	-0.12	2.76	-1.08	1.53	5886			
ISR	-6.75	-0.76	0.30	1.60	8.06	0.51	1.84	95397			
ITA	-6.75	-0.90	0.22	1.64	6.36	0.41	1.92	61869			
JOR	-3.72	-0.80	0.01	1.18	6.76	0.27	1.57	16405			
JPN	-9.57	-0.79	0.25	1.54	6.96	0.48	1.73	817152			
KAZ	-6.07	-2.18	-0.62	1.17	5.50	-0.56	1.92	1465			
KOR	-11.49	-0.50	0.29	1.34	6.96	0.48	1.82	333045			
KWT	-2.55	-0.30	0.61	1.44	5.14	0.68	1.34	24561			
LKA	-6.58	-0.79	0.15	1.30	5.26	0.28	1.53	18554			
LTU	-4.47	-0.87	0.19	1.18	4.08	0.17	1.53	5661			
LUX	-6.75	-2.49	-0.95	0.19	4.33	-1.09	2.07	4479			
LVA	-5.38	-1.46	-0.22	2.18	5.91	0.24	2.40	4654			
MAR	-6.58	-1.28	-0.15	1.66	4.76	0.10	1.81	11086			
MEX	-6.58	-1.19	0.19	1.55	5.16	0.11	1.97	21875			
MKD	-6.46	-1.30	0.19	1.34	5.35	0.09	1.87	4151			
MLT	-4.07	-1.00	-0.20	0.98	2.31	-0.07	1.32	1873			
MYS	-4.25	-0.18	0.69	1.80	6.47	0.85	1.56	198058			
NGA	-6.58	-1.39	-0.17	1.65	6.16	0.01	2.16	17787			
NLD	-6.75	-1.89	-0.37	1.09	5.99	-0.28	2.17	38059			
NOR	-6.75	-0.92	0.15	1.39	6.65	0.26	1.73	47416			
NZL	-5.78	-1.54	-0.08	1.05	5.12	-0.22	1.88	19565			
PAK	-6.58	-1.08	0.86	2.82	7.45	0.83	2.55	46403			
PER	-6.58	-1.02	0.34	1.82	5.54	0.35	2.02	14133			
PHL	-6.58	-1.43	-0.31	1.04	5.06	-0.11	1.80	42721			
POL	-5.28	-1.01	0.20	1.59	8.06	0.35	2.00	67405			
PRT	-6.75	-1.95	-0.32	1.28	4.56	-0.43	2.47	16343			
ROM	-6.75	-0.99	0.26	1.59	8.06	0.26	2.11	45558			
RUS	-6.75	-1.72	-0.25	1.28	8.06	-0.06	2.45	29503			
SAU	-4.48	-0.76	0.15	1.46	5.34	0.41	1.59	16068			
SGP	-4.35	-0.61	0.36	1.60	6.96	0.59	1.70	124879			
SVK	-6.14	-0.30	1.17	3.08	8.00	1.56	2.57	3941			

GLOBAL CREDIT REVIEW VOLUME 3 129

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	Table A.8. (Continued)											
					ZE Level							
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations				
SVN	-6.75	-0.50	0.86	2.37	8.06	1.16	2.44	10856				
SWE	-6.75	-0.67	1.02	2.62	8.06	1.09	2.38	85358				
THA	-5.98	-0.88	0.09	1.23	6.49	0.29	1.61	102574				
TUR	-5.21	-1.22	-0.03	1.24	5.62	0.09	1.84	64054				
TWN	-5.78	-0.70	0.26	1.28	6.89	0.37	1.51	250096				
UKR	-6.75	-0.99	0.10	1.06	8.06	-0.01	1.63	8699				
USA	-5.98	-2.00	-0.67	0.77	6.00	-0.54	2.00	1569893				
VEN	-6.58	-1.52	-0.08	1.18	7.45	-0.33	2.55	5337				
VNM	-4.93	-1.16	-0.22	0.81	6.29	-0.06	1.61	40028				
ZAF	-6.58	-1.57	0.17	1.87	6.55	0.16	2.34	86314				
				SI	ZE Trend							
	Min	25%	Median	75%	Max	Mean	StdDev	# Observation				
ARE	-1.83	-0.13	-0.01	0.09	2.06	-0.00	0.24	7005				
ARG	-1.87	-0.15	-0.02	0.11	2.06	-0.02	0.33	14330				
AUS	-1.58	-0.18	-0.00	0.17	1.82	0.00	0.38	312426				
AUT	-2.03	-0.12	-0.01	0.09	2.12	-0.02	0.27	23602				
BEL	-2.03	-0.11	-0.02	0.07	2.12	-0.02	0.26	37742				
BGR	-2.03	-0.15	0.00	0.16	2.12	0.02	0.37	16735				
BHR	-0.80	-0.06	0.01	0.09	2.03	0.03	0.17	3451				
BRA	-1.87	-0.15	0.00	0.15	2.06	-0.00	0.36	59280				
CAN	-1.90	-0.16	0.00	0.16	1.83	-0.00	0.37	239767				
CHE	-2.03	-0.11	-0.01	0.08	2.12	-0.02	0.24	54559				
CHL	-1.87	-0.10	-0.00	0.09	2.06	-0.00	0.22	33043				
CHN	-0.97	-0.10	-0.01	0.10	1.12	0.01	0.19	289226				
COL	-1.34	-0.09	0.00	0.10	1.93	0.01	0.22	7196				
СҮР	-2.03	-0.18	-0.00	0.16	2.12	-0.00	0.35	20496				
CZE	-2.03	-0.13	0.00	0.12	2.12	-0.01	0.26	8939				
DEU	-2.03	-0.17	-0.03	0.09	2.12	-0.06	0.35	210512				
DNK	-2.03	-0.15	-0.02	0.09	2.12	-0.03	0.27	46094				
EGY	-1.87	-0.13	-0.01	0.11	2.06	0.02	0.28	17615				
ESP	-2.03	-0.11	-0.00	0.10	2.12	0.00	0.26	40961				
EST	-2.03	-0.13	-0.01	0.13	2.12	-0.00	0.32	2663				
FIN	-2.03	-0.12	0.00	0.14	2.12	0.00	0.27	29719				
FRA	-2.03	-0.12	0.00	0.12	2.12	0.00	0.28	190957				
GBR	-2.03	-0.15	-0.00	0.13	2.12	-0.02	0.34	412607				
GRC	-2.03	-0.19	-0.03	0.13	2.12	-0.02	0.32	59060				
HKG	-1.58	-0.18	-0.02	0.14	1.82	-0.00	0.35	224713				

130 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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					Description (Continued)			
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations
HRV	-2.03	-0.14	-0.02	0.08	2.12	-0.02	0.24	15242
HUN	-1.80	-0.20	-0.05	0.08	2.12	-0.06	0.30	8352
IDN	-1.87	-0.18	-0.03	0.13	2.06	-0.01	0.35	68953
IND	-1.71	-0.22	-0.04	0.13	2.06	-0.03	0.36	531335
IRL	-2.03	-0.10	0.02	0.13	2.12	0.01	0.29	10816
ISL	-2.03	-0.11	0.00	0.12	2.12	0.01	0.30	5886
ISR	-2.03	-0.15	-0.02	0.10	2.12	-0.03	0.31	95397
ITA	-2.03	-0.12	-0.01	0.09	2.12	-0.00	0.24	61869
JOR	-1.87	-0.15	0.11	0.35	2.06	0.09	0.44	16405
JPN	-1.58	-0.12	-0.01	0.09	1.82	-0.01	0.22	817152
KAZ	-1.87	-0.11	0.00	0.13	2.06	0.03	0.41	1465
KOR	-1.58	-0.18	-0.03	0.13	1.82	-0.02	0.34	333045
KWT	-1.87	-0.12	-0.02	0.10	2.06	-0.01	0.24	24561
LKA	-1.73	-0.13	-0.01	0.10	2.06	0.00	0.24	18554
LTU	-2.03	-0.14	-0.01	0.12	2.12	-0.01	0.32	5661
LUX	-2.03	-0.08	0.00	0.11	2.12	0.03	0.28	4479
LVA	-2.03	-0.12	0.01	0.17	2.12	0.04	0.32	4654
MAR	-1.87	-0.09	-0.00	0.08	2.06	-0.00	0.20	11086
MEX	-1.87	-0.13	-0.01	0.09	2.06	-0.02	0.25	21875
MKD	-1.45	-0.10	0.00	0.07	1.26	-0.01	0.19	4151
MLT	-1.23	-0.07	0.00	0.08	2.05	0.02	0.24	1873
MYS	-1.87	-0.14	-0.03	0.09	2.06	-0.02	0.26	198058
NGA	-1.87	-0.16	-0.01	0.14	2.06	0.01	0.35	17787
NLD	-2.03	-0.11	0.00	0.11	2.12	-0.01	0.26	38059
NOR	-2.03	-0.13	-0.00	0.13	2.12	0.00	0.33	47416
NZL	-1.58	-0.09	0.01	0.11	1.82	0.02	0.25	19565
PAK	-1.87	-0.19	-0.04	0.09	2.06	-0.03	0.29	46403
PER	-1.87	-0.14	0.00	0.12	2.06	0.00	0.28	14133
PHL	-1.87	-0.15	-0.01	0.13	2.06	0.01	0.32	42721
POL	-2.03	-0.21	-0.04	0.12	2.12	-0.05	0.36	67405
PRT	-2.03	-0.14	-0.02	0.08	2.12	-0.02	0.25	16343
ROM	-2.03	-0.13	0.00	0.21	2.12	0.06	0.39	45558
RUS	-2.03	-0.13	0.00	0.12	2.12	-0.01	0.30	29503
SAU	-1.87	-0.11	-0.00	0.11	2.06	0.01	0.23	16068
SGP	-1.58	-0.14	-0.02	0.10	1.82	-0.02	0.26	124879
SVK	-2.03	-0.05	0.02	0.13	2.12	0.05	0.31	3941
SVN	-2.03	-0.16	-0.03	0.07	1.98	-0.06	0.30	10856

 Table A.8.
 (Continued)

GLOBAL CREDIT REVIEW VOLUME 3 131

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			Tab	le A.8. (C	Continued)			
				SI	ZE Trend			
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations
SWE	-2.03	-0.15	-0.01	0.14	2.12	-0.01	0.34	85358
THA	-1.87	-0.15	-0.02	0.12	2.06	-0.00	0.28	102574
TUR	-2.03	-0.17	-0.03	0.12	2.12	-0.01	0.30	64054
TWN	-1.58	-0.14	-0.02	0.11	1.82	-0.01	0.24	250096
UKR	-2.03	-0.16	0.00	0.18	2.12	0.01	0.42	8699
USA	-1.90	-0.15	-0.01	0.13	1.83	-0.02	0.33	1569893
VEN	-1.87	-0.17	-0.02	0.11	2.06	0.00	0.41	5337
VNM	-1.87	-0.20	-0.05	0.08	2.06	-0.06	0.26	40028
ZAF	-1.87	-0.16	-0.01	0.13	2.06	-0.03	0.36	86314
					M/B			
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations
ARE	0.36	0.86	1.02	1.30	8.35	1.18	0.62	6100
ARG	0.18	0.82	1.00	1.26	520.99	1.74	8.92	12709
AUS*	0.19	0.93	1.33	2.37	14.61	2.26	2.62	286945
AUT	0.20	0.94	1.06	1.37	83.80	1.35	1.93	20902
BEL	0.14	0.93	1.09	1.45	119.97	1.53	3.60	29960
BGR	0.14	0.66	0.92	1.29	50.58	1.24	1.86	10258
BHR	0.40	0.93	1.06	1.26	5.54	1.18	0.48	2754
BRA	0.18	0.83	1.07	1.60	553.67	12.50	67.96	51292
CAN	0.22	0.98	1.31	2.09	58.66	2.22	3.99	215521
CHE	0.16	0.99	1.14	1.61	119.97	1.58	1.87	51024
CHL	0.18	0.86	1.14	1.68	553.67	2.99	22.06	29461
CHN	0.66	1.47	2.08	3.13	41.43	2.69	2.42	266200
COL	0.23	0.79	1.03	1.27	553.67	1.53	13.45	5967
СҮР	0.14	0.61	0.80	1.04	46.50	1.10	1.98	16601
CZE	0.15	0.66	0.92	1.16	9.28	1.02	0.59	6492
DEU	0.14	1.00	1.21	1.66	119.97	1.81	4.06	177136
DNK	0.14	0.96	1.06	1.41	119.97	1.57	2.65	42732
EGY	0.21	0.98	1.24	1.83	92.90	1.71	2.46	14608
ESP	0.14	0.95	1.10	1.46	119.97	1.45	2.58	36987
EST	0.17	0.94	1.17	1.81	43.24	1.80	2.54	2463
FIN	0.20	1.01	1.23	1.73	119.97	1.65	2.24	28147
FRA	0.14	0.94	1.13	1.56	119.97	1.68	4.15	159530
GBR	0.14	0.97	1.33	2.07	119.97	2.25	5.15	375674
GRC	0.14	0.87	1.10	1.62	119.97	1.82	4.61	55498
HKG*	0.19	0.72	0.99	1.55	14.61	1.53	1.89	209618

*Winsorization levels are at 0.1 and 99.5 percentiles instead of 0.1 and 99.9 percentiles.

(Continued)

132 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

b1728_Ch-07.indd 132

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			140	le A.8. (C	Continued) M/B			
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations
HRV	0.14	0.71	0.93	1.19	24.60	1.10	1.21	11831
HUN	0.14	0.74	1.00	1.36	119.97	1.25	2.57	7323
IDN	0.18	0.87	1.07	1.49	553.67	1.51	6.89	60466
IND**	0.19	0.77	0.99	1.49	11.46	1.47	1.60	455646
IRL	0.14	0.99	1.22	1.71	45.56	1.71	2.07	10069
ISL	0.31	1.09	1.22	1.63	119.97	2.10	6.86	4718
ISR	0.14	0.91	1.04	1.36	119.97	1.92	7.01	72419
ITA	0.19	0.95	1.06	1.35	119.97	1.32	3.40	59035
JOR	0.19	0.79	1.00	1.31	70.95	1.25	1.97	13254
JDN JPN*	0.19	0.85	1.02	1.24	14.61	1.23	1.00	795167
KAZ	0.23	0.90	1.00	1.23	9.35	1.17	0.63	882
KOR*	0.19	0.80	0.99	1.32	14.61	1.33	1.36	282862
KWT	0.18	0.90	1.17	1.58	38.02	1.38	0.95	22002
LKA	0.24	0.95	1.15	1.62	553.67	1.70	6.59	16193
LTU	0.31	0.81	1.00	1.37	5.32	1.18	0.61	4326
LUX	0.14	0.74	0.97	1.18	9.28	1.08	0.67	3090
LVA	0.14	0.55	0.75	1.01	5.86	0.85	0.53	3140
MAR	0.18	1.07	1.28	1.84	15.95	1.65	1.00	10386
MEX	0.18	0.78	1.03	1.43	10.84	1.21	0.71	19295
MKD	0.15	0.68	0.92	1.20	119.97	3.29	15.24	2317
MLT	0.25	0.98	1.08	1.41	15.76	1.34	0.94	1330
MYS	0.18	0.77	0.99	1.41	553.67	1.44	5.88	185867
NGA	0.18	0.91	1.21	1.85	179.42	1.86	5.00	14547
NLD	0.14	1.00	1.22	1.68	119.97	1.68	2.55	36129
NOR	1.14	0.95	1.14	1.71	119.97	1.87	3.66	42915
NZL*	0.19	0.97	1.27	1.96	14.61	1.90	2.10	17458
PAK	0.18	0.84	1.00	1.32	50.46	1.32	1.71	28051
PER	0.18	0.80	1.12	1.67	29.63	1.51	1.37	11472
PHL	0.18	0.76	1.05	1.69	553.67	6.15	43.72	38165
POL	0.14	0.86	1.11	1.64	119.97	1.79	5.05	56985
PRT	0.14	0.90	1.02	1.25	27.99	1.16	0.72	14088
ROM	0.14	0.67	0.91	1.25	119.97	1.85	9.20	15944
RUS	0.14	0.84	1.11	1.64	119.97	4.56	19.01	20892
SAU	0.19	1.18	1.71	2.84	49.17	2.45	2.32	14590
SGP*	0.19	0.82	1.04	1.45	14.61	1.37	1.29	115458
SVK	0.26	0.73	0.90	1.05	3.18	0.92	0.33	1382
SVN	0.14	0.67	0.85	1.03	4.99	0.92	0.47	6981

 Table A.8.
 (Continued)

**Winsorization levels are at 1 and 99 percentiles instead of 0.1 and 99.9 percentiles.

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GLOBAL CREDIT REVIEW VOLUME 3 133

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			Tab	ole A.8. (C	Continued)			
					M/B			
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations
SWE	0.14	1.03	1.37	2.16	119.97	2.19	4.14	77958
THA	0.18	0.85	1.06	1.45	109.43	1.31	1.10	92653
TUR	0.14	076	0.99	1.32	119.97	1.30	3.23	59875
TWN*	0.29	0.93	1.16	1.65	14.61	1.48	1.06	227830
UKR	0.14	0.88	1.22	1.94	66.21	1.83	2.68	5506
USA	0.22	1.02	1.30	2.09	58.66	2.10	3.14	1506686
VEN	0.18	0.58	0.87	1.06	553.67	25.05	85.55	3817
VNM	0.18	0.82	0.95	1.19	18.99	1.15	0.79	34541
ZAF	0.18	0.89	1.19	1.84	553.67	2.41	14.85	79989
				S	SIGMA			
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations
ARE	0.01	0.09	0.11	0.16	0.43	0.13	0.06	6132
ARG	0.03	0.09	0.11	0.14	0.64	0.12	0.06	13011
AUS	0.02	0.13	0.22	0.32	0.93	0.24	0.15	295702
AUT	0.01	0.06	0.09	0.13	1.14	0.12	0.10	21775
BEL	0.02	0.07	0.09	0.12	1.41	0.11	0.08	33381
BGR	0.02	0.14	0.18	0.28	1.09	0.22	0.13	11218
BHR	0.03	0.07	0.09	0.12	0.24	0.10	0.03	2761
BRA	0.01	0.10	0.14	0.20	1.19	0.18	0.13	52177
CAN	0.03	0.11	0.17	0.27	1.01	0.21	0.15	227954
CHE	0.01	0.07	0.09	0.13	1.41	0.11	0.07	51919
CHL	0.01	0.06	0.08	0.10	0.71	0.09	0.06	29731
CHN	0.03	0.08	0.10	0.13	0.43	0.11	0.05	281900
COL	0.01	0.06	0.09	0.12	0.47	0.10	0.06	6078
СҮР	0.02	0.15	0.21	0.28	1.41	0.25	0.17	17602
CZE	0.03	0.09	0.13	0.17	0.42	0.13	0.05	7145
DEU	0.01	0.10	0.14	0.24	1.41	0.21	0.22	194443
DNK	0.01	0.07	0.10	0.15	1.24	0.13	0.10	43321
EGY	0.01	0.09	0.13	0.19	1.19	0.15	0.09	15577
ESP	0.01	0.06	0.09	0.12	0.93	0.10	0.06	37885
EST	0.01	0.08	0.13	0.20	0.65	0.16	0.10	2522
FIN	0.03	0.08	0.11	0.15	1.41	0.13	0.09	28691
FRA	0.01	0.08	0.11	0.16	1.41	0.13	0.09	172286
GBR	0.01	0.08	0.12	0.18	1.28	0.15	0.10	385233
GRC	0.01	0.10	0.14	0.18	0.90	0.15	0.08	57250
HKG	0.02	0.11	0.16	0.24	0.93	0.19	0.11	217971

*Winsorization levels are at 0.1 and 99.5 percentiles instead of 0.1 and 99.9 percentiles.

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134 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

b1728_Ch-07.indd 134

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			Tab	ole A.8. (C	Continued)			
				5	SIGMA			
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations
HRV	0.01	0.11	0.14	0.18	0.75	0.16	0.08	12104
HUN	0.02	0.09	0.13	0.18	0.66	0.15	0.09	7647
IDN	0.01	0.13	0.18	0.27	1.19	0.22	0.13	62258
IND	0.05	0.15	0.19	0.24	0.99	0.22	0.13	480192
IRL	0.03	0.08	0.11	0.17	1.41	0.15	0.14	10151
ISL	0.03	0.08	0.11	0.14	0.60	0.12	0.07	4985
ISR	0.01	0.10	0.13	0.20	1.13	0.16	0.09	80390
ITA	0.01	0.07	0.09	0.13	0.75	0.11	0.05	60334
JOR	0.01	0.10	0.12	0.14	0.80	0.13	0.05	14836
JPN	0.02	0.08	0.11	0.16	0.93	0.13	0.07	802610
KAZ	0.01	0.11	0.15	0.23	1.01	0.19	0.13	877
KOR	0.02	0.11	0.15	0.21	0.79	0.17	0.08	324689
KWT	0.01	0.10	0.12	0.16	0.76	0.13	0.05	22313
LKA	0.03	0.11	0.15	0.20	0.98	0.17	0.09	17127
LTU	0.03	0.09	0.13	0.18	1.00	0.15	0.10	5179
LUX	0.02	0.07	0.09	0.12	0.51	0.10	0.05	3189
LVA	0.03	0.11	0.15	0.22	0.95	0.17	0.09	3222
MAR	0.02	0.08	0.09	0.12	0.53	0.10	0.05	10524
MEX	0.01	0.07	0.10	0.13	1.17	0.11	0.07	19681
MKD	0.01	0.08	0.12	0.17	0.53	0.14	0.07	2401
MLT	0.02	0.05	0.07	0.09	0.54	0.08	0.05	1401
MYS	0.02	0.10	0.14	0.21	1.19	0.17	0.11	193736
NGA	0.01	0.11	0.14	0.17	0.54	0.14	0.06	15500
NLD	0.02	0.07	0.09	0.13	1.24	0.11	0.08	37092
NOR	0.03	0.10	0.14	0.20	1.08	0.17	0.10	43879
NZL	0.02	0.07	0.09	0.13	0.93	0.12	0.11	18063
PAK	0.03	0.11	0.15	0.24	1.19	0.22	0.20	37075
PER	0.02	0.09	0.12	0.16	0.61	0.13	0.07	11549
PHL	0.01	0.12	0.17	0.25	0.89	0.20	0.12	39128
POL	0.01	0.12	0.16	0.23	1.41	0.19	0.12	63844
PRT	0.01	0.07	0.10	0.14	1.16	0.12	0.08	14347
ROM	0.01	0.17	0.22	0.33	1.41	0.26	0.15	20711
RUS	0.01	0.10	0.12	0.20	1.22	0.16	0.11	20998
SAU	0.02	0.07	0.10	0.14	0.63	0.11	0.06	15520
SGP	0.02	0.10	0.15	0.23	0.93	0.19	0.14	120940
SVK	0.01	0.08	0.11	0.16	0.58	0.13	0.08	1368
SVN	0.01	0.07	0.10	0.14	1.12	0.13	0.11	8391

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GLOBAL CREDIT REVIEW VOLUME 3 135

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			Tal	ole A.8. (C	Continued)								
		SIGMA											
	Min	25%	Median	75%	Max	Mean	StdDev	# Observations					
SWE	0.01	0.10	0.14	0.25	1.41	0.20	0.16	81508					
THA	0.01	0.09	0.13	0.18	1.19	0.15	0.10	99419					
TUR	0.01	0.10	0.14	0.19	1.25	0.15	0.07	62680					
TWN	0.02	0.09	0.11	0.14	0.86	0.12	0.05	242951					
UKR	0.01	0.15	0.19	0.29	1.03	0.24	0.15	5468					
USA	0.03	0.09	0.15	0.23	1.01	0.18	0.12	1532110					
VEN	0.01	0.13	0.18	0.23	0.68	0.19	0.10	4065					
VNM	0.01	0.11	0.14	0.19	0.67	0.16	0.07	36605					
ZAF	0.01	0.09	0.13	0.22	1.19	0.19	0.18	81441					

Table A.9. Exits classified as "Defaults".

	Default						
Action Type	Subcategory						
Bankruptcy filling	Administration, Arrangement, Canadian CCAA, Chapter 7, Chapter 11, Chapter 15, Conservatorship,						
	Insolvency, Japanese CRL, Judicial Management, Liquidation, Pre-Negotiation Chapter 11, Protection,						
	Receivership, Rehabilitation, Rehabilitation (Thailand 1997), Reorganization, Restructuring, Section 304,						
	Supreme Court declaration, Winding up, Work out, Other, Unknown						
Delisting	Bankruptcy						
Default Corporate Action	Bankruptcy, Coupon & Principal Payment, Coupon Payment Only, Debt Restructuring, Interest Payment, Loan Payment, Principal Payment, ADR (Japan only), Declared Sick (India Only), Unknown						

Table A.10. Exits classified as "Other Exits".

	Other Exits								
Action Type	Subcategory								
Delisting	Unknown, Acquired/Merged, Assimilated with underlying shares, Bid price below minimum, Cancellation of listing, End of When-issued trading, Expired, Failure to meet listing requirements, Failure to pay listing fees, Inactive security, Insufficient assets, Insufficient capital and surplus, Insufficient number of market makers, Issue postponed, Lack of market maker interest, Lack of public interest, Liquidated, Matured, Not available, Not current in required filings, NP/FP finished, Privatized, Reorganization security called for redemptions, the company's request, Scheme of arrangement, Insufficient spread of holders, Selective capital reduction of the company								

 $136 \quad \text{nus-rmi credit research initiative technical report}$

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		ECOIIC	my:ARE				E	conol	my:ARG
		D	efaults	0	others			D	efaults
Year	Active	#	%	#	%	Year	Active	#	%
1992	0	0	NaN	0	NaN	1992	1	0	0.00
1993	0	0	NaN	0	NaN	1993	1	0	0.00
1994	0	0	NaN	0	NaN	1994	23	0	0.00
1995	0	0	NaN	0	NaN	1995	87	0	0.00
1996	0	0	NaN	0	NaN	1996	95	0	0.00
1997	0	0	NaN	0	NaN	1997	78	0	0.00
1998	0	0	NaN	0	NaN	1998	66	2	1.87
1999	0	0	NaN	0	NaN	1999	71	1	1.06
2000	0	0	NaN	0	NaN	2000	68	1	1.10
2001	0	0	NaN	0	NaN	2001	51	2	2.30
2002	0	0	NaN	0	NaN	2002	65	9	10.23
2003	0	0	NaN	0	NaN	2003	69	2	2.41
2004	0	0	NaN	0	NaN	2004	64	0	0.00
2005	0	0	NaN	0	NaN	2005	69	0	0.00
2006	70	0	0.00	10	12.50	2006	71	0	0.00
2007	85	0	0.00	10	10.53	2007	75	0	0.00
2008	78	0	0.00	20	20.41	2008	65	0	0.00
2009	81	0	0.00	19	19.00	2009	65	1	1.35
2010	80	0	0.00	27	25.23	2010	69	1	1.35
2011	83	0	0.00	21	20.19	2011	65	0	0.00
2012	82	1	0.96	21	20.19	2012	66	0	0.00

 Table A.11.
 Number of defaults and other exits of 71 economies from 1992 to 2012.

		Econo	my:AUS				I	Econor	ny:AUT		
		De	faults	0	thers			D	efaults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	696	0	0.00	117	14.39	1992	87	0	0.00	3	3.33
1993	828	0	0.00	48	5.48	1993	101	0	0.00	9	8.18
1994	915	0	0.00	93	9.23	1994	111	0	0.00	1	0.89
1995	953	1	0.10	82	7.92	1995	118	0	0.00	2	1.67
1996	1002	1	0.09	65	6.09	1996	117	1	0.81	5	4.07
1997	1012	3	0.27	102	9.13	1997	118	0	0.00	6	4.84
1998	1009	1	0.09	109	9.74	1998	112	0	0.00	15	11.81
1999	1073	2	0.17	99	8.43	1999	108	0	0.00	17	13.60
2000	1187	11	0.84	105	8.06	2000	119	0	0.00	15	11.19
2001	1164	27	2.08	109	8.38	2001	113	2	1.39	29	20.14
2002	1180	9	0.70	101	7.83	2002	109	0	0.00	14	11.38
2003	1226	8	0.60	94	7.08	2003	108	0	0.00	21	16.28
2004	1348	3	0.21	76	5.33	2004	100	0	0.00	23	18.70
2005	1461	6	0.39	87	5.60	2005	99	0	0.00	18	15.38
2006	1578	5	0.29	117	6.88	2006	104	0	0.00	10	8.77
2007	1752	5	0.27	114	6.09	2007	106	0	0.00	11	9.40
2008	1685	27	1.45	149	8.01	2008	105	2	1.71	10	8.55
2009	1670	30	1.65	115	6.34	2009	100	1	0.88	13	11.40
2010	1690	4	0.22	137	7.48	2010	94	1	0.87	20	17.39
2011	1682	0	0.00	178	9.57	2011	84	0	0.00	16	16.00
2012	1663	3	0.16	164	8.96	2012	81	1	1.08	11	11.83

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Others

%

0.00

0.00

8.00

17.14

18.80

28.44

36.45

23.40

24.18

39.08

15.91

14.46

16.88

5.48

10.13

11.76

20.73

10.81

5.41

12.16

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GLOBAL CREDIT REVIEW VOLUME 3 137

	E	conor	my:BEL				Eco	nomy
		De	faults	0	thers			Def
Year	Active	#	%	#	%	Year	Active	#
1992	134	0	0.00	8	5.63	1992	0	0
1993	139	0	0.00	6	4.14	1993	0	0
1994	146	0	0.00	12	7.59	1994	0	0
1995	151	0	0.00	10	6.21	1995	0	0
1996	164	0	0.00	11	6.29	1996	0	0
1997	163	0	0.00	18	9.94	1997	0	0
1998	175	0	0.00	17	8.85	1998	0	0
1999	191	2	1.01	6	3.02	1999	0	0
2000	192	0	0.00	10	4.95	2000	14	0
2001	186	2	1.00	12	6.00	2001	23	0
2002	176	3	1.55	15	7.73	2002	28	0
2003	177	1	0.52	14	7.29	2003	31	0
2004	170	1	0.54	13	7.07	2004	36	0
2005	172	2	1.08	12	6.45	2005	127	1
2006	183	2	1.03	9	4.64	2006	236	0
2007	217	1	0.36	61	21.86	2007	241	2
2008	201	1	0.36	79	28.11	2008	211	0
2009	201	2	0.78	52	20.39	2009	198	0
2010	203	0	0.00	54	21.01	2010	175	1
2011	175	1	0.42	63	26.36	2011	168	0
2012	185	1	0.44	39	17.33	2012	158	0

Table A.11.(Continued)

	Eco	nomy	BGR:			
		De	faults	Ot	hers	
Year	Active	#	%	#	%	
1992	0	0	NaN	0	NaN	
1993	0	0	NaN	0	NaN	
1994	0	0	NaN	0	NaN	
1995	0	0	NaN	0	NaN	
1996	0	0	NaN	0	NaN	
1997	0	0	NaN	0	NaN	
1998	0	0	NaN	0	NaN	
1999	0	0	NaN	0	NaN	
2000	14	0	0.00	10	41.67	
2001	23	0	0.00	8	25.81	
2002	28	0	0.00	7	20.00	
2003	31	0	0.00	9	22.50	
2004	36	0	0.00	3	7.69	
2005	127	1	0.65	25	16.34	
2006	236	0	0.00	36	13.24	
2007	241	2	0.64	71	22.61	
2008	211	0	0.00	101	32.37	
2009	198	0	0.00	88	30.77	
2010	175	1	0.37	91	34.08	
2011	168	0	0.00	78	31.71	
2012	158	0	0.00	69	30.40	

Economy:BRA

		De	faults	Ot	hers			De	efaults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	9
1992	0	0	NaN	0	NaN	1992	0	0	NaN	0	N
1993	0	0	NaN	0	NaN	1993	0	0	NaN	0	Ν
1994	0	0	NaN	0	NaN	1994	261	0	0.00	26	9
1995	0	0	NaN	0	NaN	1995	273	0	0.00	92	25
1996	0	0	NaN	0	NaN	1996	282	0	0.00	99	25
1997	0	0	NaN	0	NaN	1997	259	0	0.00	133	33
1998	0	0	NaN	0	NaN	1998	284	2	0.45	158	35
1999	0	0	NaN	0	NaN	1999	323	2	0.47	104	24
2000	0	0	NaN	0	NaN	2000	294	1	0.24	122	29
2001	0	0	NaN	0	NaN	2001	274	1	0.24	145	34
2002	0	0	NaN	0	NaN	2002	241	2	0.54	126	34
2003	0	0	NaN	0	NaN	2003	264	2	0.56	90	25
2004	28	0	0.00	2	6.67	2004	271	0	0.00	83	23
2005	37	0	0.00	1	2.63	2005	266	1	0.30	71	21
2006	30	0	0.00	10	25.00	2006	282	0	0.00	55	16
2007	36	0	0.00	5	12.20	2007	350	0	0.00	40	10
2008	31	0	0.00	10	24.39	2008	331	0	0.00	56	14
2009	32	0	0.00	13	28.89	2009	334	0	0.00	35	9
2010	31	0	0.00	14	31.11	2010	326	0	0.00	42	11
2011	27	0	0.00	19	41.30	2011	317	1	0.28	40	11
2012	29	0	0.00	19	39.58	2012	297	4	1.17	41	11

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 $138 \quad {\rm nus-rmi\ credit\ research\ initiative\ technical\ report}$

Economy:BHR

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	E	conon	ny:CAN	Ĩ	
		Def	faults	0	thers
Year	Active	#	%	#	%
1992	946	1	0.10	103	9.81
1993	1156	0	0.00	72	5.86
1994	1328	0	0.00	54	3.91
1995	1449	0	0.00	83	5.42
1996	1635	0	0.00	80	4.66
1997	1781	5	0.26	144	7.46
1998	1752	8	0.40	254	12.61
1999	1194	10	0.52	716	37.29
2000	1105	8	0.61	196	14.97
2001	949	20	1.68	223	18.71
2002	936	5	0.49	74	7.29
2003	934	14	1.36	81	7.87
2004	993	5	0.47	74	6.90
2005	1044	3	0.26	86	7.59
2006	1096	3	0.25	100	8.34
2007	1131	3	0.24	125	9.93
2008	1110	13	1.05	115	9.29
2009	1039	14	1.17	143	11.96
2010	1073	5	0.43	97	8.26
2011	1079	6	0.50	122	10.11
2012	1053	9	0.76	127	10.68

 Table A.11. (Continued)

	Ec	conom	y:CHE		
		De	faults	C	Others
Year	Active	#	%	#	%
1992	143	0	0.00	31	17.82
1993	174	0	0.00	10	5.43
1994	177	0	0.00	20	10.15
1995	189	0	0.00	15	7.35
1996	210	0	0.00	15	6.67
1997	221	1	0.43	11	4.72
1998	227	0	0.00	14	5.81
1999	243	0	0.00	15	5.81
2000	262	0	0.00	12	4.38
2001	258	2	0.72	16	5.80
2002	250	0	0.00	17	6.37
2003	244	2	0.77	13	5.02
2004	238	1	0.40	12	4.78
2005	247	1	0.39	7	2.75
2006	251	0	0.00	17	6.34
2007	255	0	0.00	9	3.41
2008	253	0	0.00	16	5.95
2009	256	0	0.00	16	5.88
2010	251	0	0.00	18	6.69
2011	247	1	0.38	18	6.77
2012	376	1	0.22	87	18.75

Economy:CHN

	Ec	conom	y:CHL				E	conom	y:CHN		
		De	faults	0	thers			Def	aults	0	Others
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN	1992	45	0	0.00	2	4.20
1993	0	0	NaN	0	NaN	1993	165	0	0.00	0	0.00
1994	141	0	0.00	9	6.00	1994	283	1	0.35	1	0.35
1995	166	0	0.00	26	13.54	1995	317	6	1.85	1	0.31
1996	168	0	0.00	46	21.50	1996	522	10	1.88	0	0.00
1997	181	0	0.00	35	16.20	1997	727	15	2.02	2	0.27
1998	168	0	0.00	56	25.00	1998	841	34	3.88	2	0.23
1999	177	0	0.00	41	18.81	1999	935	23	2.40	1	0.10
2000	168	0	0.00	43	20.38	2000	1079	26	2.35	1	0.09
2001	168	1	0.47	43	20.28	2001	1151	51	4.21	8	0.66
2002	155	1	0.48	54	25.71	2002	1203	49	3.85	21	1.65
2003	152	0	0.00	57	27.27	2003	1266	46	3.45	23	1.72
2004	164	0	0.00	32	16.33	2004	1353	110	7.39	26	1.75
2005	166	0	0.00	39	19.02	2005	1351	100	6.79	21	1.43
2006	171	0	0.00	41	19.34	2006	1372	73	4.82	70	4.62
2007	187	0	0.00	27	12.62	2007	1457	65	4.02	93	5.76
2008	146	0	0.00	52	26.26	2008	1581	45	2.68	55	3.27
2009	167	0	0.00	28	14.36	2009	1681	54	3.06	31	1.76
2010	161	0	0.00	41	20.30	2010	1992	42	2.02	46	2.21
2011	170	0	0.00	36	17.48	2011	2272	14	0.60	61	2.60
2012	167	0	0.00	53	24.09	2012	2425	18	0.72	57	2.28

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GLOBAL CREDIT REVIEW VOLUME 3 139

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	E	conor	ny:COL				E	conoi	ny:CYP		
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN	1992	0	0	NaN	0	NaN
1993	0	0	NaN	0	NaN	1993	0	0	NaN	0	NaN
1994	1	0	0.00	0	0.00	1994	0	0	NaN	0	NaN
1995	49	0	0.00	29	37.18	1995	0	0	NaN	0	NaN
1996	43	0	0.00	46	51.69	1996	35	0	0.00	3	7.89
1997	45	0	0.00	43	48.86	1997	43	0	0.00	0	0.00
1998	58	0	0.00	64	52.46	1998	48	0	0.00	2	4.00
1999	47	0	0.00	55	53.92	1999	58	0	0.00	2	3.33
2000	39	0	0.00	40	50.63	2000	116	0	0.00	4	3.33
2001	49	0	0.00	19	27.94	2001	137	0	0.00	6	4.20
2002	50	0	0.00	21	29.58	2002	144	0	0.00	10	6.49
2003	53	0	0.00	15	22.06	2003	134	0	0.00	21	13.55
2004	53	0	0.00	12	18.46	2004	134	0	0.00	29	17.79
2005	55	0	0.00	17	23.61	2005	139	0	0.00	22	13.66
2006	47	0	0.00	18	27.69	2006	137	0	0.00	11	7.43
2007	50	0	0.00	16	24.24	2007	137	0	0.00	8	5.52
2008	35	0	0.00	24	40.68	2008	121	0	0.00	31	20.39
2009	42	0	0.00	10	19.23	2009	109	0	0.00	26	19.26
2010	44	0	0.00	14	24.14	2010	107	0	0.00	27	20.15
2011	40	0	0.00	11	21.57	2011	83	0	0.00	50	37.59
2012	40	1	1.96	10	19.61	2012	75	0	0.00	55	42.31

Table A.11. (Continued)

	Ε	conor	ny:CZE				E	Conor	ny:DEU	ſ	
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN	1992	399	0	0.00	38	8.70
1993	0	0	NaN	0	NaN	1993	420	0	0.00	28	6.25
1994	1	0	0.00	0	0.00	1994	574	0	0.00	63	9.89
1995	50	0	0.00	3	5.66	1995	592	0	0.00	65	9.89
1996	148	0	0.00	6	3.90	1996	623	4	0.58	63	9.13
1997	263	0	0.00	369	58.39	1997	630	2	0.28	74	10.48
1998	235	1	0.36	42	15.11	1998	727	2	0.26	53	6.78
1999	142	4	1.60	104	41.60	1999	905	1	0.10	51	5.33
2000	101	6	3.45	67	38.51	2000	1036	2	0.18	56	5.12
2001	75	2	1.33	73	48.67	2001	1033	26	2.33	55	4.94
2002	46	1	0.98	55	53.92	2002	952	37	3.40	100	9.18
2003	33	0	0.00	38	53.52	2003	893	16	1.62	76	7.72
2004	42	0	0.00	25	37.31	2004	886	8	0.85	47	4.99
2005	25	0	0.00	25	50.00	2005	915	4	0.42	41	4.27
2006	19	0	0.00	18	48.65	2006	1076	4	0.36	35	3.14
2007	12	0	0.00	13	52.00	2007	1222	5	0.39	60	4.66
2008	15	0	0.00	7	31.82	2008	1277	17	1.20	120	8.49
2009	13	0	0.00	8	38.10	2009	1247	11	0.77	170	11.90
2010	20	0	0.00	2	9.09	2010	1298	0	0.00	153	10.54
2011	16	1	4.35	6	26.09	2011	1315	5	0.30	336	20.29
2012	15	0	0.00	4	21.05	2012	954	9	0.62	485	33.49

(Continued)

140 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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Economy:DNK										
		De	faults	0	thers					
Year	Active	#	%	#	%					
1992	157	0	0.00	19	10.80					
1993	173	0	0.00	12	6.49					
1994	176	0	0.00	24	12.00					
1995	202	1	0.46	16	7.31					
1996	219	0	0.00	11	4.78					
1997	212	0	0.00	21	9.01					
1998	213	0	0.00	29	11.98					
1999	210	0	0.00	24	10.26					
2000	208	1	0.44	20	8.73					
2001	191	5	2.20	31	13.66					
2002	175	3	1.44	30	14.42					
2003	174	1	0.52	19	9.79					
2004	170	1	0.54	14	7.57					
2005	169	0	0.00	9	5.06					
2006	190	0	0.00	6	3.06					
2007	216	1	0.45	4	1.81					
2008	213	1	0.43	16	6.96					
2009	208	3	1.36	9	4.09					
2010	198	0	0.00	15	7.04					
2011	187	2	1.00	12	5.97					
2012	176	2	1.06	11	5.82					

 Table A.11.
 (Continued)

Economy:EGY											
	De	faults	0	thers							
Active	#	%	#	%							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
0	0	NaN	0	NaN							
171	0	0.00	78	31.33							
200	0	0.00	124	38.27							
254	0	0.00	141	35.70							
208	0	0.00	90	30.20							
198	0	0.00	59	22.96							
235	0	0.00	11	4.47							
214	0	0.00	51	19.25							
	Active 0 0 0 0 0 0 0 0 0 0 0 0 0	Active # 0 0 171 0 200 0 254 0 208 0 198 0 235 0	Defaults Active # % 0 0 NaN 0 0 0.00 200 0 0.00 204 0 0.00 205 <	Defaults O Active # $%$ # 0 0 NaN 0 0							

	E	cono	my:ESP				E	cono	my:EST		
		De	faults	0	thers			De	faults	(Dt
Year	Active	#	%	#	%	Year	Active	#	%	#	
1992	148	0	0.00	39	20.86	1992	0	0	NaN	0	
1993	113	0	0.00	96	45.93	1993	0	0	NaN	0	
1994	239	0	0.00	18	7.00	1994	0	0	NaN	0	
1995	237	0	0.00	97	29.04	1995	0	0	NaN	0	
1996	265	0	0.00	64	19.45	1996	0	0	NaN	0	
1997	273	0	0.00	59	17.77	1997	17	0	0.00	0	
1998	232	0	0.00	104	30.95	1998	19	0	0.00	1	
1999	214	0	0.00	78	26.71	1999	19	0	0.00	1	
2000	213	0	0.00	55	20.52	2000	16	0	0.00	4	
2001	195	0	0.00	79	28.83	2001	14	0	0.00	2	
2002	207	2	0.74	60	22.30	2002	11	0	0.00	3	
2003	172	0	0.00	81	32.02	2003	11	0	0.00	0	
2004	162	0	0.00	41	20.20	2004	11	0	0.00	0	
2005	160	0	0.00	47	22.71	2005	13	0	0.00	1	
2006	163	0	0.00	44	21.26	2006	13	0	0.00	2	
2007	154	1	0.51	41	20.92	2007	16	0	0.00	0	
2008	144	2	1.16	27	15.61	2008	17	0	0.00	0	
2009	141	0	0.00	24	14.55	2009	15	0	0.00	2	
2010	144	1	0.63	15	9.38	2010	15	0	0.00	1	
2011	142	0	0.00	15	9.55	2011	15	0	0.00	0	
2012	137	2	1.27	18	11.46	2012	16	0	0.00	0	

GLOBAL CREDIT REVIEW VOLUME 3 141

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	E	cono	my:FIN				E	Conor	ny:FRA	L L	
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	92	0	0.00	0	0.00	1992	626	0	0.00	72	10.32
1993	94	0	0.00	2	2.08	1993	640	0	0.00	78	10.86
1994	96	0	0.00	6	5.88	1994	691	0	0.00	102	12.86
1995	102	0	0.00	5	4.67	1995	718	0	0.00	119	14.22
1996	110	0	0.00	3	2.65	1996	780	0	0.00	116	12.95
1997	124	0	0.00	1	0.80	1997	802	1	0.11	142	15.03
1998	127	1	0.75	6	4.48	1998	812	0	0.00	193	19.20
1999	146	0	0.00	8	5.19	1999	855	0	0.00	93	9.81
2000	153	0	0.00	12	7.27	2000	917	2	0.20	97	9.55
2001	148	1	0.62	13	8.02	2001	915	8	0.79	93	9.15
2002	142	1	0.65	10	6.54	2002	870	5	0.51	113	11.44
2003	138	1	0.68	9	6.08	2003	860	4	0.41	102	10.56
2004	131	0	0.00	11	7.75	2004	841	3	0.32	106	11.16
2005	132	0	0.00	6	4.35	2005	846	5	0.53	98	10.33
2006	132	0	0.00	8	5.71	2006	913	6	0.60	75	7.55
2007	130	0	0.00	5	3.70	2007	951	7	0.66	101	9.54
2008	127	1	0.76	3	2.29	2008	904	12	1.13	149	13.99
2009	125	1	0.78	2	1.56	2009	891	6	0.58	142	13.67
2010	123	0	0.00	4	3.15	2010	849	2	0.20	169	16.57
2011	121	1	0.81	1	0.81	2011	801	2	0.21	153	16.00
2012	119	0	0.00	5	4.03	2012	776	0	0.00	154	16.56

Table A.11. (Continued)

	Ε	conon	ny:GBR	2			Ε	conor	ny:GRC	l ,	
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	1079	0	0.00	89	7.62	1992	90	0	0.00	0	0.00
1993	1192	0	0.00	39	3.17	1993	97	0	0.00	0	0.00
1994	1284	0	0.00	49	3.68	1994	162	0	0.00	2	1.22
1995	1415	0	0.00	63	4.26	1995	182	0	0.00	2	1.09
1996	1623	0	0.00	62	3.68	1996	196	0	0.00	6	2.97
1997	1721	0	0.00	114	6.21	1997	208	0	0.00	4	1.89
1998	1688	0	0.00	198	10.50	1998	231	0	0.00	5	2.12
1999	1560	2	0.11	295	15.89	1999	264	0	0.00	6	2.22
2000	1675	2	0.11	210	11.13	2000	308	0	0.00	8	2.53
2001	1686	12	0.65	148	8.02	2001	312	0	0.00	14	4.29
2002	1641	13	0.71	169	9.27	2002	309	0	0.00	19	5.79
2003	1600	6	0.33	187	10.43	2003	313	0	0.00	9	2.80
2004	1786	2	0.10	151	7.79	2004	313	0	0.00	10	3.10
2005	2036	2	0.09	205	9.14	2005	299	0	0.00	21	6.56
2006	2186	0	0.00	251	10.30	2006	285	0	0.00	16	5.32
2007	2225	2	0.08	264	10.60	2007	279	0	0.00	13	4.45
2008	2029	30	1.24	355	14.71	2008	274	0	0.00	18	6.16
2009	1827	32	1.47	318	14.61	2009	265	0	0.00	24	8.30
2010	1788	3	0.15	253	12.38	2010	257	0	0.00	30	10.45
2011	1666	11	0.56	278	14.22	2011	232	0	0.00	35	13.11
2012	1582	19	1.05	211	11.64	2012	223	0	0.00	42	15.85

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142 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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	E	conor	ny:HKG	T			
		De	faults	0	thers		
Year	Active	#	%	#	%	Year	A
1992	356	0	0.00	11	3.00	1992	
1993	423	0	0.00	7	1.63	1993	
1994	464	0	0.00	13	2.73	1994	
1995	491	0	0.00	9	1.80	1995	
1996	529	0	0.00	15	2.76	1996	
1997	601	0	0.00	24	3.84	1997	
1998	628	2	0.30	30	4.55	1998	
1999	661	6	0.87	21	3.05	1999	
2000	743	4	0.52	21	2.73	2000	
2001	813	9	1.06	30	3.52	2001	
2002	908	4	0.42	36	3.80	2002	
2003	957	4	0.40	51	5.04	2003	
2004	995	0	0.00	56	5.33	2004	
2005	1028	3	0.27	65	5.93	2005	
2006	1080	2	0.18	44	3.91	2006	
2007	1170	2	0.17	24	2.01	2007	
2008	1174	8	0.66	33	2.72	2008	
2009	1233	3	0.24	23	1.83	2009	
2010	1309	1	0.07	28	2.09	2010	
2011	1366	1	0.07	26	1.87	2011	
2012	1409	2	0.14	43	2.96	2012	

Table A.11. (Continued)

	Economy:HRV											
		De	faults	0	thers							
Year	Active	#	%	#	%							
1992	0	0	NaN	0	NaN							
1993	0	0	NaN	0	NaN							
1994	0	0	NaN	0	NaN							
1995	0	0	NaN	0	NaN							
1996	0	0	NaN	0	NaN							
1997	0	0	NaN	0	NaN							
1998	0	0	NaN	0	NaN							
1999	0	0	NaN	0	NaN							
2000	0	0	NaN	0	NaN							
2001	0	0	NaN	0	NaN							
2002	28	0	0.00	4	12.50							
2003	34	0	0.00	8	19.05							
2004	48	0	0.00	9	15.79							
2005	54	0	0.00	8	12.90							
2006	211	0	0.00	20	8.66							
2007	240	0	0.00	62	20.53							
2008	164	0	0.00	120	42.25							
2009	160	0	0.00	72	31.03							
2010	162	1	0.47	50	23.47							
2011	133	0	0.00	65	32.83							
2012	127	1	0.56	50	28.09							

	Ε	conor	ny:HUN				ŀ	Econor	my:IDN		
		De	faults	0	thers			De	faults	C	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN	1992	124	0	0.00	30	19.48
1993	0	0	NaN	0	NaN	1993	155	0	0.00	23	12.92
1994	0	0	NaN	0	NaN	1994	184	0	0.00	49	21.03
1995	34	0	0.00	2	5.56	1995	204	0	0.00	41	16.73
1996	38	0	0.00	7	15.56	1996	236	0	0.00	26	9.92
1997	39	0	0.00	10	20.41	1997	255	0	0.00	33	11.46
1998	46	0	0.00	4	8.00	1998	250	17	5.48	43	13.87
1999	53	0	0.00	7	11.67	1999	257	15	5.07	24	8.1
2000	50	1	1.72	7	12.07	2000	262	6	1.99	34	11.20
2001	47	0	0.00	8	14.55	2001	272	8	2.43	49	14.89
2002	39	0	0.00	11	22.00	2002	270	2	0.60	63	18.8
2003	40	0	0.00	6	13.04	2003	297	1	0.31	28	8.5
2004	38	0	0.00	7	15.56	2004	286	1	0.28	64	18.2
2005	36	0	0.00	6	14.29	2005	277	0	0.00	77	21.7
2006	37	0	0.00	5	11.90	2006	303	0	0.00	53	14.8
2007	35	0	0.00	4	10.26	2007	321	0	0.00	71	18.1
2008	37	0	0.00	1	2.63	2008	281	0	0.00	93	24.8
2009	40	0	0.00	0	0.00	2009	323	4	1.04	56	14.6
2010	45	0	0.00	1	2.17	2010	360	2	0.50	40	9.9
2011	44	0	0.00	6	12.00	2011	386	0	0.00	42	9.8
2012	46	1	1.96	4	7.84	2012	414	1	0.22	37	8.1

GLOBAL CREDIT REVIEW VOLUME 3 143

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]	Econo	my:INI)			E	Cono	my:IRL		
		De	faults	Ot	hers			De	faults	0	Others
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	1495	1	0.06	143	8.72	1992	31	0	0.00	4	11.4
1993	1842	0	0.00	201	9.84	1993	37	0	0.00	4	9.7
1994	2801	0	0.00	270	8.79	1994	37	0	0.00	5	11.9
1995	4073	2	0.05	349	7.89	1995	37	0	0.00	1	2.6
1996	4078	5	0.10	1060	20.61	1996	43	0	0.00	0	0.0
1997	3011	11	0.22	1958	39.32	1997	51	0	0.00	3	5.5
1998	2685	8	0.19	1469	35.30	1998	50	0	0.00	5	9.0
1999	3060	13	0.31	1112	26.57	1999	51	0	0.00	6	10.5
2000	2558	10	0.26	1334	34.19	2000	59	0	0.00	5	7.8
2001	2266	6	0.18	1122	33.06	2001	54	0	0.00	6	10.0
2002	3348	6	0.15	656	16.36	2002	48	0	0.00	6	11.1
2003	2683	14	0.34	1383	33.90	2003	43	0	0.00	5	10.4
2004	2514	7	0.20	931	26.97	2004	42	0	0.00	3	6.6
2005	2523	7	0.22	611	19.45	2005	42	0	0.00	2	4.5
2006	2730	11	0.34	483	14.98	2006	47	0	0.00	2	4.0
2007	2995	16	0.51	130	4.14	2007	52	0	0.00	2	3.7
2008	3041	23	0.64	542	15.03	2008	49	0	0.00	3	5.7
2009	3146	37	1.07	283	8.17	2009	43	1	2.04	5	10.2
2010	3605	9	0.20	955	20.90	2010	40	0	0.00	4	9.0
2011	3404	9	0.21	866	20.24	2011	38	0	0.00	2	5.0
2012	3582	24	0.59	429	10.63	2012	33	0	0.00	5	13.1

Table A.11. (Continued)

]	Econ	omy:ISL				I	Econo	my:ISR		
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN	1992	0	0	NaN	0	NaN
1993	0	0	NaN	0	NaN	1993	0	0	NaN	0	NaN
1994	0	0	NaN	0	NaN	1994	9	0	0.00	0	0.00
1995	0	0	NaN	0	NaN	1995	82	0	0.00	2	2.38
1996	26	0	0.00	0	0.00	1996	627	0	0.00	7	1.10
1997	33	0	0.00	3	8.33	1997	632	0	0.00	21	3.22
1998	50	0	0.00	3	5.66	1998	610	0	0.00	45	6.87
1999	59	0	0.00	9	13.24	1999	622	0	0.00	50	7.44
2000	58	0	0.00	17	22.67	2000	594	0	0.00	84	12.39
2001	64	0	0.00	12	15.79	2001	582	0	0.00	171	22.71
2002	55	0	0.00	17	23.61	2002	541	2	0.28	176	24.48
2003	40	0	0.00	22	35.48	2003	525	0	0.00	157	23.02
2004	32	0	0.00	13	28.89	2004	510	2	0.34	85	14.24
2005	25	0	0.00	11	30.56	2005	524	0	0.00	47	8.23
2006	25	0	0.00	5	16.67	2006	552	0	0.00	39	6.60
2007	26	0	0.00	5	16.13	2007	592	0	0.00	28	4.52
2008	11	4	13.79	14	48.28	2008	565	0	0.00	46	7.53
2009	10	1	6.67	4	26.67	2009	562	0	0.00	30	5.07
2010	8	0	0.00	4	33.33	2010	553	1	0.17	43	7.20
2011	9	0	0.00	4	30.77	2011	538	1	0.17	50	8.49
2012	12	0	0.00	1	7.69	2012	485	0	0.00	77	13.70

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144 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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				ny:ITA	conor	E	
)th	(aults	De		
Ye			#	%	#	Active	Year
19	3		5	0.00	0	185	1992
19	3		8	0.00	0	181	1993
19	5		13	0.00	0	195	1994
19	8		14	0.00	0	209	1995
19	8		15	0.43	1	219	1996
19	8		28	0.00	0	218	1997
19	1		14	0.00	0	227	1998
19	0		7	0.00	0	252	1999
20	8		25	0.00	0	277	2000
20	2		17	0.00	0	280	2001
20	8		14	0.34	1	278	2002
20	1		23	2.09	6	258	2003
20	3		7	0.75	2	257	2004
20	8		15	0.00	0	264	2005
20	8		16	0.00	0	276	2006
20	1		13	0.00	0	296	2007
20	2		22	0.32	1	286	2008
20	2		20	1.32	4	278	2009
20	3		15	0.00	0	283	2010
20	8		26	0.00	0	277	2011
20	1		29	0.66	2	274	2012

Table A.11.(Continued)

	E	conor	ny:JOR		
		De	faults	0	thers
Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN
1993	0	0	NaN	0	NaN
1994	0	0	NaN	0	NaN
1995	0	0	NaN	0	NaN
1996	0	0	NaN	0	NaN
1997	0	0	NaN	0	NaN
1998	0	0	NaN	0	NaN
1999	0	0	NaN	0	NaN
2000	0	0	NaN	0	NaN
2001	0	0	NaN	0	NaN
2002	0	0	NaN	0	NaN
2003	0	0	NaN	0	NaN
2004	0	0	NaN	0	NaN
2005	0	0	NaN	0	NaN
2006	0	0	NaN	0	NaN
2007	204	0	0.00	13	5.99
2008	217	0	0.00	17	7.26
2009	223	0	0.00	23	9.35
2010	215	0	0.00	29	11.89
2011	210	0	0.00	27	11.39
2012	208	0	0.00	31	12.97
	E	conor	ny:KAZ		

	F	Econor	ny:JPN				F	Econo	my:KAZ		
		De	faults	0	thers			De	efaults	C	others
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	2535	2	0.08	21	0.82	1992	0	0	NaN	0	NaN
1993	2616	3	0.11	24	0.91	1993	0	0	NaN	0	NaN
1994	2762	0	0.00	17	0.61	1994	0	0	NaN	0	NaN
1995	2943	1	0.03	17	0.57	1995	0	0	NaN	0	NaN
1996	3096	4	0.13	21	0.67	1996	0	0	NaN	0	NaN
1997	3215	6	0.18	31	0.95	1997	0	0	NaN	0	NaN
1998	3269	12	0.36	37	1.12	1998	0	0	NaN	0	NaN
1999	3332	6	0.18	47	1.39	1999	0	0	NaN	0	NaN
2000	3474	12	0.34	59	1.66	2000	0	0	NaN	0	NaN
2001	3581	16	0.44	64	1.75	2001	0	0	NaN	0	NaN
2002	3605	29	0.78	93	2.50	2002	7	0	0.00	6	46.15
2003	3636	20	0.53	98	2.61	2003	9	0	0.00	4	30.77
2004	3739	11	0.29	85	2.22	2004	12	0	0.00	13	52.00
2005	3824	9	0.23	89	2.27	2005	3	0	0.00	13	81.25
2006	3953	2	0.05	85	2.10	2006	2	0	0.00	4	66.67
2007	3982	6	0.15	104	2.54	2007	23	0	0.00	12	34.29
2008	3907	35	0.86	107	2.64	2008	21	0	0.00	15	41.67
2009	3779	28	0.71	136	3.45	2009	14	4	10.26	21	53.85
2010	3682	9	0.24	129	3.38	2010	11	0	0.00	17	60.71
2011	3627	5	0.13	100	2.68	2011	13	0	0.00	10	43.48
2012	3584	6	0.16	101	2.74	2012	15	0	0.00	10	40.00

GLOBAL CREDIT REVIEW VOLUME 3 145

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	Ε	conon	ıy:KOR	1			E	conor	ny:KWI		
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	635	0	0.00	1	0.16	1992	0	0	NaN	0	NaN
1993	643	0	0.00	0	0.00	1993	0	0	NaN	0	NaN
1994	674	0	0.00	0	0.00	1994	0	0	NaN	0	NaN
1995	701	0	0.00	2	0.28	1995	0	0	NaN	0	NaN
1996	747	6	0.79	3	0.40	1996	51	0	0.00	2	3.77
1997	1003	50	4.67	17	1.59	1997	66	0	0.00	1	1.49
1998	915	78	7.08	109	9.89	1998	67	0	0.00	1	1.47
1999	1012	22	2.05	39	3.63	1999	72	0	0.00	7	8.86
2000	1156	12	0.99	44	3.63	2000	70	0	0.00	10	12.50
2001	1302	18	1.34	26	1.93	2001	73	0	0.00	0	0.00
2002	1450	14	0.94	32	2.14	2002	76	0	0.00	6	7.32
2003	1494	13	0.85	31	2.02	2003	94	0	0.00	0	0.00
2004	1510	7	0.45	55	3.50	2004	104	0	0.00	4	3.70
2005	1560	8	0.49	52	3.21	2005	139	0	0.00	4	2.80
2006	1634	2	0.12	12	0.73	2006	156	0	0.00	6	3.70
2007	1706	0	0.00	16	0.93	2007	176	0	0.00	13	6.88
2008	1734	8	0.45	34	1.91	2008	176	1	0.52	17	8.76
2009	1724	7	0.38	90	4.94	2009	182	1	0.47	31	14.49
2010	1745	8	0.43	95	5.14	2010	182	0	0.00	44	19.47
2011	1750	3	0.16	89	4.83	2011	164	1	0.44	62	27.3
2012	1723	7	0.39	75	4.16	2012	184	0	0.00	38	17.12

Table A.11. (Continued)

	Ε	conor	ny:LKA				E	cono	my:LTU		
		De	faults	C	thers			De	faults	0)
Year	Active	#	%	#	%	Year	Active	#	%	#	-
1992	0	0	NaN	0	NaN	1992	0	0	NaN	0	-
1993	1	0	0.00	0	0.00	1993	0	0	NaN	0	
1994	0	0	0.00	1	100.00	1994	0	0	NaN	0	
1995	139	0	0.00	1	0.71	1995	0	0	NaN	0	
1996	132	0	0.00	42	24.14	1996	0	0	NaN	0	
1997	140	0	0.00	33	19.08	1997	0	0	NaN	0	
1998	153	0	0.00	30	16.39	1998	0	0	NaN	0	
1999	149	0	0.00	35	19.02	1999	0	0	NaN	0	
2000	149	0	0.00	35	19.02	2000	34	0	0.00	5	
2001	163	0	0.00	25	13.30	2001	32	0	0.00	12	
2002	172	0	0.00	24	12.24	2002	42	0	0.00	4	
2003	176	0	0.00	28	13.73	2003	39	0	0.00	7	
2004	191	0	0.00	10	4.98	2004	40	0	0.00	1	
2005	204	0	0.00	10	4.67	2005	40	0	0.00	0	
2006	210	0	0.00	12	5.41	2006	39	0	0.00	2	
2007	216	0	0.00	10	4.42	2007	37	0	0.00	3	
2008	215	0	0.00	13	5.70	2008	38	0	0.00	0	
2009	219	0	0.00	10	4.37	2009	36	0	0.00	3	
2010	239	0	0.00	0	0.00	2010	37	0	0.00	2	
2011	259	0	0.00	9	3.36	2011	32	1	2.63	5	
2012	276	0	0.00	4	1.43	2012	32	0	0.00	0	

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 $146 \quad {\rm nus-rmi\ credit\ research\ initiative\ technical\ report}$

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	E	conor	ny:LUX				E	Cono	my:LVA		
		De	faults	0	thers			De	efaults	C	Others
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	2	0	0.00	1	33.33	1992	0	0	NaN	0	Nal
1993	2	0	0.00	1	33.33	1993	0	0	NaN	0	Nal
1994	2	0	0.00	0	0.00	1994	0	0	NaN	0	Nal
1995	30	0	0.00	12	28.57	1995	0	0	NaN	0	Nal
1996	28	0	0.00	15	34.88	1996	0	0	NaN	0	Nal
1997	37	0	0.00	9	19.57	1997	0	0	NaN	0	Nal
1998	34	0	0.00	13	27.66	1998	0	0	NaN	0	Nal
1999	34	0	0.00	14	29.17	1999	0	0	NaN	0	Nal
2000	32	0	0.00	13	28.89	2000	12	0	0.00	5	29.4
2001	28	0	0.00	13	31.71	2001	33	0	0.00	6	15.3
2002	27	0	0.00	10	27.03	2002	35	0	0.00	3	7.8
2003	28	0	0.00	11	28.21	2003	26	0	0.00	11	29.7
2004	36	0	0.00	7	16.28	2004	29	0	0.00	3	9.3
2005	37	0	0.00	7	15.91	2005	33	0	0.00	3	8.3
2006	36	0	0.00	16	30.77	2006	31	0	0.00	5	13.8
2007	34	0	0.00	8	19.05	2007	32	0	0.00	6	15.7
2008	26	0	0.00	14	35.00	2008	27	0	0.00	7	20.5
2009	23	0	0.00	8	25.81	2009	28	0	0.00	9	24.3
2010	18	1	3.45	10	34.48	2010	31	0	0.00	4	11.4
2011	14	0	0.00	8	36.36	2011	28	1	2.78	7	19.4
2012	14	0	0.00	5	26.32	2012	29	0	0.00	5	14.7

Table A.11. (Continued)

	E	conon	ny:MAR				E	conor	ny:MEX	-	
		De	faults	C	Others			De	efaults	0	ther
Year	Active	#	%	#	%	Year	Active	#	%	#	Ģ
1992	0	0	NaN	0	NaN	1992	0	0	NaN	0	N
1993	0	0	NaN	0	NaN	1993	0	0	NaN	0	Ν
1994	0	0	NaN	0	NaN	1994	90	0	0.00	37	29
1995	0	0	NaN	0	NaN	1995	97	0	0.00	22	18
1996	16	0	0.00	0	0.00	1996	107	0	0.00	19	15
1997	40	0	0.00	3	6.98	1997	117	0	0.00	23	16
1998	48	0	0.00	1	2.04	1998	113	0	0.00	21	15
1999	46	0	0.00	6	11.54	1999	112	1	0.74	23	16
2000	51	0	0.00	1	1.92	2000	107	1	0.79	18	14
2001	52	0	0.00	7	11.86	2001	110	1	0.77	19	14
2002	52	0	0.00	7	11.86	2002	98	1	0.78	29	22
2003	49	0	0.00	8	14.04	2003	103	2	1.65	16	13
2004	48	0	0.00	7	12.73	2004	108	0	0.00	11	9
2005	52	0	0.00	3	5.45	2005	101	0	0.00	23	18
2006	59	0	0.00	5	7.81	2006	105	0	0.00	10	8
2007	72	0	0.00	2	2.70	2007	103	0	0.00	18	14
2008	77	0	0.00	1	1.28	2008	95	2	1.67	23	19
2009	75	0	0.00	2	2.60	2009	102	3	2.61	10	8
2010	73	0	0.00	4	5.19	2010	107	2	1.63	14	11
2011	75	0	0.00	2	2.60	2011	104	0	0.00	23	18
2012	76	0	0.00	0	0.00	2012	105	0	0.00	11	9

% NaN NaN NaN NaN NaN NaN NaN NaN 29.41 15.38 7.89 29.73 9.38 8.33 13.89 15.79 20.59 24.32 11.43 19.44 14.71

GLOBAL CREDIT REVIEW VOLUME 3 147

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	E	conon	ny:MKD)			E	conor	ny:MLT		
		De	faults	0	thers			De	faults	C	Others
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN	1992	0	0	NaN	0	NaN
1993	0	0	NaN	0	NaN	1993	0	0	NaN	0	NaN
1994	0	0	NaN	0	NaN	1994	0	0	NaN	0	NaN
1995	0	0	NaN	0	NaN	1995	0	0	NaN	0	NaN
1996	0	0	NaN	0	NaN	1996	5	0	0.00	0	0.00
1997	0	0	NaN	0	NaN	1997	6	0	0.00	0	0.00
1998	0	0	NaN	0	NaN	1998	7	0	0.00	0	0.00
1999	0	0	NaN	0	NaN	1999	7	0	0.00	1	12.50
2000	0	0	NaN	0	NaN	2000	9	0	0.00	0	0.00
2001	0	0	NaN	0	NaN	2001	9	0	0.00	2	18.18
2002	0	0	NaN	0	NaN	2002	10	0	0.00	2	16.67
2003	0	0	NaN	0	NaN	2003	10	0	0.00	3	23.08
2004	13	0	0.00	0	0.00	2004	11	0	0.00	2	15.38
2005	62	0	0.00	70	53.03	2005	11	0	0.00	2	15.38
2006	84	0	0.00	69	45.10	2006	13	0	0.00	0	0.00
2007	92	0	0.00	75	44.91	2007	14	0	0.00	3	17.65
2008	73	0	0.00	76	51.01	2008	13	0	0.00	7	35.00
2009	68	0	0.00	66	49.25	2009	12	0	0.00	3	20.00
2010	66	0	0.00	61	48.03	2010	12	0	0.00	2	14.29
2011	58	0	0.00	64	52.46	2011	14	0	0.00	2	12.50
2012	51	1	0.85	65	55.56	2012	19	0	0.00	1	5.00

Table A.11. (Continued)

	Ε	conon	ny:MYS				Ε	conor	ny:NGA		
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	356	0	0.00	10	2.73	1992	0	0	NaN	0	NaN
1993	405	0	0.00	0	0.00	1993	0	0	NaN	0	NaN
1994	466	0	0.00	7	1.48	1994	0	0	NaN	0	NaN
1995	523	0	0.00	2	0.38	1995	0	0	NaN	0	NaN
1996	615	0	0.00	0	0.00	1996	0	0	NaN	0	NaN
1997	702	0	0.00	2	0.28	1997	0	0	NaN	0	NaN
1998	696	14	1.92	21	2.87	1998	0	0	NaN	0	NaN
1999	704	8	1.10	14	1.93	1999	0	0	NaN	0	NaN
2000	735	8	1.06	12	1.59	2000	0	0	NaN	0	NaN
2001	739	10	1.30	18	2.35	2001	0	0	NaN	0	NaN
2002	767	7	0.87	30	3.73	2002	114	0	0.00	31	21.38
2003	835	3	0.35	20	2.33	2003	77	0	0.00	58	42.96
2004	911	3	0.32	17	1.83	2004	115	0	0.00	41	26.28
2005	984	0	0.00	27	2.67	2005	127	0	0.00	27	17.53
2006	994	5	0.48	36	3.48	2006	139	0	0.00	30	17.75
2007	968	6	0.58	65	6.26	2007	171	0	0.00	25	12.76
2008	927	17	1.68	66	6.53	2008	177	0	0.00	47	20.98
2009	923	12	1.23	41	4.20	2009	190	0	0.00	22	10.38
2010	922	16	1.64	36	3.70	2010	184	0	0.00	24	11.54
2011	924	2	0.21	33	3.44	2011	160	0	0.00	38	19.19
2012	907	4	0.42	39	4.11	2012	162	0	0.00	17	9.50

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148 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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	E	conon	ny:NLD				E	conor	ny:NOR	l	
		Def	faults	C	Others			De	faults	C	Others
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	161	0	0.00	6	3.59	1992	81	0	0.00	8	8.9
1993	171	0	0.00	5	2.84	1993	101	0	0.00	2	1.94
1994	174	0	0.00	5	2.79	1994	115	0	0.00	3	2.54
1995	187	0	0.00	5	2.60	1995	137	0	0.00	2	1.44
1996	196	1	0.50	4	1.99	1996	159	0	0.00	4	2.4
1997	198	0	0.00	17	7.91	1997	201	0	0.00	12	5.6
1998	210	0	0.00	10	4.55	1998	216	0	0.00	19	8.0
1999	213	0	0.00	19	8.19	1999	201	0	0.00	28	12.2
2000	201	1	0.45	21	9.42	2000	195	1	0.44	33	14.4
2001	176	7	3.40	23	11.17	2001	213	3	1.21	32	12.9
2002	158	10	5.43	16	8.70	2002	206	4	1.63	35	14.2
2003	153	0	0.00	14	8.38	2003	184	3	1.31	42	18.3
2004	145	0	0.00	11	7.05	2004	198	0	0.00	18	8.3
2005	141	0	0.00	8	5.37	2005	235	0	0.00	23	8.9
2006	137	1	0.68	8	5.48	2006	262	0	0.00	37	12.3
2007	134	0	0.00	9	6.29	2007	270	0	0.00	47	14.8
2008	126	1	0.72	11	7.97	2008	243	4	1.38	42	14.5
2009	122	3	2.33	4	3.10	2009	225	6	2.23	38	14.1
2010	120	0	0.00	5	4.00	2010	224	1	0.40	24	9.64
2011	115	0	0.00	8	6.50	2011	221	1	0.43	13	5.5
2012	111	0	0.00	9	7.50	2012	213	1	0.44	14	6.14

Table A.11.(Continued)

	E	conor	ny:NZL				E	cono	my:PAK	-	
		De	faults	C	Others			De	efaults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	29	0	0.00	1	3.33	1992	0	0	NaN	0	NaN
1993	33	0	0.00	0	0.00	1993	0	0	NaN	0	NaN
1994	41	0	0.00	0	0.00	1994	0	0	NaN	0	NaN
1995	42	0	0.00	2	4.55	1995	0	0	NaN	0	NaN
1996	46	0	0.00	2	4.17	1996	0	0	NaN	0	NaN
1997	49	0	0.00	0	0.00	1997	0	0	NaN	0	NaN
1998	50	0	0.00	1	1.96	1998	248	0	0.00	132	34.74
1999	56	0	0.00	0	0.00	1999	371	0	0.00	112	23.19
2000	63	0	0.00	1	1.56	2000	381	0	0.00	157	29.18
2001	72	0	0.00	0	0.00	2001	326	0	0.00	180	35.57
2002	77	0	0.00	0	0.00	2002	431	0	0.00	97	18.37
2003	89	0	0.00	0	0.00	2003	469	0	0.00	61	11.51
2004	103	0	0.00	1	0.96	2004	503	0	0.00	48	8.71
2005	108	0	0.00	0	0.00	2005	491	0	0.00	89	15.34
2006	114	0	0.00	0	0.00	2006	472	0	0.00	97	17.05
2007	124	0	0.00	1	0.80	2007	492	0	0.00	77	13.53
2008	114	0	0.00	14	10.94	2008	304	0	0.00	296	49.33
2009	122	0	0.00	9	6.87	2009	518	0	0.00	68	11.60
2010	121	0	0.00	15	11.03	2010	491	0	0.00	53	9.74
2011	122	0	0.00	12	8.96	2011	503	0	0.00	105	17.27
2012	121	0	0.00	18	12.95	2012	454	0	0.00	87	16.08

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% 8.99 1.94 2.54 1.44 2.45 5.63 8.09 12.23 14.41 12.90 14.29 18.34 8.33 8.91 12.37 14.83 14.53 14.13 9.64 5.53 6.14

GLOBAL CREDIT REVIEW VOLUME 3 149

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	Ε	conor	ny:PER				E	conor	ny:PHL		
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	1	0	0.00	0	0.00	1992	80	0	0.00	24	23.08
1993	1	0	0.00	0	0.00	1993	108	1	0.80	16	12.80
1994	63	0	0.00	2	3.08	1994	129	0	0.00	26	16.77
1995	97	0	0.00	22	18.49	1995	158	0	0.00	15	8.67
1996	94	0	0.00	46	32.86	1996	178	0	0.00	14	7.29
1997	118	0	0.00	40	25.32	1997	186	0	0.00	22	10.58
1998	107	0	0.00	59	35.54	1998	179	1	0.48	28	13.40
1999	92	0	0.00	67	42.14	1999	186	4	1.98	12	5.94
2000	74	0	0.00	74	50.00	2000	173	1	0.48	36	17.14
2001	64	0	0.00	54	45.76	2001	170	3	1.42	39	18.40
2002	75	0	0.00	50	40.00	2002	152	6	2.82	55	25.82
2003	66	0	0.00	46	41.07	2003	173	4	1.90	33	15.7
2004	75	0	0.00	40	34.78	2004	177	7	3.04	46	20.00
2005	78	0	0.00	42	35.00	2005	181	3	1.35	38	17.12
2006	78	0	0.00	40	33.90	2006	188	2	0.92	28	12.84
2007	90	0	0.00	26	22.41	2007	193	2	0.90	27	12.16
2008	79	0	0.00	50	38.76	2008	183	1	0.46	34	15.60
2009	94	0	0.00	33	25.98	2009	203	2	0.90	18	8.07
2010	89	0	0.00	31	25.83	2010	207	0	0.00	18	8.00
2011	78	0	0.00	39	33.33	2011	220	0	0.00	14	5.98
2012	82	0	0.00	37	31.09	2012	215	0	0.00	22	9.28

 Table A.11 (Continued)

	Ε	conor	ny:POL				E	conor	ny:PRT		
		De	faults	C	Others			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN	1992	1	0	0.00	0	0.00
1993	0	0	NaN	0	NaN	1993	69	0	0.00	13	15.85
1994	31	0	0.00	19	38.00	1994	79	0	0.00	12	13.19
1995	59	0	0.00	0	0.00	1995	92	0	0.00	18	16.36
1996	77	0	0.00	0	0.00	1996	96	0	0.00	21	17.95
1997	137	0	0.00	3	2.14	1997	93	0	0.00	29	23.77
1998	191	0	0.00	3	1.55	1998	84	0	0.00	34	28.81
1999	212	0	0.00	3	1.40	1999	88	0	0.00	24	21.43
2000	218	1	0.44	6	2.67	2000	84	0	0.00	17	16.83
2001	220	1	0.44	5	2.21	2001	69	0	0.00	20	22.47
2002	203	2	0.88	22	9.69	2002	62	0	0.00	19	23.46
2003	191	3	1.44	14	6.73	2003	65	0	0.00	6	8.45
2004	212	0	0.00	8	3.64	2004	67	0	0.00	7	9.46
2005	233	1	0.41	9	3.70	2005	64	0	0.00	7	9.86
2006	252	0	0.00	9	3.45	2006	62	0	0.00	11	15.07
2007	327	0	0.00	10	2.97	2007	58	0	0.00	9	13.43
2008	425	0	0.00	5	1.16	2008	57	0	0.00	8	12.31
2009	453	0	0.00	11	2.37	2009	56	0	0.00	9	13.85
2010	548	0	0.00	8	1.44	2010	56	0	0.00	7	11.11
2011	732	0	0.00	14	1.88	2011	54	2	3.23	6	9.68
2012	814	8	0.94	31	3.63	2012	54	0	0.00	5	8.47

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150 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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	E	conon	ny:ROM	[E	conor	my:RUS		
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN	1992	0	0	NaN	0	Nal
1993	0	0	NaN	0	NaN	1993	0	0	NaN	0	NaN
1994	0	0	NaN	0	NaN	1994	0	0	NaN	0	NaN
1995	0	0	NaN	0	NaN	1995	0	0	NaN	0	NaN
1996	0	0	NaN	0	NaN	1996	0	0	NaN	0	NaN
1997	0	0	NaN	0	NaN	1997	74	0	0.00	28	27.4
1998	78	0	0.00	1	1.27	1998	35	1	0.76	96	72.7
1999	340	0	0.00	30	8.11	1999	38	0	0.00	42	52.50
2000	355	0	0.00	100	21.98	2000	53	0	0.00	70	56.9
2001	333	0	0.00	177	34.71	2001	84	0	0.00	63	42.8
2002	279	0	0.00	197	41.39	2002	55	0	0.00	111	66.8
2003	275	0	0.00	165	37.50	2003	78	0	0.00	64	45.0
2004	311	0	0.00	113	26.65	2004	105	4	2.33	63	36.6
2005	305	1	0.20	198	39.29	2005	169	0	0.00	83	32.94
2006	514	0	0.00	160	23.74	2006	226	1	0.25	176	43.6
2007	719	0	0.00	698	49.26	2007	317	0	0.00	177	35.8
2008	557	0	0.00	444	44.36	2008	273	1	0.19	253	48.0
2009	445	0	0.00	445	50.00	2009	298	8	1.92	110	26.4
2010	509	0	0.00	371	42.16	2010	306	1	0.23	131	29.9
2011	444	1	0.11	438	49.60	2011	271	2	0.43	195	41.6
2012	347	0	0.00	349	50.14	2012	482	1	0.16	144	22.9

Table A.11. (Continued)

	E	conor	my:SAU				F	Conor	ny:SGP		
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN	1992	177	0	0.00	11	5.85
1993	0	0	NaN	0	NaN	1993	203	0	0.00	3	1.46
1994	0	0	NaN	0	NaN	1994	233	0	0.00	4	1.69
1995	0	0	NaN	0	NaN	1995	250	1	0.39	6	2.33
1996	0	0	NaN	0	NaN	1996	272	1	0.36	8	2.85
1997	0	0	NaN	0	NaN	1997	299	1	0.31	18	5.66
1998	0	0	NaN	0	NaN	1998	318	4	1.19	14	4.17
1999	0	0	NaN	0	NaN	1999	358	4	1.06	15	3.98
2000	62	0	0.00	4	6.06	2000	428	0	0.00	18	4.04
2001	63	0	0.00	4	5.97	2001	437	2	0.43	31	6.60
2002	66	0	0.00	3	4.35	2002	450	2	0.41	34	7.00
2003	69	0	0.00	2	2.82	2003	500	1	0.19	15	2.91
2004	72	0	0.00	0	0.00	2004	576	2	0.34	9	1.53
2005	76	0	0.00	0	0.00	2005	623	4	0.62	17	2.64
2006	86	0	0.00	0	0.00	2006	667	1	0.14	25	3.61
2007	109	0	0.00	2	1.80	2007	711	0	0.00	19	2.60
2008	126	0	0.00	1	0.79	2008	703	4	0.53	41	5.48
2009	134	0	0.00	1	0.74	2009	717	14	1.84	31	4.07
2010	145	0	0.00	0	0.00	2010	721	0	0.00	36	4.76
2011	148	0	0.00	1	0.67	2011	704	0	0.00	53	7.00
2012	156	0	0.00	1	0.64	2012	700	0	0.00	40	5.41

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 ${\rm GLOBAL\ CREDIT\ REVIEW\ VOLUME\ 3} \quad 151$

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	E	conor	ny:SVK				E	cono	my:SVN		
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN	1992	0	0	NaN	0	Nal
1993	0	0	NaN	0	NaN	1993	0	0	NaN	0	Nal
1994	0	0	NaN	0	NaN	1994	0	0	NaN	0	Nal
1995	0	0	NaN	0	NaN	1995	0	0	NaN	0	Nal
1996	0	0	NaN	0	NaN	1996	0	0	NaN	0	Nal
1997	0	0	NaN	0	NaN	1997	0	0	NaN	0	Nal
1998	7	0	0.00	13	65.00	1998	75	0	0.00	8	9.6
1999	8	0	0.00	24	75.00	1999	98	0	0.00	9	8.4
2000	12	0	0.00	11	47.83	2000	94	0	0.00	34	26.5
2001	15	0	0.00	14	48.28	2001	112	0	0.00	29	20.5
2002	23	0	0.00	15	39.47	2002	103	0	0.00	36	25.9
2003	40	0	0.00	26	39.39	2003	104	0	0.00	19	15.4
2004	44	0	0.00	31	41.33	2004	108	0	0.00	22	16.9
2005	41	0	0.00	26	38.81	2005	87	0	0.00	34	28.1
2006	48	0	0.00	42	46.67	2006	77	0	0.00	28	26.6
2007	24	0	0.00	47	66.20	2007	65	0	0.00	22	25.2
2008	34	0	0.00	26	43.33	2008	67	0	0.00	15	18.2
2009	38	0	0.00	31	44.93	2009	60	2	2.30	25	28.7
2010	55	0	0.00	21	27.63	2010	69	0	0.00	11	13.7
2011	53	0	0.00	42	44.21	2011	59	1	1.33	15	20.0
2012	47	0	0.00	31	39.74	2012	56	1	1.67	3	5.0

Table A.11. (Continued)

	E	conor	ny:SWE				E	conon	ny:THA		
		De	faults	Ot	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	121	0	0.00	2	1.63	1992	280	0	0.00	1	0.36
1993	144	0	0.00	2	1.37	1993	331	0	0.00	1	0.30
1994	171	0	0.00	2	1.16	1994	378	0	0.00	1	0.26
1995	184	0	0.00	0	0.00	1995	402	1	0.24	9	2.18
1996	223	0	0.00	15	6.30	1996	420	7	1.56	21	4.69
1997	270	0	0.00	28	9.40	1997	370	21	4.59	67	14.63
1998	299	1	0.31	19	5.96	1998	346	16	3.87	51	12.35
1999	338	1	0.27	27	7.38	1999	322	15	4.04	34	9.16
2000	370	1	0.25	35	8.62	2000	303	18	5.19	26	7.49
2001	362	4	1.01	31	7.81	2001	304	8	2.42	19	5.74
2002	350	7	1.81	29	7.51	2002	322	4	1.17	17	4.96
2003	338	3	0.82	26	7.08	2003	352	3	0.83	8	2.20
2004	353	1	0.27	22	5.85	2004	392	1	0.24	21	5.07
2005	389	2	0.50	12	2.98	2005	428	3	0.66	22	4.86
2006	433	0	0.00	22	4.84	2006	443	0	0.00	12	2.64
2007	502	1	0.19	14	2.71	2007	443	2	0.43	19	4.09
2008	510	2	0.37	30	5.54	2008	441	0	0.00	24	5.16
2009	497	4	0.75	32	6.00	2009	455	7	1.49	8	1.70
2010	501	2	0.38	29	5.45	2010	458	4	0.85	9	1.91
2011	495	3	0.56	34	6.39	2011	463	1	0.21	10	2.11
2012	474	0	0.00	47	9.02	2012	478	0	0.00	7	1.44

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152 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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	E	conor	ny:TUR				E	conor	ny:TWN	1	
		De	faults	0	thers			De	faults		C
Year	Active	#	%	#	%	Year	Active	#	%	#	
1992	8	0	0.00	0	0.00	1992	236	0	0.00	2	_
1993	15	0	0.00	0	0.00	1993	259	0	0.00	1	
1994	33	0	0.00	1	2.94	1994	291	0	0.00	1	
1995	200	0	0.00	3	1.48	1995	371	0	0.00	0	
1996	220	0	0.00	3	1.35	1996	442	0	0.00	1	
1997	256	0	0.00	8	3.03	1997	496	0	0.00	3	
1998	275	0	0.00	3	1.08	1998	572	4	0.68	12	
1999	273	0	0.00	13	4.55	1999	695	7	0.99	8	
2000	294	2	0.64	16	5.13	2000	794	8	0.98	17	
2001	282	0	0.00	17	5.69	2001	886	9	0.98	23	
2002	285	0	0.00	7	2.40	2002	996	7	0.67	38	
2003	283	0	0.00	6	2.08	2003	1082	2	0.18	21	
2004	295	0	0.00	0	0.00	2004	1350	6	0.43	28	
2005	302	0	0.00	3	0.98	2005	1363	8	0.56	67	
2006	313	0	0.00	6	1.88	2006	1389	3	0.21	41	
2007	317	0	0.00	5	1.55	2007	1444	3	0.20	35	
2008	314	0	0.00	5	1.57	2008	1446	9	0.60	49	
2009	314	0	0.00	4	1.26	2009	1503	4	0.26	22	
2010	336	0	0.00	0	0.00	2010	1601	2	0.12	24	
2011	361	0	0.00	3	0.82	2011	1666	1	0.06	40	
2012	391	0	0.00	9	2.25	2012	1722	2	0.11	48	

Table A.11. (Continued)

	Ε	conor	ny:UKR				1	Econon	ny:USA		
		De	faults	0	thers			Def	aults	0	the
Year	Active	#	%	#	%	Year	Active	#	%	#	
1992	0	0	NaN	0	NaN	1992	5321	18	0.33	107	
1993	0	0	NaN	0	NaN	1993	6006	25	0.40	172	
1994	0	0	NaN	0	NaN	1994	6666	18	0.26	279	
1995	0	0	NaN	0	NaN	1995	7031	20	0.27	389	
1996	0	0	NaN	0	NaN	1996	7582	19	0.24	404	
1997	0	0	NaN	0	NaN	1997	7769	54	0.64	564	
1998	20	0	0.00	31	60.78	1998	7405	85	1.02	872	1
1999	38	0	0.00	37	49.33	1999	7058	89	1.10	929	1
2000	58	0	0.00	30	34.09	2000	6769	133	1.73	786	1
2001	25	0	0.00	75	75.00	2001	6064	199	2.85	716	1
2002	12	0	0.00	35	74.47	2002	5616	137	2.19	506	
2003	16	0	0.00	18	52.94	2003	5276	91	1.56	462	
2004	33	0	0.00	24	42.11	2004	5247	34	0.60	378	
2005	56	0	0.00	29	34.12	2005	5211	39	0.69	380	
2006	120	0	0.00	44	26.83	2006	5171	21	0.38	379	
2007	157	0	0.00	78	33.19	2007	5115	29	0.52	447	
2008	103	0	0.00	129	55.60	2008	4817	77	1.46	367	
2009	76	1	0.48	131	62.98	2009	4561	96	1.93	313	
2010	53	0	0.00	73	57.94	2010	4484	35	0.73	308	
2011	61	0	0.00	33	35.11	2011	4327	38	0.81	312	
2012	58	0	0.00	42	42.00	2012	4251	37	0.81	263	:

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GLOBAL CREDIT REVIEW VOLUME 3 153

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	Ε	conor	ny:VEN				E	conor	ny:VNM	[
		De	faults	0	thers			De	faults	0	thers
Year	Active	#	%	#	%	Year	Active	#	%	#	%
1992	0	0	NaN	0	NaN	1992	0	0	NaN	0	NaN
1993	7	0	0.00	0	0.00	1993	0	0	NaN	0	NaN
1994	12	0	0.00	1	7.69	1994	0	0	NaN	0	NaN
1995	15	0	0.00	3	16.67	1995	0	0	NaN	0	NaN
1996	15	0	0.00	2	11.76	1996	0	0	NaN	0	NaN
1997	47	0	0.00	17	26.56	1997	0	0	NaN	0	NaN
1998	40	0	0.00	25	38.46	1998	0	0	NaN	0	NaN
1999	39	0	0.00	18	31.58	1999	0	0	NaN	0	NaN
2000	37	0	0.00	12	24.49	2000	5	0	0.00	0	0.00
2001	30	1	2.33	12	27.91	2001	10	0	0.00	0	0.00
2002	20	0	0.00	19	48.72	2002	20	0	0.00	0	0.00
2003	25	0	0.00	10	28.57	2003	22	0	0.00	0	0.00
2004	27	0	0.00	9	25.00	2004	25	0	0.00	0	0.00
2005	29	0	0.00	7	19.44	2005	31	0	0.00	0	0.00
2006	27	0	0.00	7	20.59	2006	91	0	0.00	0	0.00
2007	24	0	0.00	8	25.00	2007	222	0	0.00	3	1.33
2008	24	0	0.00	32	57.14	2008	296	0	0.00	2	0.67
2009	26	0	0.00	21	44.68	2009	397	0	0.00	29	6.81
2010	23	0	0.00	14	37.84	2010	675	0	0.00	26	3.71
2011	30	0	0.00	16	34.78	2011	739	1	0.13	52	6.57
2012	15	0	0.00	18	54.55	2012	739	0	0.00	86	10.42

 Table A.11. (Continued)

		Defaults		Others		
Year	Active	#	%	#	# % 0 0.00 89 9.01 22 4.98 25 5.08 13 2.57 23 4.11 50 9.45 48 7.45 56 10.89 98 17.28 13 24.20 49 13.00 48 14.16 36 10.88 23 6.91 40 10.67 26 7.18 31 8.81	
1992	395	0	0.00	0	0.00	
1993	394	0	0.00	39	9.01	
1994	420	0	0.00	22	4.98	
1995	467	0	0.00	25	5.08	
1996	492	0	0.00	13	2.57	
1997	537	0	0.00	23	4.11	
1998	573	2	0.31	60	9.45	
1999	593	3	0.47	48	7.45	
2000	534	6	0.99	66	10.89	
2001	460	9	1.59	98	17.28	
2002	346	8	1.71	113	24.20	
2003	327	1	0.27	49	13.00	
2004	289	2	0.59	48	14.16	
2005	293	2	0.60	36	10.88	
2006	310	0	0.00	23	6.91	
2007	335	0	0.00	40	10.67	
2008	336	0	0.00	26	7.18	
2009	320	1	0.28	31	8.81	
2010	312	2	0.60	22	6.55	
2011	310	1	0.30	21	6.33	
2012	294	5	1.56	22	6.85	

154 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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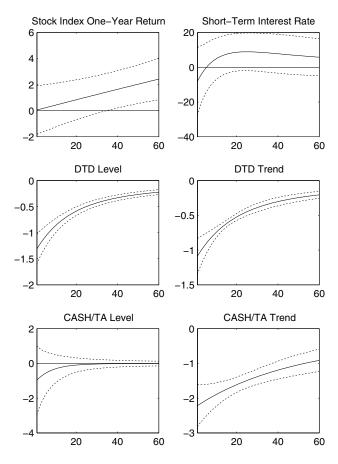
APPENDIX B: PERFORMANCE ANALYSIS

Economy	AR				AUROC				
	1 mth	1 yr	2 yr	5 yr	1 mth	1 yr	2 yr	5 yr	
AUS	0.809	0.641	0.513	0.359	0.905	0.821	0.758	0.684	
CHN	0.582	0.515	0.435	0.308	0.791	0.760	0.724	0.673	
HKG	0.691	0.417	0.318	0.205	0.846	0.709	0.660	0.606	
IND	0.701	0.617	0.523	0.441	0.850	0.809	0.762	0.722	
IDN	0.771	0.688	0.576	0.435	0.886	0.845	0.790	0.725	
JPN	0.910	0.824	0.765	0.635	0.955	0.912	0.883	0.819	
MYS	0.842	0.734	0.644	0.406	0.921	0.867	0.823	0.708	
PHL	0.661	0.597	0.558	0.434	0.831	0.799	0.781	0.725	
SGP	0.765	0.627	0.439	0.270	0.882	0.814	0.721	0.639	
KOR	0.882	0.719	0.646	0.616	0.941	0.860	0.825	0.813	
TWN	0.875	0.729	0.666	0.500	0.938	0.865	0.834	0.753	
THA	0.838	0.753	0.702	0.563	0.919	0.877	0.853	0.789	
USA	0.938	0.815	0.702	0.514	0.969	0.908	0.853	0.762	
CAN	0.928	0.785	0.654	0.475	0.964	0.893	0.828	0.742	
DNK	0.880	0.797	0.630	0.459	0.940	0.899	0.816	0.733	
FRA	0.868	0.678	0.613	0.522	0.934	0.840	0.807	0.763	
DEU	0.885	0.724	0.597	0.499	0.943	0.863	0.800	0.755	
NLD	0.809	0.752	0.640	0.525	0.905	0.877	0.822	0.767	
NOR	0.962	0.809	0.624	0.311	0.981	0.905	0.813	0.659	
GBR	0.899	0.722	0.563	0.382	0.949	0.861	0.782	0.694	
AsiaDev	0.860	0.721	0.639	0.533	0.930	0.861	0.821	0.769	
EMR	0.826	0.740	0.670	0.508	0.913	0.871	0.836	0.758	
EU	0.876	0.722	0.588	0.421	0.938	0.861	0.795	0.713	

Table B.1. Accuracy Ratios (AR) and Area Under Receiver Operating Characteristic (AUROC) for different economies.

Notes: The calibration groups, Developed Asia, Emerging Markets and Europe, are indicated by AsiaDev, EMR and EU. Only economies with more than 20 defaults are listed.

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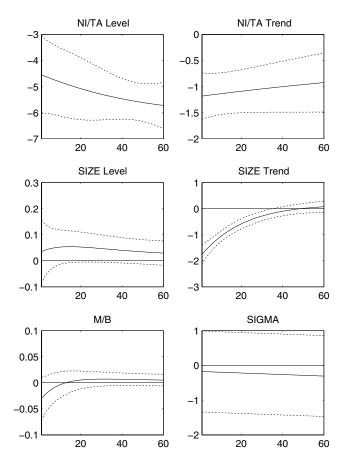
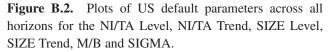


Figure B.1. Plots of US default parameters across all horizons for the Stock index one-year return, Short-term interest rate, DTD Level, DTD Trend, CASH/TA Level and CASH/TA Trend.

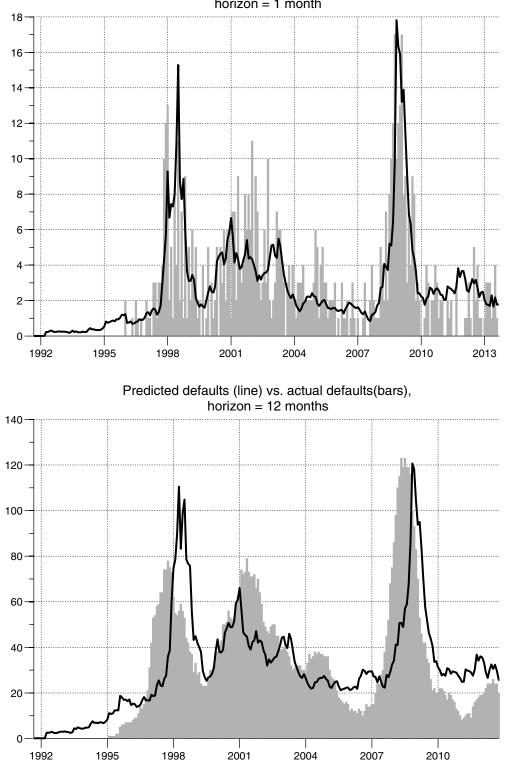
Notes: Solid lines are the parameter estimates and dashed lines are the 90% confidence level. The horizontal axis is the horizon in months.



Notes: Solid lines are the parameter estimates and dashed lines are the 90% confidence level. The horizontal axis is the horizon in months.

156 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

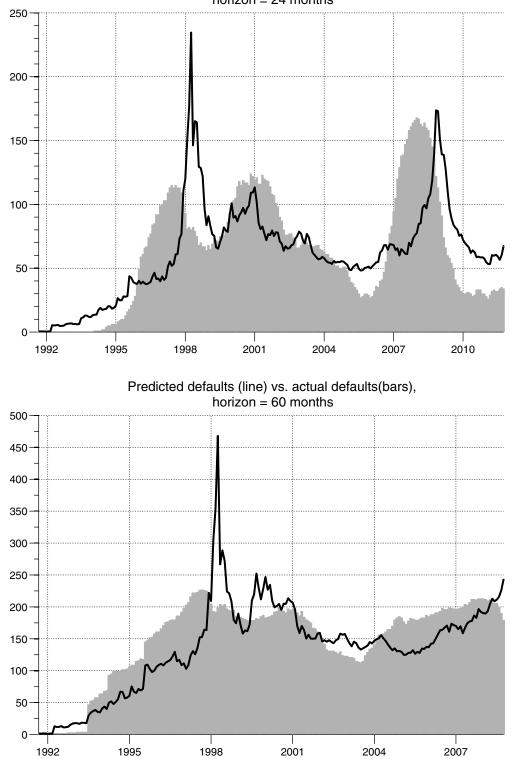
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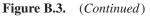


Predicted defaults (line) vs. actual defaults(bars), horizon = 1 month

Figure B.3. Performance test for Developed Asia, in sample.

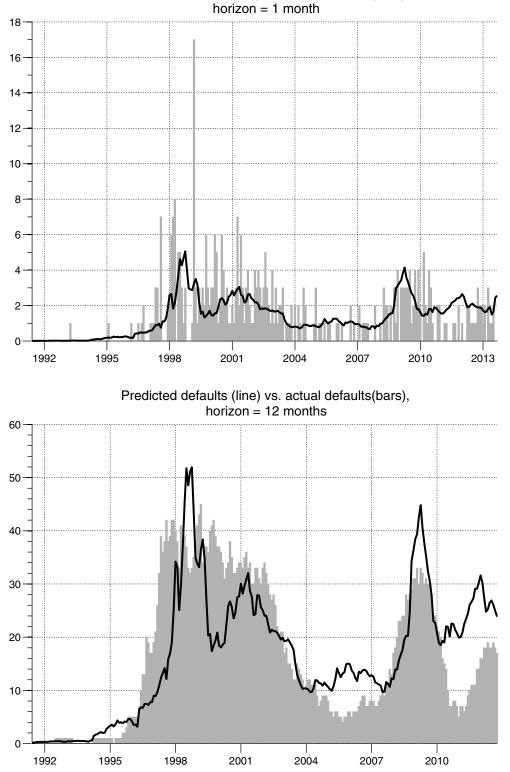
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158 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

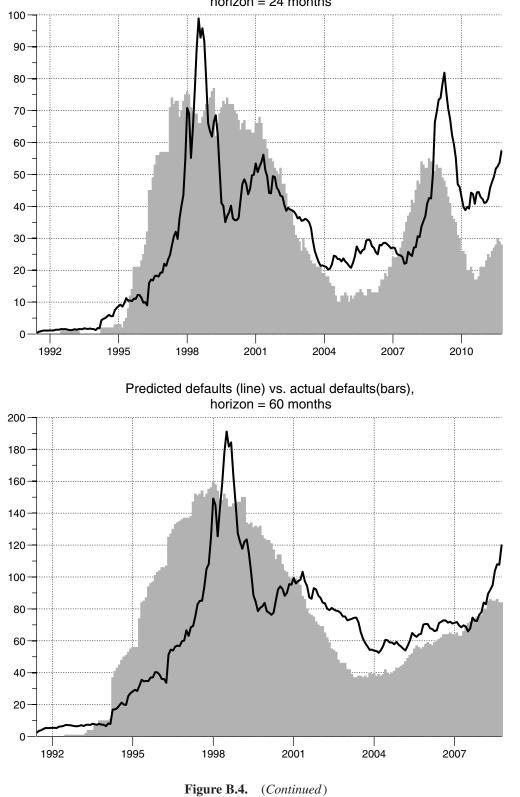
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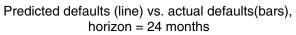


Predicted defaults (line) vs. actual defaults(bars),

Figure B.4. Performance test for the Emerging Markets group, in sample.

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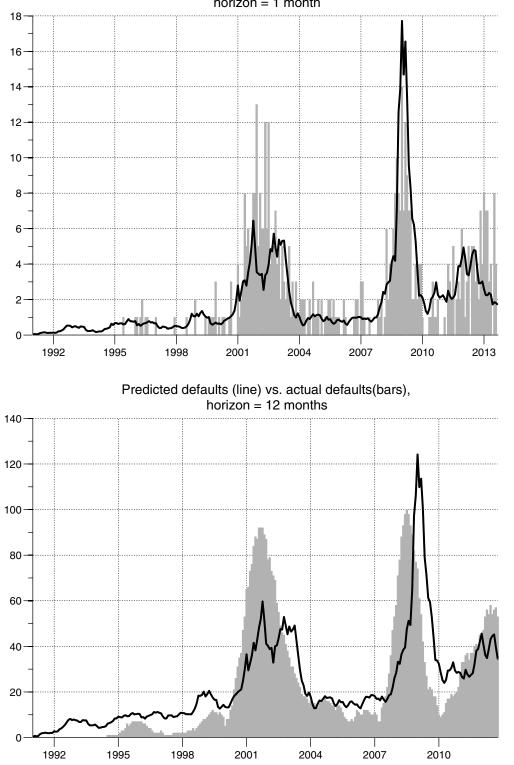




160 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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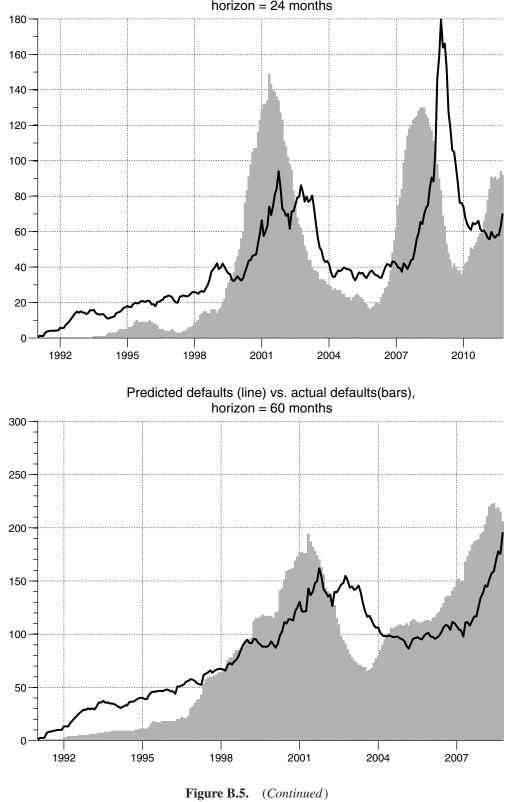
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Predicted defaults (line) vs. actual defaults(bars), horizon = 1 month

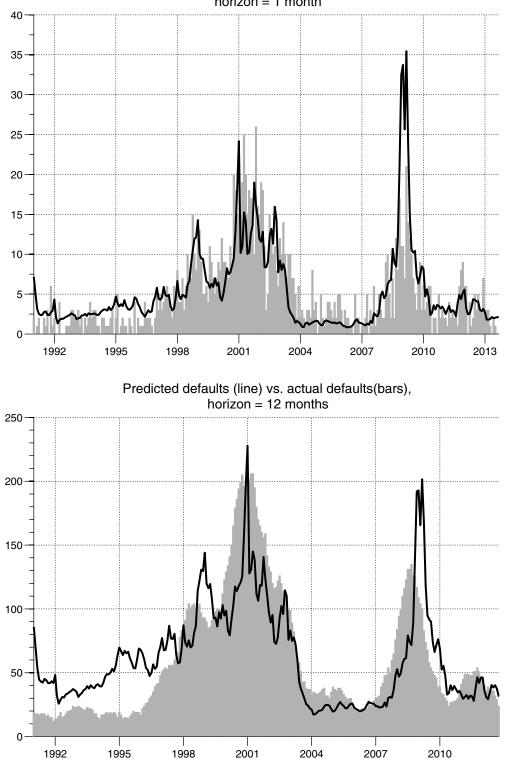
Figure B.5. Performance test for the Europe group, in sample.

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162 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

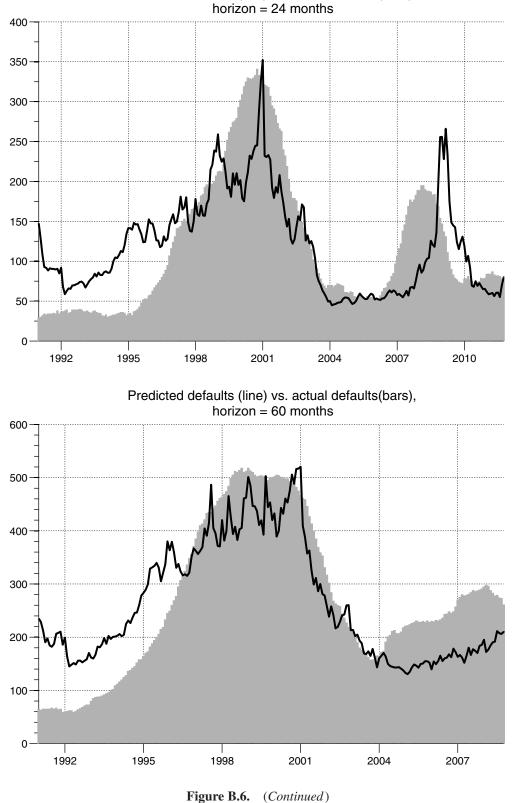
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Predicted defaults (line) vs. actual defaults(bars), horizon = 1 month

Figure B.6. Performance test for the North America group, in sample.

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Predicted defaults (line) vs. actual defaults(bars),

164 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

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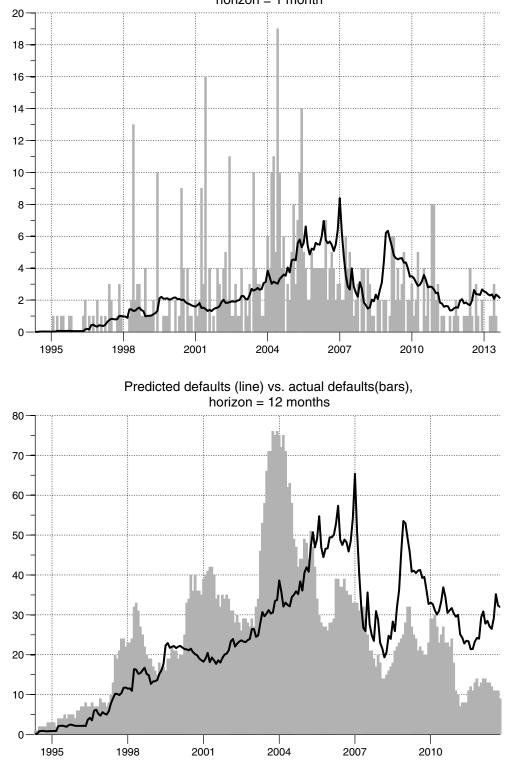
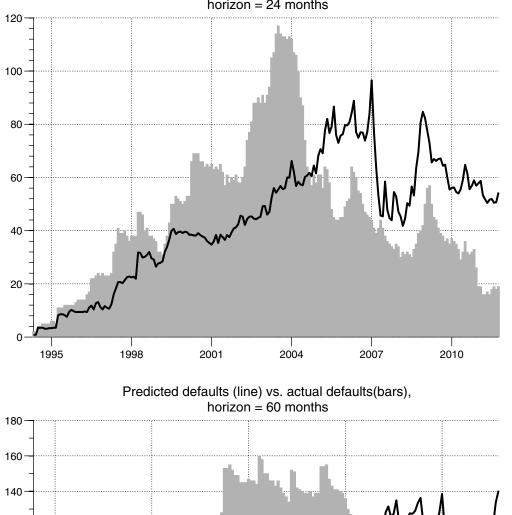


Figure B.7. Performance test for China, in sample.

GLOBAL CREDIT REVIEW VOLUME 3 165

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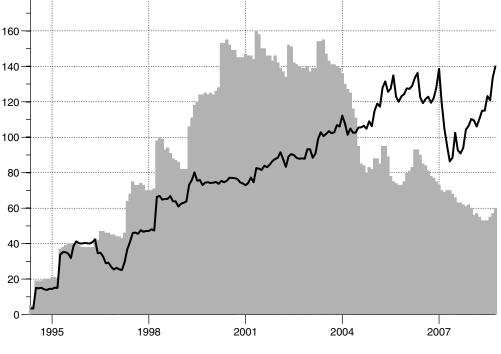
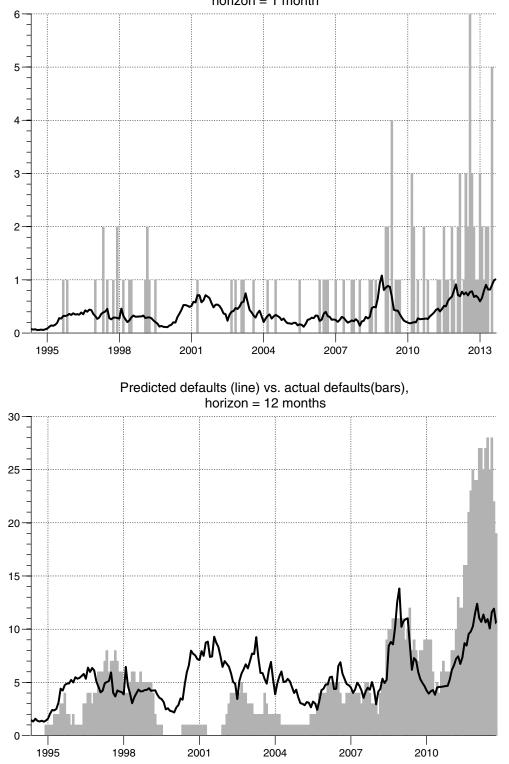


Figure B.7. (Continued)

166 NUS-RMI CREDIT RESEARCH INITIATIVE TECHNICAL REPORT

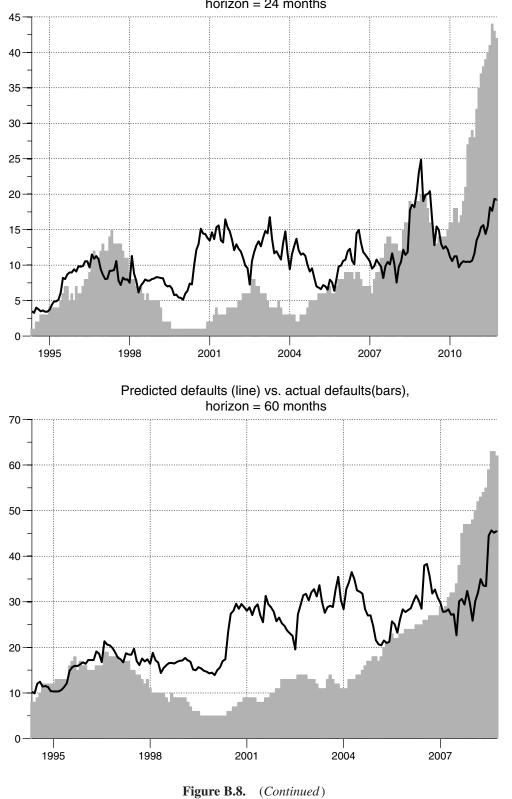
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Predicted defaults (line) vs. actual defaults(bars), horizon = 1 month

Figure B.8. Performance test for India, in sample.

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 $168 \quad {\rm nus-rmi\ credit\ research\ initiative\ technical\ report}$

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